Balanced Allocations with Heterogeneous Bins: The Power of Memory

<u>Dimitrios Los</u>¹, Thomas Sauerwald¹, John Sylvester²

 $^1\mathrm{University}$ of Cambridge, UK, $^2\mathrm{University}$ of Liverpool, UK





Allocate m tasks (balls) sequentially into n machines (bins).

Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the maximum load $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the maximum load $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.



Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the maximum load $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.



Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t. \Leftrightarrow minimise the **gap**, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the maximum load $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t. \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Allocate m tasks (balls) sequentially into n machines (bins).

<u>Goal</u>: minimise the maximum load $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t. \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Applications in hashing [PR01], load balancing [Wie16] and routing [GKK88].

<u>ONE-CHOICE Process</u>: Iteration: For each $t \ge 0$, sample one bin uniformly at random (u.a.r.) and place the ball there.

<u>ONE-CHOICE Process</u>: Iteration: For each $t \ge 0$, sample one bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

<u>ONE-CHOICE Process</u>: Iteration: For each $t \ge 0$, sample one bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case
$$(m = n)$$
, w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
Meaning with probability
at least $1 - n^{-c}$ for constant $c > 0$.

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

<u>Two-Choice Process</u>: **Iteration**: For each $t \ge 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

<u>Two-Choice Process</u>: Iteration: For each $t \ge 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n}}, \log n\right)$ (e.g. [RS98]).

<u>TWO-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \ge 0$, sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \log_2 \log n + \Theta(1)$ [BCSV06].

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

<u>Two-Choice Process</u>: **Iteration**: For each $t \ge 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \log_2 \log n + \Theta(1)$ [KLMadH96, ABKU99].

In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \log_2 \log n + \Theta(1)$ [BCSV06].

<u>ONE-CHOICE Process</u>: **Iteration**: For each $t \ge 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

<u>*d*-CHOICE Process:</u> **Iteration**: For each $t \ge 0$, sample <u>*d*</u> bins independently u.a.r. and place the ball in the least loaded of the two.

In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \log_d \log n + \bigoplus'(1)$ [KLMadH96, ABKU99].

In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \log_d \log n + \Theta(1)$ [BCSV06].

Power of two choices: Visualisation



The Memory process

Several different variants of *d*-CHOICE have been studied: $(1 + \beta)$ [PTW15], THINNING [FGGL21].

The Memory process

- Several different variants of *d*-CHOICE have been studied: $(1 + \beta)$ [PTW15], THINNING [FGGL21].
- Shah and Prabhakar [SP02] introduced a variant of d-CHOICE maintaining M cached bins.

- Several different variants of *d*-CHOICE have been studied: $(1 + \beta)$ [PTW15], THINNING [FGGL21].
- Shah and Prabhakar [SP02] introduced a variant of d-CHOICE maintaining M cached bins.

 $\begin{array}{l} \displaystyle \underbrace{\text{MEMORY Process } (M=1):}_{\text{Initialization: Set the cache } c^0 = 1.\\ \text{Iteration: For each step } t \geq 0:\\ \hline & \text{Sample bins } i_1, \ldots, i_d \text{ uniformly at random.}\\ \hline & \text{Allocate to bin } j = \operatorname{argmin}_{k \in \{c^t, i_1, \ldots, i_d\}} x_k^t. \end{array}$

• Update the cache to $c^{t+1} = \operatorname{argmin}_{k \in \{c^t, i_1, \dots, i_d\}} x_k^{t+1}$.

















In the lightly-loaded case, MEMORY with d = 1 w.h.p. achieves an $\mathcal{O}(\log \log n)$ gap [MPS02].



- In the lightly-loaded case, MEMORY with d = 1 w.h.p. achieves an $\mathcal{O}(\log \log n)$ gap [MPS02].
- For general $d \ge 1$, the bound becomes $\log_{f(d)} \log n + \Theta(1)$ for $f(d) \in (2d, 2d + 1)$.



In the lightly-loaded case, MEMORY with d = 1 w.h.p. achieves an $\mathcal{O}(\log \log n)$ gap [MPS02].

For general $d \ge 1$, the bound becomes $\log_{f(d)} \log n + \Theta(1)$ for $f(d) \in (2d, 2d + 1)$.

What happens in the heavily-loaded case $(m \ge n)$?

Heterogeneous sampling distributions

Heterogeneous sampling distributions

Several different settings for d**-CHOICE**:
Several different settings for *d*-CHOICE: outdated information [BCE⁺12],

 Several different settings for *d*-CHOICE: outdated information [BCE⁺12], graphical [BK22],

Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],

Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],

- Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.

- Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



- Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



- Several different settings for d-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



In particular, (a, b)-biased sampling distributions s satisfy $\frac{1}{an} \leq s_i \leq \frac{b}{n}$.

- Several different settings for *d*-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



In particular, (a, b)-biased sampling distributions s satisfy ¹/_{an} ≤ s_i ≤ ^b/_n.
Given a, b > 1, Wieder showed that there exists d' > 0, such that for any ε > 0:

- Several different settings for *d*-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



In particular, (a, b)-biased sampling distributions s satisfy 1/an ≤ s_i ≤ b/n.
Given a, b > 1, Wieder showed that there exists d' > 0, such that for any ε > 0:
For any d ≥ (1 + ε) ⋅ d', then d-CHOICE w.h.p. achieves Gap(m) = O(log log n).

- Several different settings for *d*-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



In particular, (a, b)-biased sampling distributions s satisfy ¹/_{an} ≤ s_i ≤ ^b/_n.
Given a, b > 1, Wieder showed that there exists d' > 0, such that for any ε > 0:
For any d ≥ (1 + ε) ⋅ d', then d-CHOICE w.h.p. achieves Gap(m) = O(log log n).

For any $d \leq (1 - \epsilon) \cdot d'$, then d-CHOICE has a gap that grows with m.

- Several different settings for *d*-CHOICE: outdated information [BCE⁺12], graphical [BK22], adversarial noise [LS22b],
- Wieder [Wie07] studied *d*-CHOICE with non-uniform sampling distributions.



In particular, (a, b)-biased sampling distributions s satisfy ¹/_{an} ≤ s_i ≤ ^b/_n.
Given a, b > 1, Wieder showed that there exists d' > 0, such that for any ε > 0:

- ▶ For any $d \ge (1 + \epsilon) \cdot d'$, then *d*-CHOICE w.h.p. achieves $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- ▶ For any $d \leq (1 \epsilon) \cdot d'$, then *d*-CHOICE has a gap that grows with *m*.

How does MEMORY deal with heterogeneous sampling distributions?

■ In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$.

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.
- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $Gap(m) = O(\log \log n)$.

- In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.
- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to TWO-CHOICE, where the gap grows with m, for a = b = 2.

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to Two-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to Two-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

Challenges:

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to Two-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

Challenges: (i) long-term dependencies due to cache

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to Two-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

Challenges: (i) long-term dependencies due to cache and (ii) biased sampling.

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to TWO-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

Challenges: (i) long-term dependencies due to cache and (ii) biased sampling.

■ *d*-RESET-MEMORY, a variant of MEMORY where the cache resets every *d* steps has w.h.p. $Gap(m) = O(\log n)$

In the heavily-loaded case $(m \ge n)$, [LSS22] proved that MEMORY (with d = M = 1) achieves w.h.p. $\mathcal{O}(\log n)$. We improve this to $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Further, we show that w.h.p. $\operatorname{Gap}(m) = \Omega(\log \log n)$ for any $m \ge n$.

- For (a, b)-biased distributions with any const a, b > 1, w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$. \leftarrow In contrast to TWO-CHOICE, where the gap grows with m, for a = b = 2.
- For any a := a(n) and b := b(n), the gap is independent of m.

Challenges: (i) long-term dependencies due to cache and (ii) biased sampling.

■ *d*-RESET-MEMORY, a variant of MEMORY where the cache resets every *d* steps has w.h.p. $Gap(m) = O(\log n)$, even in the presence of weights.

Power of memory: Visualisation



Balanced allocations: Background

	Two-Choice														
	000000														
									000000	0000		000			

Memory

Balanced allocations: Background



In Two-Choice, there is a set of bins that receives > m/n balls in expectation.



■ In Two-Choice, there is a set of bins that receives > m/n balls in expectation.

In MEMORY, w.h.p. we sample every bin roughly every $an \log n$ steps.

Balanced allocations: Background



In Two-Choice, there is a set of bins that receives > m/n balls in expectation.

■ In MEMORY, w.h.p. we sample every bin roughly every $an \log n$ steps.

Balanced allocations: Background

Upper Bound for MEMORY

■ Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i: x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants $z > 0, \alpha \in (0, 1)$ and v > 1.

■ Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i: x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants z > 0, $\alpha \in (0, 1)$ and v > 1.

• When $\Phi_j^t = \mathcal{O}(n)$, then $\operatorname{Gap}(t) = \mathcal{O}(j \cdot z + \frac{\log n}{\alpha \cdot v^j})$.

■ Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i: x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants z > 0, $\alpha \in (0, 1)$ and v > 1.

When
$$\Phi_j^t = \mathcal{O}(n)$$
, then $\operatorname{Gap}(t) = \mathcal{O}(j \cdot z + \frac{\log n}{\alpha \cdot v^j})$.

For $j = \Theta(\log \log n)$, when $\Phi_j = \mathcal{O}(n)$, then $\operatorname{Gap}(m) = \Theta(\log \log n)$.

Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i:x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants $z > 0, \alpha \in (0, 1)$ and v > 1.

- When $\Phi_j^t = \mathcal{O}(n)$, then $\operatorname{Gap}(t) = \mathcal{O}(j \cdot z + \frac{\log n}{\alpha \cdot v^j})$.
- For $j = \Theta(\log \log n)$, when $\Phi_j = \mathcal{O}(n)$, then $\operatorname{Gap}(m) = \Theta(\log \log n)$.
- Further, when $\Phi_j^t = \mathcal{O}(n)$, then also number of bins with load at least z_{j+1} is at most $\mathcal{O}(n \cdot e^{-\alpha \cdot v^j \cdot z})$.

Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i:x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants $z > 0, \alpha \in (0, 1)$ and v > 1.

- When $\Phi_j^t = \mathcal{O}(n)$, then $\operatorname{Gap}(t) = \mathcal{O}(j \cdot z + \frac{\log n}{\alpha \cdot v^j})$.
- For $j = \Theta(\log \log n)$, when $\Phi_j = \mathcal{O}(n)$, then $\operatorname{Gap}(m) = \Theta(\log \log n)$.
- Further, when $\Phi_j^t = \mathcal{O}(n)$, then also number of bins with load at least z_{j+1} is at most $\mathcal{O}(n \cdot e^{-\alpha \cdot v^j \cdot z})$.
- We group steps into rounds (at most $e^{v^{j+2}} \cdot \log^3 n$ steps each) and show that

$$\mathbf{E}\left[\left.\Phi_{j+1}^{r+1}\right|\,\mathfrak{F}^{r},\Phi_{j}^{r}=\mathcal{O}(n)\right] \leq \Phi_{j+1}^{r}\cdot\left(1-\frac{e^{v^{j+2}}}{n}\right)+e^{-v^{j+1}/2}$$
Outline for the $\mathcal{O}(\log \log n)$ bound

■ Define the super-exponential potentials Φ_j for $0 \le j = \mathcal{O}(\log \log n)$,

$$\Phi_j^t := \Phi_j^t(\alpha \cdot v^j, z_j) := \sum_{i:x_i^t \ge z_j} e^{\alpha \cdot v^j \cdot (x_i^t - z_j)},$$

where $z_j := \frac{t}{n} + j \cdot z$ for constants $z > 0, \alpha \in (0, 1)$ and v > 1.

- When $\Phi_j^t = \mathcal{O}(n)$, then $\operatorname{Gap}(t) = \mathcal{O}(j \cdot z + \frac{\log n}{\alpha \cdot v^j})$.
- For $j = \Theta(\log \log n)$, when $\Phi_j = \mathcal{O}(n)$, then $\operatorname{Gap}(m) = \Theta(\log \log n)$.
- Further, when $\Phi_j^t = \mathcal{O}(n)$, then also number of bins with load at least z_{j+1} is at most $\mathcal{O}(n \cdot e^{-\alpha \cdot v^j \cdot z})$.
- We group steps into rounds (at most $e^{v^{j+2}} \cdot \log^3 n$ steps each) and show that

$$\mathbf{E}\left[\left.\Phi_{j+1}^{r+1}\right| \,\mathfrak{F}^{r}, \Phi_{j}^{r} = \mathcal{O}(n)\right] \le \Phi_{j+1}^{r} \cdot \left(1 - \frac{e^{v^{j+2}}}{n}\right) + e^{-v^{j+1}/2}.$$

 The base case follows by an involved analysis of the hyperbolic cosine potential function [PTW15, LS22a].

Upper Bound for MEMORY

















Conclusion

We have shown that:

We have shown that:

• MEMORY with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

We have shown that:

MEMORY with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.

Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.

We have shown that:

- MEMORY with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- *d*-Reset-Memory has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

Several avenues for future work:

• What is the gap for the optimal caching strategy at step m?

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

- What is the gap for the optimal caching strategy at step m?
- Are there any weighted settings where **MEMORY** is superior to *d*-CHOICE?

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

- What is the gap for the optimal caching strategy at step m?
- Are there any weighted settings where **MEMORY** is superior to *d*-CHOICE?
- Obtaining tight bounds for (a, b)-biased distributions for **non-const** a, b.

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

- What is the gap for the optimal caching strategy at step m?
- Are there any weighted settings where **MEMORY** is superior to *d*-CHOICE?
- Obtaining tight bounds for (a, b)-biased distributions for **non-const** a, b.
- Obtaining tight bounds up to lower order terms (as in [MPS02]).

We have shown that:

- **MEMORY** with d = M = 1 has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log \log n)$.
- Same upper bound for (a, b)-biased sampling distributions with any const a, b > 1.
- A matching lower bound holds for any $m \ge n$.
- d-RESET-MEMORY has w.h.p. $\operatorname{Gap}(m) = \mathcal{O}(\log n)$.

- What is the gap for the optimal caching strategy at step m?
- Are there any weighted settings where **MEMORY** is superior to *d*-CHOICE?
- Obtaining tight bounds for (a, b)-biased distributions for **non-const** a, b.
- Obtaining tight bounds up to lower order terms (as in [MPS02]).
- Analyse MEMORY in settings with outdated or noisy information.

Questions?



More visualisations: dimitrioslos.com/soda23 $\,$

Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.

Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.

For ONE-CHOICE,
$$p^t = \left(\frac{1}{n}, \dots, \frac{1}{n}\right)$$
.

- Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.
- For ONE-CHOICE, $p^t = \begin{pmatrix} \frac{1}{n}, & \dots, & \frac{1}{n} \end{pmatrix}$.
- For Two-Choice,

$$p^{t} = \left(\frac{1}{n^{2}}, \dots, \frac{2i-1}{n^{2}}, \dots, \frac{2n-1}{n^{2}}\right).$$

- Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.
- For ONE-CHOICE, $p^t = \begin{pmatrix} \frac{1}{n}, & \dots, & \frac{1}{n} \end{pmatrix}$.
- For Two-Choice,

$$p^{t} = \left(\frac{1}{n^{2}}, \dots, \frac{2i-1}{n^{2}}, \dots, \frac{2n-1}{n^{2}}\right).$$

For MEMORY, if the cache is the k-th most loaded bin, then

$$p^{t} = \left(\underbrace{0, \dots, 0}_{k-1 \text{ bins}}, \underbrace{\frac{k}{n}, \frac{1}{n}, \dots, \frac{1}{n}}_{n-k \text{ bins}}\right)$$

.

- Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.
- For ONE-CHOICE, $p^t = (\frac{1}{n}, \ldots, \frac{1}{n})$.
- For Two-Choice,

$$p^{t} = \left(\frac{1}{n^{2}}, \dots, \frac{2i-1}{n^{2}}, \dots, \frac{2n-1}{n^{2}}\right).$$

For MEMORY, if the cache is the k-th most loaded bin, then

$$p^{t} = \left(\underbrace{0, \dots, 0}_{k-1 \text{ bins}}, \underbrace{\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}}_{n-k \text{ bins}}\right).$$

For $k = 1$, this is like ONE-CHOICE.

- Some processes induce a **probability allocation vector** p^t , where p_i^t gives the probability to allocate to the *i*-th most loaded bin.
- For ONE-CHOICE, $p^t = \left(\frac{1}{n}, \dots, \frac{1}{n}\right)$.
- For Two-Choice,

$$p^{t} = \left(\frac{1}{n^{2}}, \dots, \frac{2i-1}{n^{2}}, \dots, \frac{2n-1}{n^{2}}\right).$$

For MEMORY, if the cache is the k-th most loaded bin, then

$$p^{t} = \left(\underbrace{0, \dots, 0}_{k-1 \text{ bins}}, \frac{k}{n}, \underbrace{\frac{1}{n}, \dots, \frac{1}{n}}_{n-k \text{ bins}}\right)$$

A probability vector p is (δ, ϵ) -smooth if majorized by

$$\left(\underbrace{\frac{1-\epsilon}{n}, \ldots, \frac{1-\epsilon}{n}}_{\delta n \text{ bins}}, \underbrace{\frac{1+\tilde{\epsilon}}{n}, \ldots, \frac{1+\tilde{\epsilon}}{n}}_{(1-\delta)n \text{ bins}}\right).$$

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

• When
$$\Gamma^m = \text{poly}(n)$$
, then $\text{Gap}(m) = \mathcal{O}\left(\frac{\log n}{\alpha}\right)$.

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

• When
$$\Gamma^m = \operatorname{poly}(n)$$
, then $\operatorname{Gap}(m) = \mathcal{O}\left(\frac{\log n}{\alpha}\right)$.

They showed that for any (δ, ϵ) -smooth probability allocation vector p^t ,

$$\mathbf{E}\left[\Gamma^{t+1} \mid \mathfrak{F}^t\right] \leq \Gamma^t \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

When
$$\Gamma^m = \operatorname{poly}(n)$$
, then $\operatorname{Gap}(m) = \mathcal{O}\left(\frac{\log n}{\alpha}\right)$.

They showed that for any (δ, ϵ) -smooth probability allocation vector p^t ,

$$\mathbf{E}\left[\Gamma^{t+1} \mid \mathfrak{F}^t\right] \leq \Gamma^t \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

By induction, this implies that $\mathbf{E}[\Gamma^m] \leq \frac{c}{\alpha\epsilon} \cdot n$.

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

• When
$$\Gamma^m = \operatorname{poly}(n)$$
, then $\operatorname{Gap}(m) = \mathcal{O}\left(\frac{\log n}{\alpha}\right)$.

■ They showed that for any (δ, ϵ) -smooth probability allocation vector p^t ,

$$\mathbf{E}\left[\Gamma^{t+1} \mid \mathfrak{F}^t\right] \leq \Gamma^t \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

By induction, this implies that E [Γ^m] ≤ c/αε ⋅ n.
And so, by Markov's inequality Pr [Γ^m ≤ c/αε ⋅ n³] ≥ 1 − n⁻².

Peres, Talwar and Wieder [PTW15] used the hyperbolic cosine potential Γ^t , defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha(x_i^t - t/n)} + \sum_{i=1}^n e^{-\alpha(x_i^t - t/n)}.$$

• When
$$\Gamma^m = \operatorname{poly}(n)$$
, then $\operatorname{Gap}(m) = \mathcal{O}\left(\frac{\log n}{\alpha}\right)$.

■ They showed that for any (δ, ϵ) -smooth probability allocation vector p^t ,

$$\mathbf{E}\left[\Gamma^{t+1} \mid \mathfrak{F}^t\right] \leq \Gamma^t \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

By induction, this implies that $\mathbf{E} [\Gamma^m] \leq \frac{c}{\alpha \epsilon} \cdot n$. And so, by Markov's inequality $\mathbf{Pr} [\Gamma^m \leq \frac{c}{\alpha \epsilon} \cdot n^3] \geq 1 - n^{-2}$.

Problem: p^t for MEMORY may not be (δ, ϵ) -smooth
If for some (δ, ϵ) -smooth probability vector q,

$$\begin{split} \mathbf{E}\left[\begin{array}{c} \Phi^{t+1} \mid \mathfrak{F}^t \end{array} \right] &\leq \Phi^t + \sum_{i=1}^n \Phi^t_i \cdot \left(q^t_i - \frac{1}{n} \right) \cdot \alpha + \Phi^t \cdot C \cdot \frac{\alpha^2}{n}, \\ \mathbf{E}\left[\begin{array}{c} \Psi^{t+1} \mid \mathfrak{F}^t \end{array} \right] &\leq \Psi^t + \sum_{i=1}^n \Psi^t_i \cdot \left(\frac{1}{n} - q^t_i \right) \cdot \alpha + \Psi^t \cdot C \cdot \frac{\alpha^2}{n}. \end{split}$$

If for some (δ, ϵ) -smooth probability vector q,

$$\begin{split} \mathbf{E}\left[\begin{array}{c|c} \Phi^{t+1} & \mathfrak{F}^t \end{array} \right] &\leq \Phi^t + \sum_{i=1}^n \Phi^t_i \cdot \left(q^t_i - \frac{1}{n} \right) \cdot \alpha + \Phi^t \cdot C \cdot \frac{\alpha^2}{n}, \\ \mathbf{E}\left[\begin{array}{c|c} \Psi^{t+1} & \mathfrak{F}^t \end{array} \right] &\leq \Psi^t + \sum_{i=1}^n \Psi^t_i \cdot \left(\frac{1}{n} - q^t_i \right) \cdot \alpha + \Psi^t \cdot C \cdot \frac{\alpha^2}{n}. \end{split}$$

• Then, for sufficiently small $\alpha > 0$,

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

If for some (δ, ϵ) -smooth probability vector q,

$$\begin{split} \mathbf{E}\left[\begin{array}{c} \Phi^{t+1} \mid \mathfrak{F}^t\right] &\leq \Phi^t + \sum_{i=1}^n \Phi^t_i \cdot \left(q_i^t - \frac{1}{n}\right) \cdot \alpha + \Phi^t \cdot C \cdot \frac{\alpha^2}{n}, \\ \end{split} \\ \begin{aligned} & \text{Could be allocating} \\ & \text{more than one ball.} \end{aligned} \\ \mathbf{E}\left[\begin{array}{c} \Psi^{t+1} \mid \mathfrak{F}^t\right] &\leq \Psi^t + \sum_{i=1}^n \Psi^t_i \cdot \left(\frac{1}{n} - q_i^t\right) \cdot \alpha + \Psi^t \cdot C \cdot \frac{\alpha^2}{n}. \end{aligned}$$

• Then, for sufficiently small $\alpha > 0$,

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

If for some (δ, ϵ) -smooth probability vector q, \Leftrightarrow not always the prob allocation vector.

$$\begin{split} \mathbf{E}\left[\begin{array}{c} \Phi^{t+1} \mid \mathfrak{F}^t \end{array} \right] &\leq \Phi^t + \sum_{i=1}^n \Phi^t_i \cdot \left(q^t_i - \frac{1}{n} \right) \cdot \alpha + \Phi^t \cdot C \cdot \frac{\alpha^2}{n}, \\ \mathbf{E}\left[\begin{array}{c} \Psi^{t+1} \mid \mathfrak{F}^t \end{array} \right] &\leq \Psi^t + \sum_{i=1}^n \Psi^t_i \cdot \left(\frac{1}{n} - q^t_i \right) \cdot \alpha + \Psi^t \cdot C \cdot \frac{\alpha^2}{n}. \end{split}$$

• Then, for sufficiently small $\alpha > 0$,

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

If for some (δ, ϵ) -smooth probability vector q, \Leftrightarrow not always the prob allocation vector.

$$\begin{split} \mathbf{E}\left[\left.\Phi^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right] &\leq \Phi^{t} + \sum_{i=1}^{n}\Phi^{t}_{i}\cdot\left(q^{t}_{i}-\frac{1}{n}\right)\cdot\alpha + \Phi^{t}\cdot C\cdot\frac{\alpha^{2}}{n},\\ \mathbf{E}\left[\left.\Psi^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right] &\leq \Psi^{t} + \sum_{i=1}^{n}\Psi^{t}_{i}\cdot\left(\frac{1}{n}-q^{t}_{i}\right)\cdot\alpha + \Psi^{t}\cdot C\cdot\frac{\alpha^{2}}{n}. \end{split}$$

• Then, for sufficiently small $\alpha > 0$,

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c.$$

For 2-RESET-MEMORY, q is the probability allocation vector of Two-CHOICE.

If for some (δ, ϵ) -smooth probability vector q, \Leftrightarrow not always the prob allocation vector.

$$\begin{split} \mathbf{E}\left[\left.\Phi^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right] &\leq \Phi^{t} + \sum_{i=1}^{n}\Phi^{t}_{i}\cdot\left(q^{t}_{i}-\frac{1}{n}\right)\cdot\alpha + \Phi^{t}\cdot C\cdot\frac{\alpha^{2}}{n},\\ \mathbf{E}\left[\left.\Psi^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right] &\leq \Psi^{t} + \sum_{i=1}^{n}\Psi^{t}_{i}\cdot\left(\frac{1}{n}-q^{t}_{i}\right)\cdot\alpha + \Psi^{t}\cdot C\cdot\frac{\alpha^{2}}{n}. \end{split}$$

• Then, for sufficiently small $\alpha > 0$,

$$\mathbf{E}\left[\left|\Gamma^{t+1}\right| \mathfrak{F}^{t}\right] \leq \Gamma^{t} \cdot \left(1 - \frac{\alpha \epsilon}{n}\right) + c$$

For 2-RESET-MEMORY, q is the probability allocation vector of TWO-CHOICE.
which is (1/4, 1/2)-smooth, implying an O(log n) gap for MEMORY.

To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - $\,\triangleright\,$ Cache resets every d steps.

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.

- To analyze a heterogeneous sampling distribution $s (\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - ▶ Load comparisons are based on the last reset.

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - ▶ Load comparisons are based on the last reset. \leftarrow makes computation of q tractable.

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.

• Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right]\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q_{i}^{t} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.

• Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right|\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q^{t}_{i} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

since $\Phi_i^t - \Phi_j^t \le \Phi_j^t \cdot (2\alpha d)$ and

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.
- Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right|\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q^{t}_{i} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

since $\Phi_i^t - \Phi_j^t \leq \Phi_j^t \cdot (2\alpha d)$ and probability of selecting a bin twice is at most $d^2 \cdot \frac{b}{n}$.

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.
- Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left.\Phi^{t+1}\right|\left.\mathfrak{F}^{t}\right.\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q^{t}_{i} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

since $\Phi_i^t - \Phi_j^t \leq \Phi_j^t \cdot (2\alpha d)$ and probability of selecting a bin twice is at most $d^2 \cdot \frac{b}{n}$. Similarly for Ψ .

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.
- Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right]\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q^{t}_{i} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

since $\Phi_i^t - \Phi_j^t \leq \Phi_j^t \cdot (2\alpha d)$ and probability of selecting a bin twice is at most $d^2 \cdot \frac{b}{n}$. Similarly for Ψ . So for sufficiently small $\alpha := \alpha(d) > 0$, $\mathbf{E}[\Gamma^m] = \mathcal{O}(n)$.

- To analyze a heterogeneous sampling distribution s $(\frac{1}{an} \le s_i \le \frac{b}{n})$, we make two further reductions:
 - ▶ Cache resets every d steps. \leftarrow for sufficiently large d, beats the (a, b)-bias.
 - \blacktriangleright Load comparisons are based on the last reset. \leftrightsquigarrow makes computation of q tractable.
- Moving probabilities between bins with almost the same load, introduces a small additive term in the bound,

$$\mathbf{E}\left[\left|\Phi^{t+1}\right|\left|\mathfrak{F}^{t}\right]\right] \leq \Phi^{t} + \sum_{i=1}^{n} \Phi^{t}_{i} \cdot \left(q^{t}_{i} - \frac{1}{n}\right) \cdot \alpha + \Phi^{t} \cdot C \cdot \frac{\alpha^{2}}{n} + \mathcal{O}\left(\Phi^{t} \cdot \frac{\alpha^{2}}{n} \cdot (2d^{3}b)\right),$$

since $\Phi_i^t - \Phi_j^t \leq \Phi_j^t \cdot (2\alpha d)$ and probability of selecting a bin twice is at most $d^2 \cdot \frac{b}{n}$. Similarly for Ψ . So for sufficiently small $\alpha := \alpha(d) > 0$, $\mathbf{E}[\Gamma^m] = \mathcal{O}(n)$.

And so $\operatorname{Gap}(m) = \mathcal{O}((\log n)/\alpha)$ gap.

Bibliography I

- Y. Azar, A. Z. Broder, A. R. Karlin, and E. Upfal, *Balanced allocations*, SIAM J. Comput. 29 (1999), no. 1, 180–200. MR 1710347
- P. Berenbrink, A. Czumaj, M. Englert, T. Friedetzky, and L. Nagel, *Multiple-choice balanced allocation in (almost) parallel*, 16th International Workshop on Randomization and Computation (RANDOM'12) (Berlin Heidelberg), Springer-Verlag, 2012, pp. 411–422.
- P. Berenbrink, A. Czumaj, A. Steger, and B. Vöcking, Balanced allocations: the heavily loaded case, SIAM J. Comput. 35 (2006), no. 6, 1350–1385. MR 2217150
- N. Bansal and W. Kuszmaul, Balanced allocations: The heavily loaded case with deletions, 63rd Annual IEEE Symposium on Foundations of Computer Science (FOCS'22), IEEE, 2022, pp. 801–812.
- O. N. Feldheim, O. Gurel-Gurevich, and J. Li, Long-term balanced allocation via thinning, 2021, arXiv:2110.05009.

Bibliography II

- R.J. Gibbens, F.P. Kelly, and P.B. Key, *Dynamic alternative routing modelling and behavior*, Proceedings of the 12 International Teletraffic Congress, Torino, Italy, Elsevier, Amsterdam, 1988.
- G. H. Gonnet, Expected length of the longest probe sequence in hash code searching, J. Assoc. Comput. Mach. 28 (1981), no. 2, 289–304. MR 612082
- R. M. Karp, M. Luby, and F. Meyer auf der Heide, Efficient PRAM simulation on a distributed memory machine, Algorithmica 16 (1996), no. 4-5, 517–542. MR 1407587
- D. Los and T. Sauerwald, Balanced allocations in batches: Simplified and generalized, 34th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'22) (New York, NY, USA), ACM, 2022, p. 389–399.

Balanced allocations with the choice of noise, 41st Annual ACM-SIGOPT Principles of Distributed Computing (PODC'22) (New York, NY, USA), ACM, 2022, p. 164–175.

Bibliography III

- D. Los, T. Sauerwald, and J. Sylvester, Balanced Allocations: Caching and Packing, Twinning and Thinning, 33rd Annual ACM-SIAM Symposium on Discrete Algorithms (SODA'22) (Alexandria, Virginia), SIAM, 2022, pp. 1847–1874.
- M. Mitzenmacher, B. Prabhakar, and D. Shah, Load balancing with memory, 43rd Annual IEEE Symposium on Foundations of Computer Science (FOCS'02), IEEE, 2002, pp. 799–808.
- R. Pagh and F. F. Rodler, Cuckoo hashing, Algorithms—ESA 2001 (Århus), Lecture Notes in Comput. Sci., vol. 2161, Springer, Berlin, 2001, pp. 121–133. MR 1913547
- ▶ Y. Peres, K. Talwar, and U. Wieder, Graphical balanced allocations and the $(1 + \beta)$ -choice process, Random Structures Algorithms 47 (2015), no. 4, 760–775. MR 3418914
- M. Raab and A. Steger, "Balls into bins"—a simple and tight analysis, 2nd International Workshop on Randomization and Computation (RANDOM'98), vol. 1518, Springer, 1998, pp. 159–170. MR 1729169

Bibliography IV

- D. Shah and B. Prabhakar, The use of memory in randomized load balancing, IEEE International Symposium on Information Theory (ISIT'02), 2002, p. 125.
- ▶ U. Wieder, *Balanced allocations with heterogenous bins*, 19th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'07), ACM, 2007, pp. 188–193.
- <u>Hashing</u>, load balancing and multiple choice, Found. Trends Theor. Comput. Sci. 12 (2016), no. 3-4, front matter, 276–379. MR 3683828