

Introduction to Probability

Lecture 4: More discrete distributions – Poisson, Geometric,
Negative Binomial, Hypergeometric

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Poisson discrete random variable

Geometric discrete random variable

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Preliminaries:

The natural exponent e

e is a mathematical constant AKA the Euler number. e is very important for exponential functions. Here are some important identities:

$$e \approx 2.71828$$

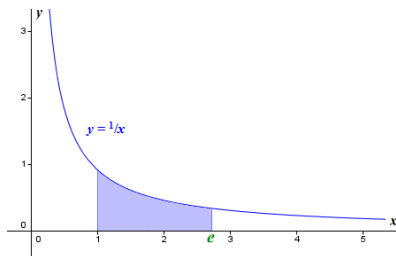
$$e = \sum_{n=0}^{\infty} \frac{1}{n!}$$

$$e = \lim_{n \rightarrow \infty} \left(1 + \frac{1}{n} \right)^n$$

$$e^{-\lambda} = \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n} \right)^n$$

$$e^r = \lim_{n \rightarrow \infty} \left(1 + \frac{r}{n} \right)^n$$

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Binomial RV example: large n , small p

We are trying to predict footfall in a store. We know, based on previous data, that on average 8 people enter the store per hour. What is the probability of k people entering the store in the next 1 hour?



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- At each **minute**, independent Bernoulli trial with 1 for a person entering the store and 0 for nobody entering the store.
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2. Break an hour into **milliseconds**.

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3. Break an hour into **infinitely small units**.

- At each **unit**, independent Bernoulli trial: 1 for enter, 0 for not enter.
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Computing Binomial in the limit

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Therefore, in our store footfall example: the probability of k people entering the store in the next 1 hour is:

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Poisson discrete random variable

A Poisson RV X approximates Binomial where n is large, p is small, and $\lambda = np$ is "moderate". Thus we no longer need to know n and p , we only need to provide **rate** λ . X is the number of successes over the duration of the experiment.

$$\mathbf{X} \sim \mathbf{Pois}(\lambda)$$

$$\text{Range: } \{0, 1, 2, \dots\}$$

$$\text{PMF: } \mathbf{P}[X = k] = \frac{\lambda^k}{k!} e^{-\lambda}$$

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Key idea: Divide time into a **large number** of small increments. Assume that during each increment, there is some **small probability** of the event happening (independent of other increments).



Earthquake example

Example

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_____ Answer _____



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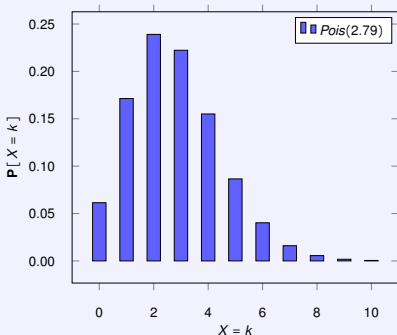
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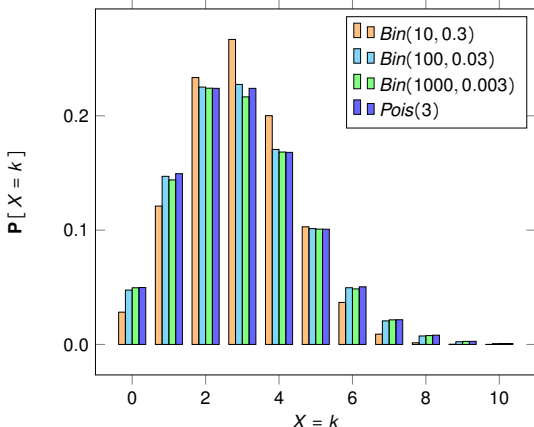
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$$= \lambda \sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} e^{-\lambda} \quad (\text{let } i = k - 1)$$

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$$\mathbf{E}[X] = \sum_{k=0}^{\infty} k \frac{\lambda^k}{k!} e^{-\lambda} = \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!} e^{-\lambda} =$$

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$$= \lambda \sum_{i=0}^{\infty} (i+1) \frac{\lambda^i}{i!} e^{-\lambda} =$$

Poisson variance

$$\begin{aligned}\mathbf{E}[X^2] &= \sum_{k=0}^{\infty} k^2 \frac{\lambda^k}{k!} e^{-\lambda} = \lambda \sum_{k=1}^{\infty} k \frac{\lambda^{k-1}}{(k-1)!} e^{-\lambda} \quad (\text{let } i = k - 1) \\ &= \lambda \sum_{i=0}^{\infty} (i+1) \frac{\lambda^i}{i!} e^{-\lambda} = \lambda \left(\underbrace{\sum_{i=0}^{\infty} i \frac{\lambda^i}{i!} e^{-\lambda}}_{\text{same as before}} + \underbrace{\sum_{i=0}^{\infty} \frac{\lambda^i}{i!} e^{-\lambda}}_{\text{sum of PMFs}=1} \right) =\end{aligned}$$



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$$\mathbf{E}[X^k] = \lambda \mathbf{E}[(X+1)^{k-1}] \quad \leftarrow \text{useful generalisation}$$



Bernoulli, Poisson, and random processes

- A Poisson process is a model for a series of discrete events where the **average time** between events is known, but the exact timing of events is random.



Bernoulli, Poisson, and random processes

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 - Two events cannot occur at the same time: each sub-interval of a Poisson process is a Bernoulli trial that is either a success or a failure.



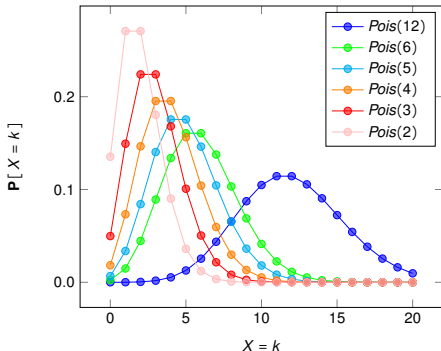
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Outline

Poisson discrete random variable

Geometric discrete random variable

Negative binomial discrete random variable

Hypergeometric discrete random variable



Geometric discrete random variable

X is a geometric RV if X is a number of independent Bernoulli trials until the **first** success, and p is the probability of success on each Bernoulli trial.

$$X \sim \text{Geo}(p)$$

$$\text{Range: } \{1, 2, \dots\}$$

$$\text{PMF: } \mathbf{P}[X = n] = (1 - p)^{n-1} p$$

$$\text{Expectation: } \mathbf{E}[X] = \frac{1}{p}$$

$$\text{Variance: } \mathbf{V}[X] = \frac{1 - p}{p^2}$$



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Examples: tossing a coin ($\mathbf{P}[\text{head}] = p$) until first heads appears, generating bits with $\mathbf{P}[\text{bit} = 1] = p$ until first 1 is generated.



PMF (E_i is the event that the i -th trial succeeds):

$$\mathbf{P}[X = n] = \mathbf{P}[E_1^c E_2^c \dots E_{n-1}^c E_n] =$$

CDF ($\mathbf{P}[X > n]$ is the probability that at least the first n trials fail):



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Die example

Example

You roll a fair 6-sided die until it comes up with #6. What is the probability that it will take 3 rolls?

_____ Answer _____



Die example

Example

You roll a fair 6-sided die until it comes up with #6. What is the probability that it will take 3 rolls?

Answer

Let X be a RV for # of rolls. Probability for any # on die is $\frac{1}{6}$.

Define RVs: $X \sim \text{Geo}(\frac{1}{6})$, want $\mathbf{P}[X = 3]$.

Solve:



Outline

Poisson discrete random variable

Geometric discrete random variable

Negative binomial discrete random variable

Hypergeometric discrete random variable



Negative binomial

Negative binomial discrete random variable

X is a negative binomial RV if X is the number of independent Bernoulli trials until r successes and p is the probability of success on each trial.

$$X \sim \text{NegBin}(r, p)$$

$$\text{Range: } \{r, r + 1, \dots\}$$

$$\text{PMF: } \mathbf{P}[X = n] = \binom{n-1}{r-1} (1-p)^{n-r} p^r$$

$$\text{Expectation: } \mathbf{E}[X] = \frac{r}{p}$$

$$\text{Variance: } \mathbf{V}[X] = \frac{r(1-p)}{p^2}$$



Negative binomial

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$$\text{Expectation: } \mathbf{E}[X] = \frac{r}{p}$$

$$\text{Variance: } \mathbf{V}[X] = \frac{r(1-p)}{p^2}$$

Examples: tossing a coin until r -th heads appears, generating bits until the first r 1's are generated.

Note: $\text{Geo}(p) = \text{NegBin}(1, p)$.



NegBin example

Example (not real life!)

A PhD student is expected to publish 2 papers to graduate. A conference accepts each paper randomly and independently with probability $p = 0.25$. On average, how many papers will the student need to submit to a conference in order to graduate?

_____ Answer _____



Adding NegBin example

Example

Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer



Adding NegBin example

Example

Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.



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Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.
- Consider the sequence of independent events tossing a coin with $\mathbf{P}[\text{heads}] = p$.



Adding NegBin example

Example

Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.
- Consider the sequence of independent events tossing a coin with $\mathbf{P}[\text{heads}] = p$.
- Let X be a RV for # of coin tosses until m heads are observed. Thus $X \sim \text{NegBin}(m, p)$.



Adding NegBin example

Example

Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.
- Consider the sequence of independent events tossing a coin with $\mathbf{P}[\text{heads}] = p$.
- Let X be a RV for # of coin tosses until m heads are observed. Thus $X \sim \text{NegBin}(m, p)$.
- Now, continue to toss a coin after m heads are observed, until n more heads are observed. Thus, for this part of the sequence, $Y \sim \text{NegBin}(n, p)$.



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Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.
- Consider the sequence of independent events tossing a coin with $\mathbf{P}[\text{heads}] = p$.
- Let X be a RV for # of coin tosses until m heads are observed. Thus $X \sim \text{NegBin}(m, p)$.
- Now, continue to toss a coin after m heads are observed, until n more heads are observed. Thus, for this part of the sequence, $Y \sim \text{NegBin}(n, p)$.
- Looking at it from the beginning we tossed independently the coin until we observed $m + n$ heads, thus $Z = X + Y$ and thus $Z \sim \text{NegBin}(m + n, p)$.



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Let $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ be two independent RVs. Define a new RV as $Z = X + Y$. Find PMF of Z .

Answer

- Need to show that $Z \sim \text{NegBin}(m + n, p)$.
- Consider the sequence of independent events tossing a coin with $\mathbf{P}[\text{heads}] = p$.
- Let X be a RV for # of coin tosses until m heads are observed. Thus $X \sim \text{NegBin}(m, p)$.
- Now, continue to toss a coin after m heads are observed, until n more heads are observed. Thus, for this part of the sequence, $Y \sim \text{NegBin}(n, p)$.
- Looking at it from the beginning we tossed independently the coin until we observed $m + n$ heads, thus $Z = X + Y$ and thus $Z \sim \text{NegBin}(m + n, p)$.
- Note: if X_1, X_2, \dots, X_m are m independent $\text{Geo}(p)$ RVs, then the RV $X = X_1 + X_2 + \dots + X_m$ has $\text{NegBin}(m, p)$ distribution.



Outline

Poisson discrete random variable

Geometric discrete random variable

Negative binomial discrete random variable

Hypergeometric discrete random variable



Hypergeometric

Hypergeometric discrete random variable

X is a hypergeometric RV that samples n objects, **without replacement**, with i successes (random draw for which the object drawn has a specified feature), from a finite population of size N that contains exactly m objects with that feature.

$$X \sim \text{Hyp}(N, n, m)$$

$$\text{Range: } \{0, 1, \dots, n\}$$

$$\text{PMF: } \mathbf{P}[X = i] = \frac{\binom{m}{i} \binom{N-m}{n-i}}{\binom{N}{n}}$$

$$\text{Expectation: } \mathbf{E}[X] = n \frac{m}{N}$$

$$\text{Variance: } \mathbf{V}[X] = n \frac{m}{N} \left(1 - \frac{m}{N}\right) \left(1 - \frac{n-1}{N-1}\right)$$

Example: an urn has N balls of which m are white and $N - m$ are black; we take a random sample **without replacement** of size n and measure X : # of white balls in the sample.



Survey sampling

Example

A street has 40 houses of which 5 houses are inhabited by families with an income below the poverty line. In a survey, 7 houses are sampled at random from this street. What is the probability that: (a) none of the 5 families with income below poverty line are sampled? (b) 4 of them are sampled? (c) no more than 2 are sampled? (d) at least 3 are sampled?

Answer



Example

A street has 40 houses of which 5 houses are inhabited by families with an income below the poverty line. In a survey, 7 houses are sampled at random from this street. What is the probability that: (a) none of the 5 families with income below poverty line are sampled? (b) 4 of them are sampled? (c) no more than 2 are sampled? (d) at least 3 are sampled?

Answer

Let X : # of families sampled which are below the poverty line.

$$X \sim \text{Hyp}(N = 40, n = 7, m = 5).$$

Summary of discrete RV

	$Ber(p)$	$Bin(n, p)$	$Pois(\lambda)$	$Geo(p)$	$NegBin(r, p)$	$Hyp(N, n, m)$
PMF	$\mathbf{P}[X=1]=p$	$\mathbf{P}[X=k]=\binom{n}{k}p^k(1-p)^{n-k}$	$\mathbf{P}[X=k]=\frac{\lambda^k}{k!}e^{-\lambda}$	$\mathbf{P}[X=n]=(1-p)^{n-1}p$	$\mathbf{P}[X=n]=\binom{n-1}{r-1}(1-p)^{n-r}p^r$	$\mathbf{P}[X=i]=\frac{\binom{m}{i}\binom{N-m}{n-i}}{\binom{N}{n}}$
$\mathbf{E}[X]$	p	np	λ	$\frac{1}{p}$	$\frac{r}{p}$	$n\frac{m}{N}$
$\mathbf{V}[X]$	$p(1-p)$	$np(1-p)$	λ	$\frac{1-p}{p^2}$	$\frac{r(1-p)}{p^2}$	$n\frac{m}{N}(1-\frac{m}{N})(1-\frac{n-1}{N-1})$
Descr.	1 experiment with prob p of success	n independent trials with prob p of success	# successes over experiment duration, $\lambda = np$ rate of success	# independent trials until first success	# independent trials until r successes	# successes of drawing item with a feature (without replacement) in a sample of size n from a population of size N with m items with the feature