

Introduction to Probability

Lecture 8: Basic Inequalities and Law of Large Numbers

Mateja Jamnik, [Thomas Sauerwald](#)

University of Cambridge, Department of Computer Science and Technology
email: {mateja.jamnik,thomas.sauerwald}@cl.cam.ac.uk

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Outline

Introduction

Markov's Inequality and Chebyshev's Inequality

Weak Law of Large Numbers

Intro: Sum of Independent (Uniform) Random Variables

Example 1

Let X_1 and X_2 be two independent random variables, both uniformly distributed on $[0, 1]$. How does the probability density of $X_1 + X_2$ look like? What happens for $X_1 + X_2 + X_3$ etc.?

Answer

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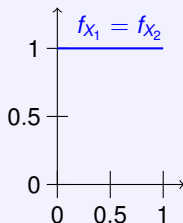
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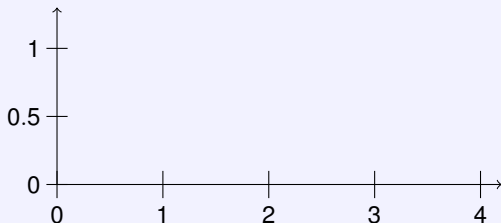
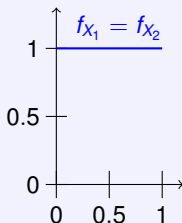
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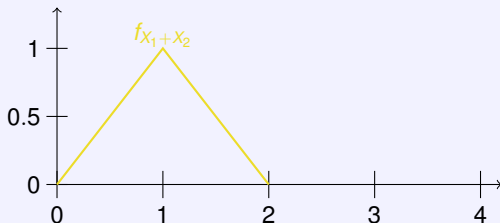
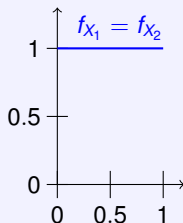
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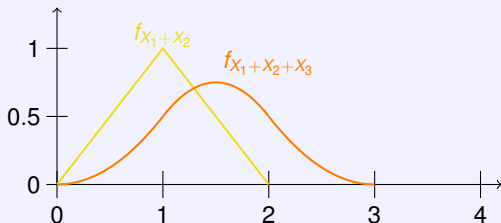
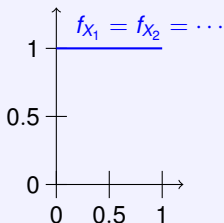
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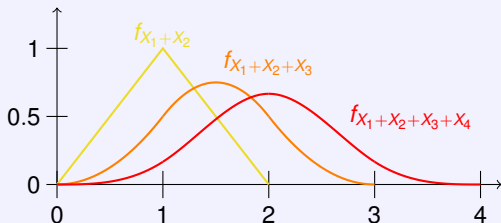
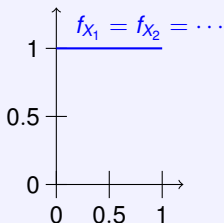
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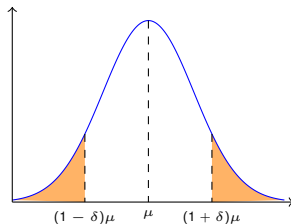
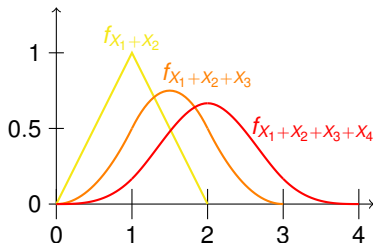
Let us try to sketch the densities without explicit computations^a



^aThis is also called “convolution”. The detailed calculation for $f_{X_1+X_2}$ can be found at the end of these slides. The exact distribution is known for any number of random variables under the name Irwin-Hall distribution.

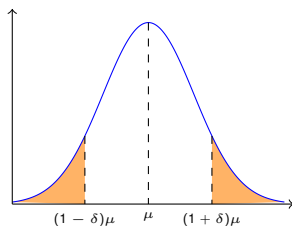
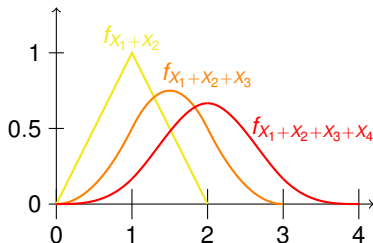
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We will study **sums** of **independent** and **identically distributed** variables. How does their **distribution** look like, and how well do they **concentrate** around the expectation?



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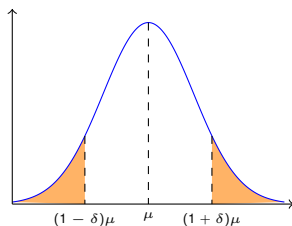
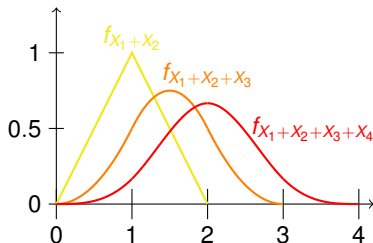
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1. Markov's inequality
2. Chebyshev's inequality
3. Law of Large Numbers
4. **Central Limit Theorem**

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2. Chebyshev's inequality
3. Law of Large Numbers
4. **Central Limit Theorem**

Re-use concepts from previous lectures:

1. Independence (Random Var.) (Lec. 1, 7)
2. Expectation and Variance (Lec. 2, 3)
3. Normal Distribution (Lec. 5)
4. Sums of Random Variables (Lec. 6)

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Weak Law of Large Numbers

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- **Advantage**: Very basic inequality, we only need to know $\mathbf{E}[X]$
- **Downside**: For many distributions, the tail bound might be quite loose
- Proof is similar to the proof of Chebyshev's inequality (Exercise!)

Applying Markov's Inequality

Example 2

Consider throwing an unbiased, six-sided dice 120 times and let X denote the number of times we obtain a six.

1. Derive an upper bound on $\mathbf{P}[X \geq 30]$.
2. Can you also derive an upper bound on $\mathbf{P}[X \leq 10]$?

Answer

Chebyshev's Inequality

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For **any** random variable X with finite $\mathbf{E}[X]$ and $\mathbf{V}[X]$, for any $a > 0$,

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The “ $\mu \pm \text{a few } \sigma$ ” rule. Most of the probability mass is within a few standard deviations from μ .

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- Chebyshev's inequality is also known as **Second Moment Method**

Derivation of Chebychev's inequality

Proof

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Proof

Exercise: Can you find a proof that uses Markov's inequality?

Example: Chebychev is (usually) much stronger than Markov

Example 3

Throw an unbiased coin n times and let X be the total number of heads. In an experiment, with n large, we would usually expect a number of heads that is close to the expectation. Can we justify that?

Answer

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Weak Law of Large Numbers

Law of Large Numbers

The Weak Law of Large Numbers

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- “Power of Averaging”: repeated samples allow us to estimate μ
- A similar statement holds even if the X_i 's are not identically distributed
- There is also a strong law of large numbers:

$$\mathbf{P} \left[\lim_{n \rightarrow \infty} \bar{X}_n = \mu \right] = 1.$$

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How does a “typical” realisation look like?

Illustration of Weak Law of Large Numbers (2/4)

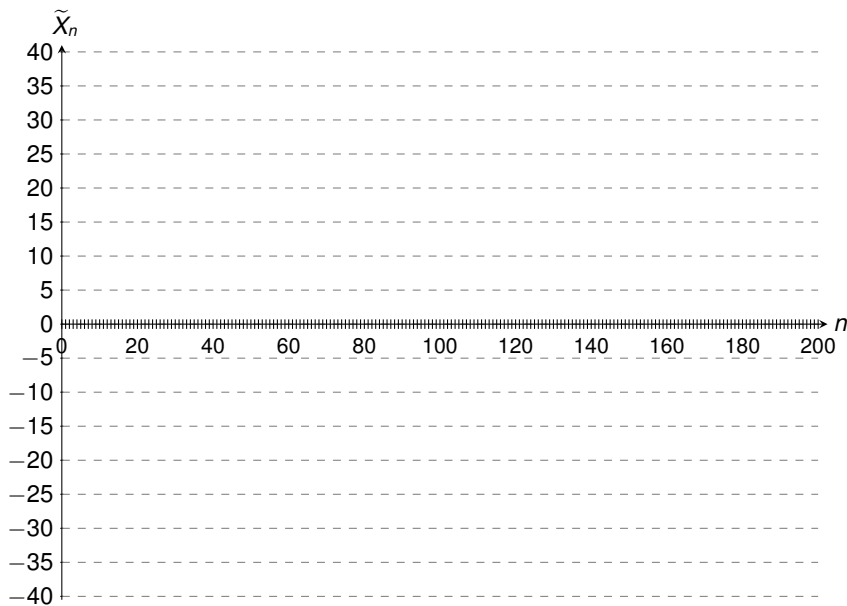


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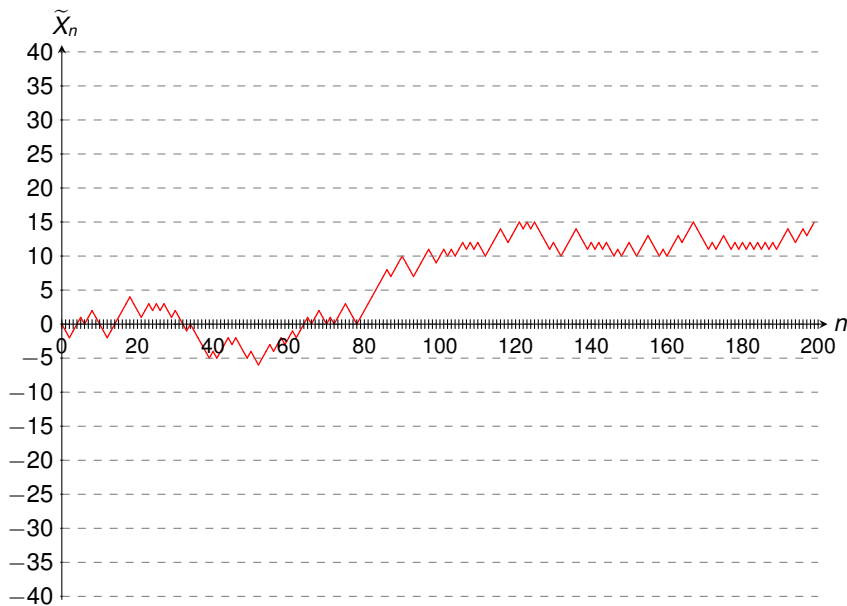


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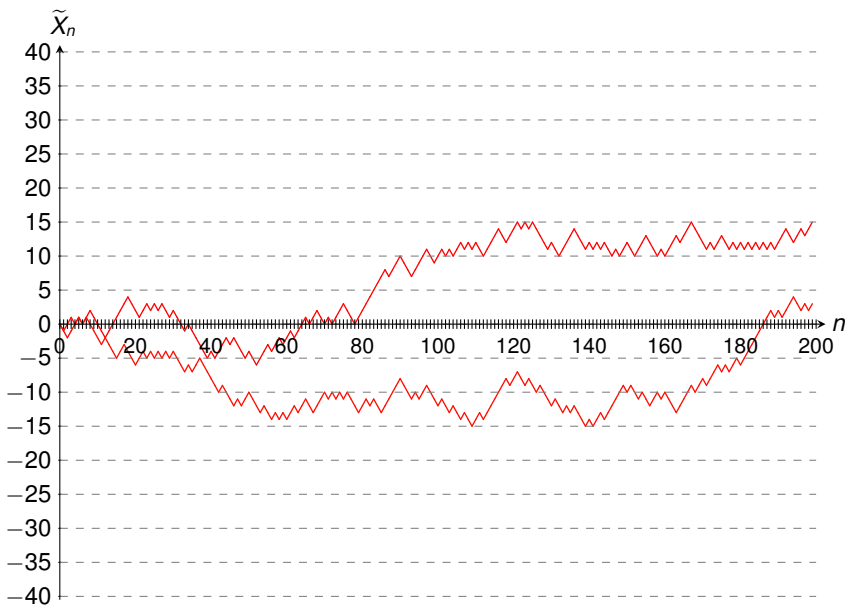


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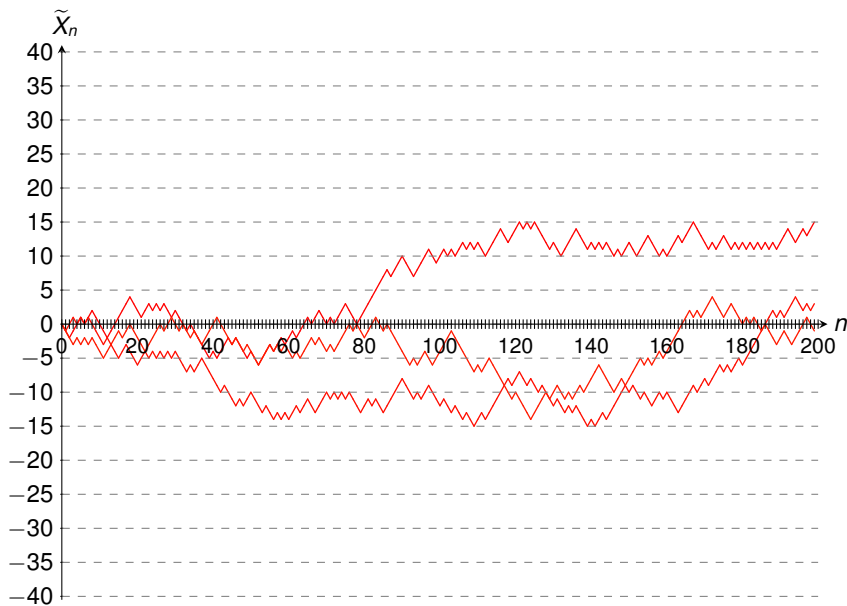


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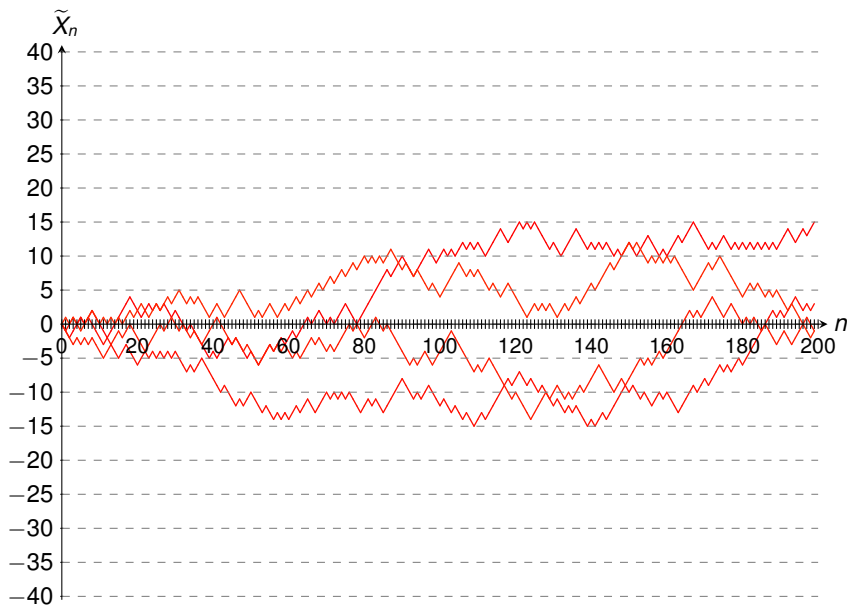


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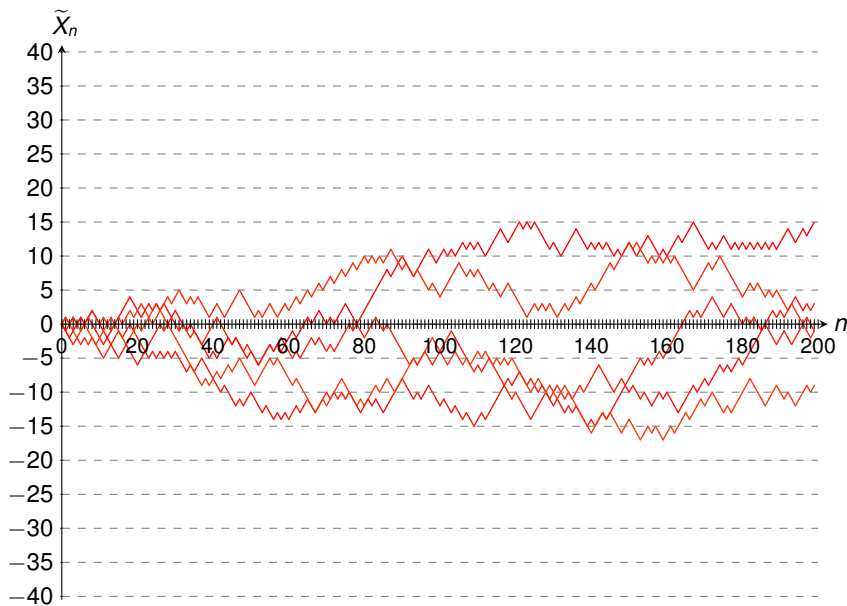


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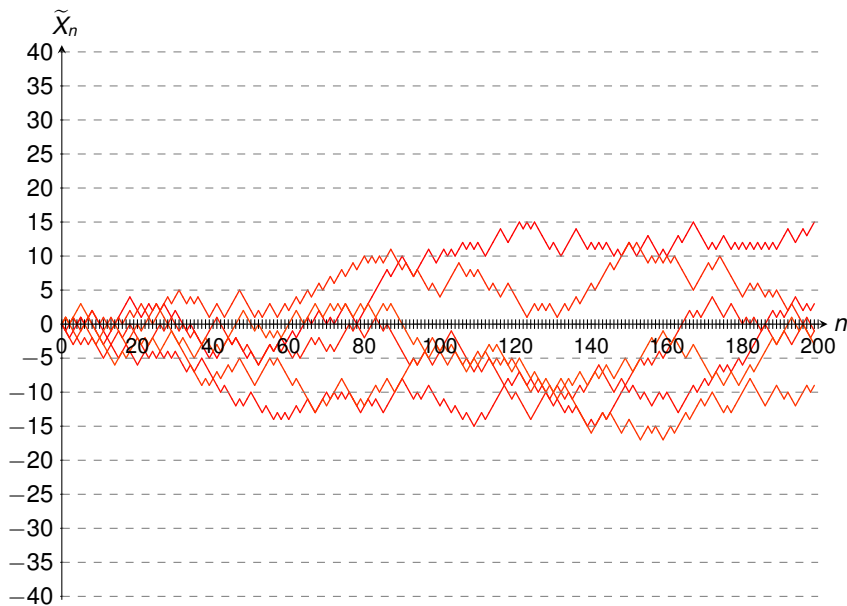


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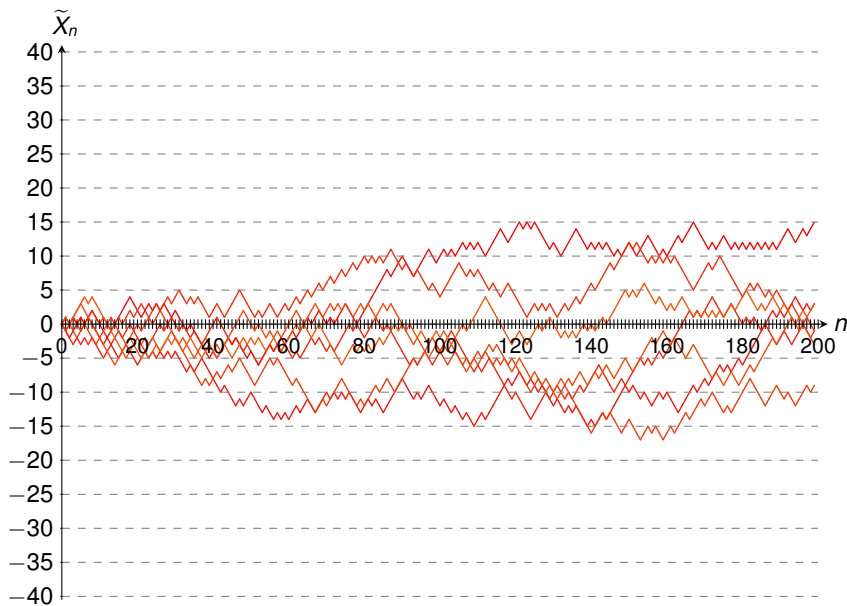


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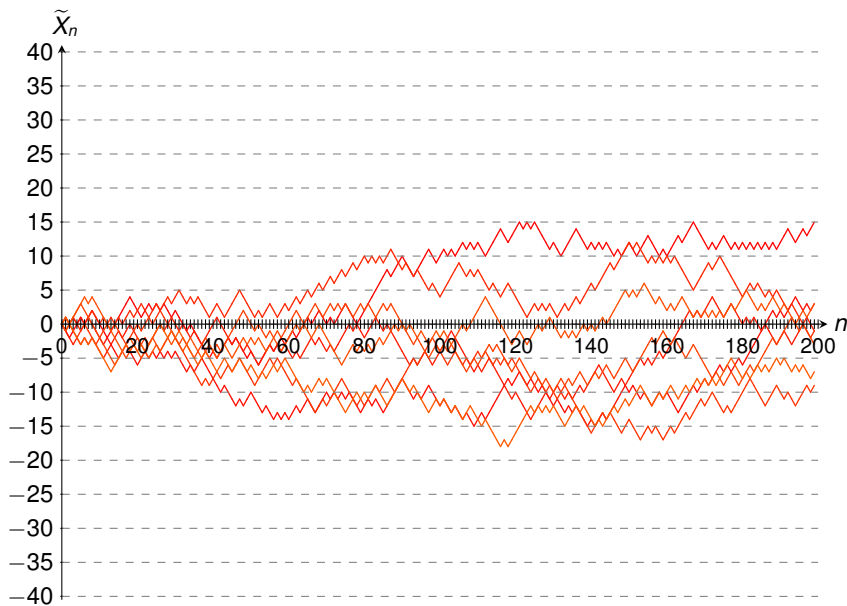


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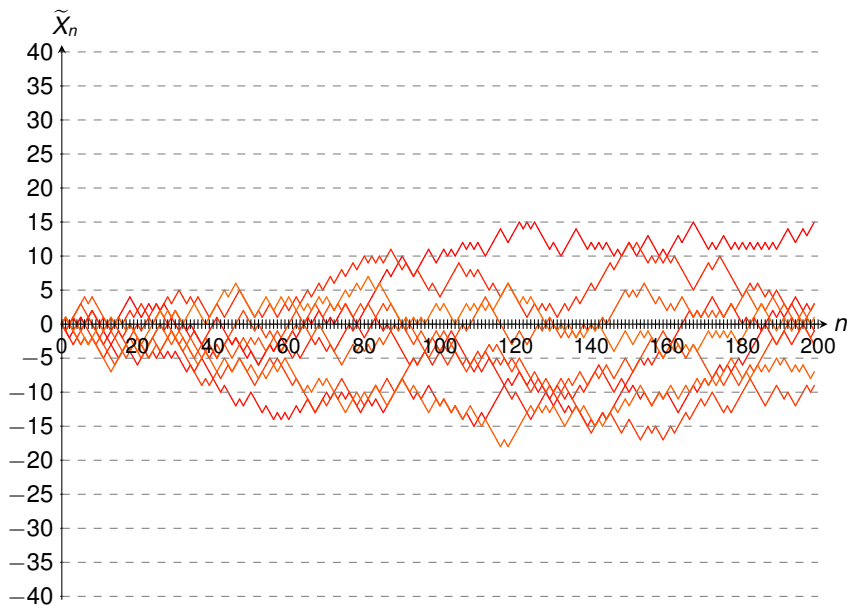


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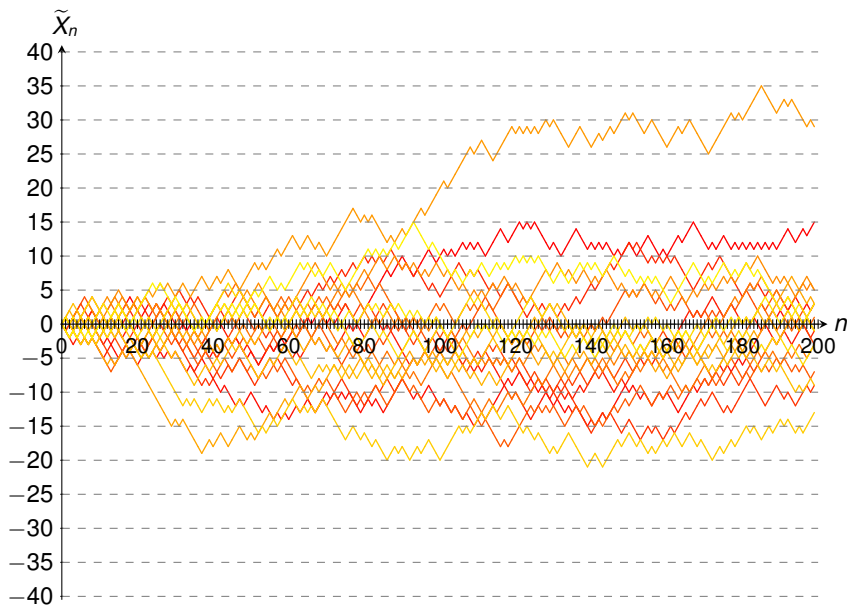
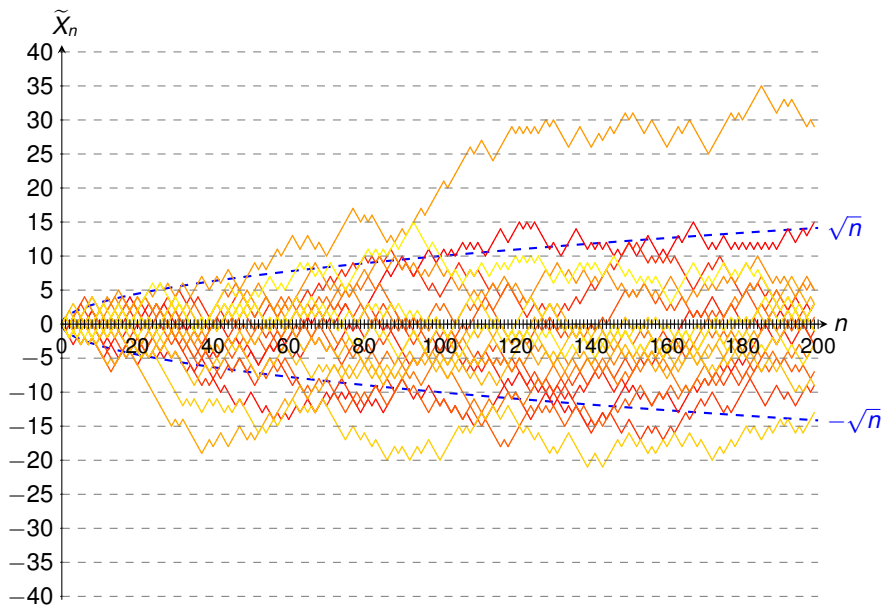
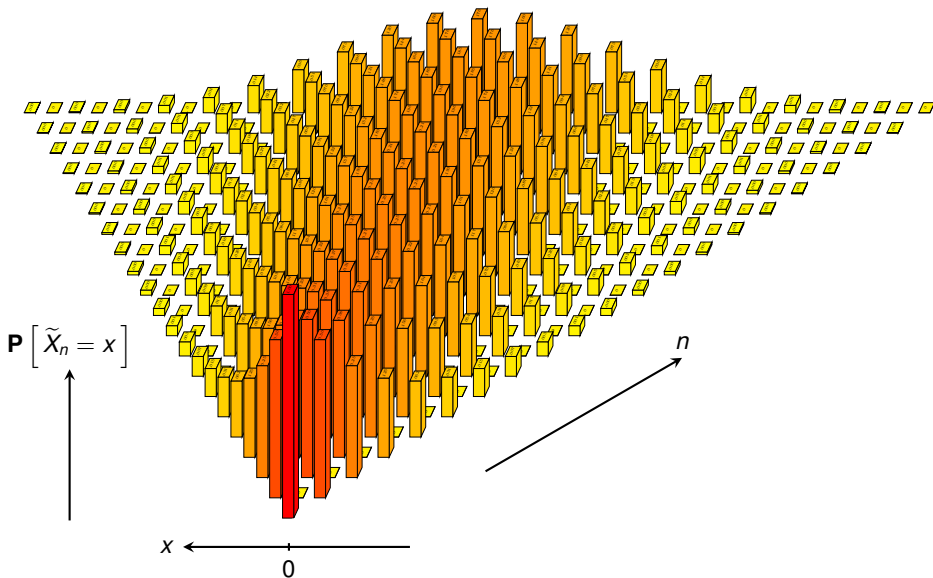


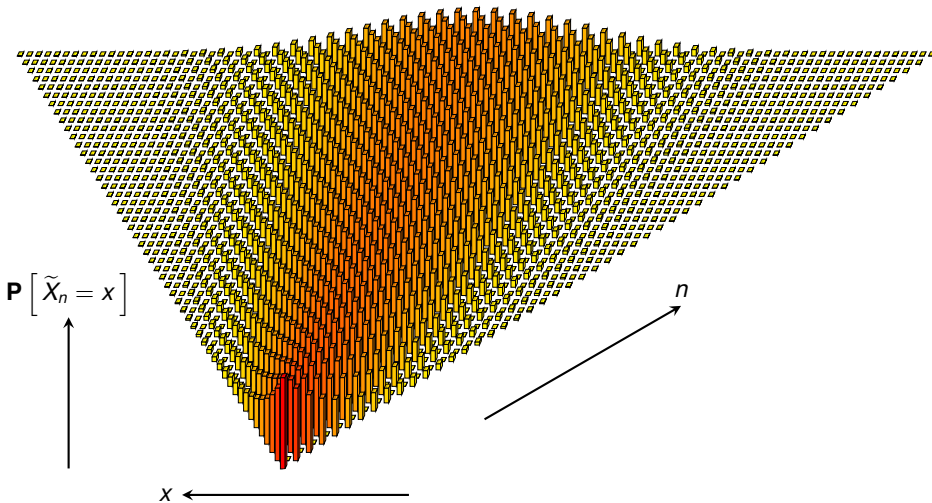
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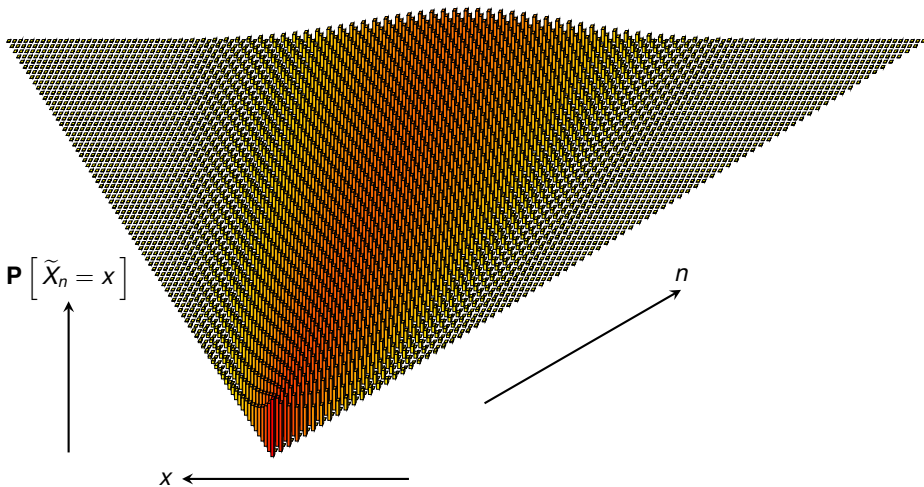
Plot of the Distributions for $n = 0, 1, \dots, 20$



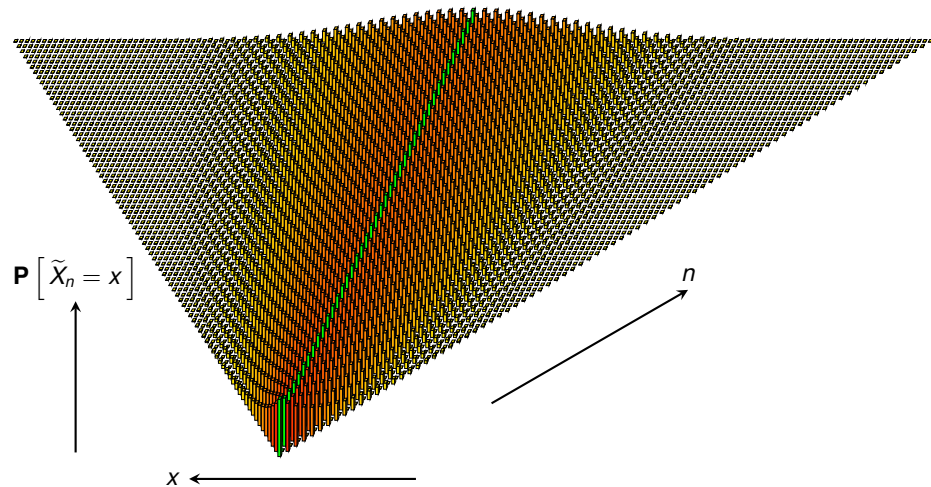
Plot of the Distributions for $n = 0, 1, \dots, 50$



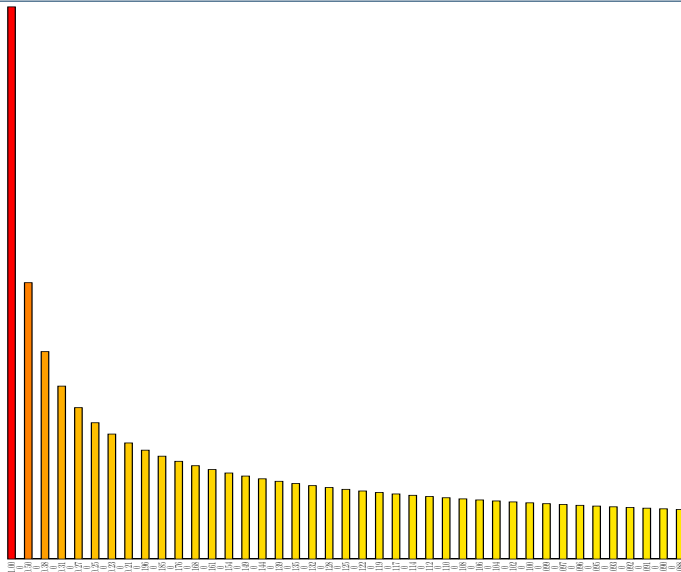
Plot of the Distributions for $n = 0, 1, \dots, 80$



Plot of the Distributions for $n = 0, 1, \dots, 80$



Interlude: Approximation of $P[\tilde{X}_n = 0]$



Interlude: Approximation of $\mathbf{P}[\tilde{X}_n = 0]$

Exercise

Try to find an expression for $\mathbf{P}[\tilde{X}_n = 0]$. Using Stirling's approximation for $n!$, conclude that $\mathbf{P}[\tilde{X}_n = 0] = \Theta(1/\sqrt{n})$ for even integers n .

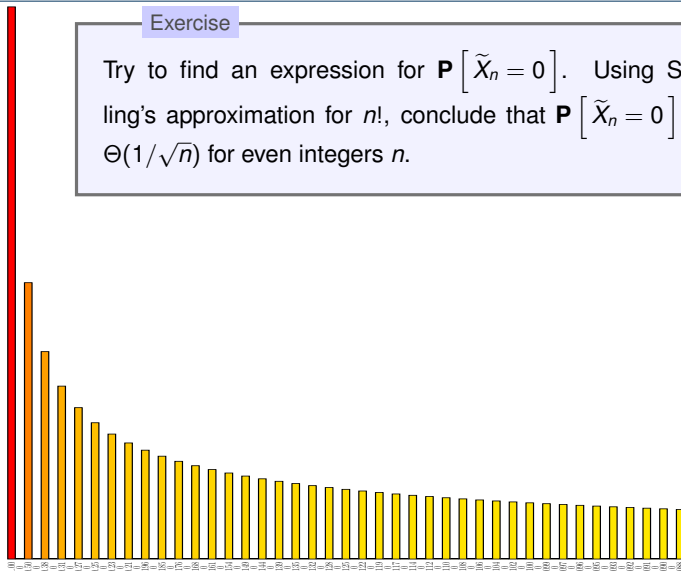


Illustration of Weak Law of Large Numbers (3/4)

- Let X_i be independent random variables taking values $\in \{-1, +1\}$ with probability $1/2$ each
- Consider $\tilde{X}_n := \sum_{i=1}^n X_i$ for any for any $n = 0, 1, \dots, 200$

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Consider now the **average (sample mean)**: $\bar{X}_n := 1/n \cdot \sum_{i=1}^n X_i$.

Illustration of Weak Law of Large Numbers (4/4)

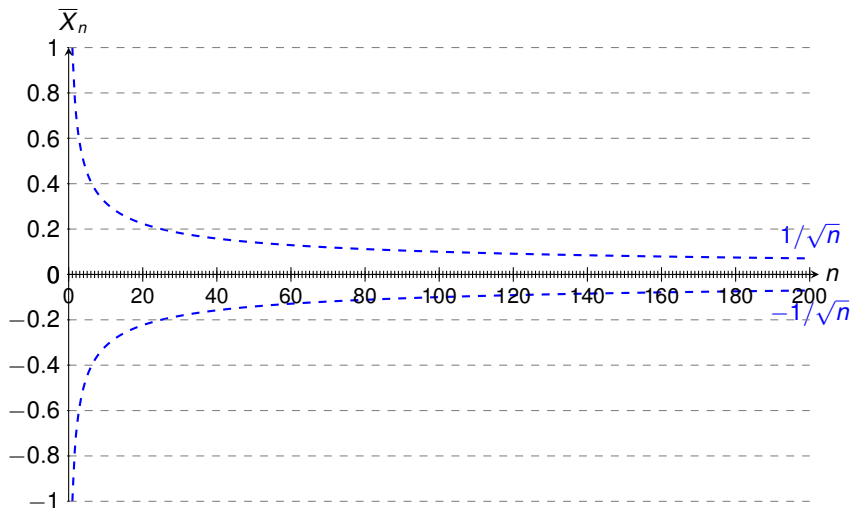


Illustration of Weak Law of Large Numbers (4/4)

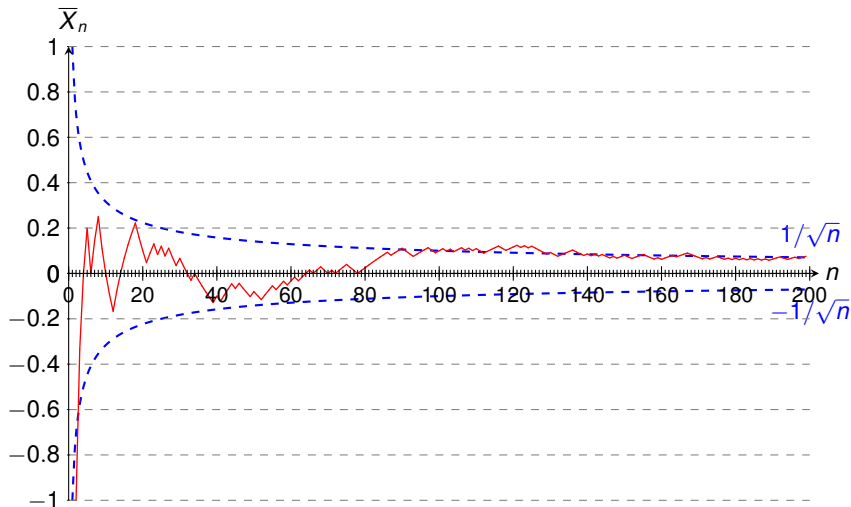


Illustration of Weak Law of Large Numbers (4/4)

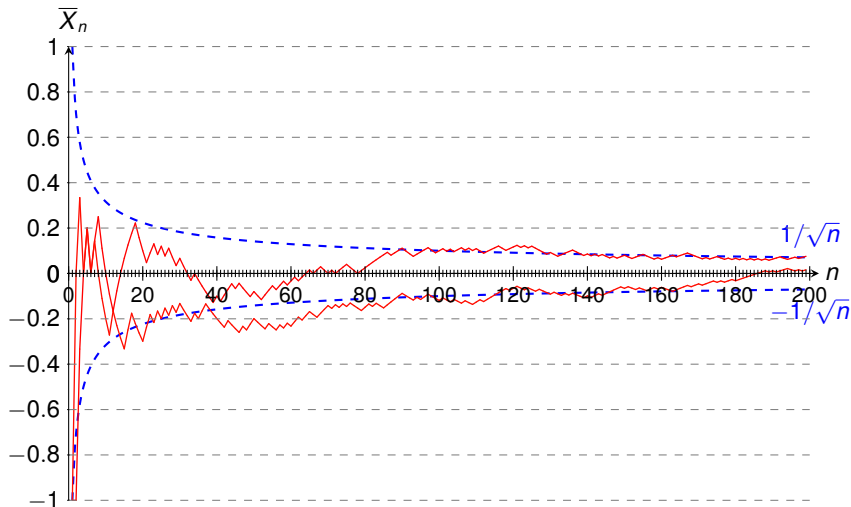


Illustration of Weak Law of Large Numbers (4/4)

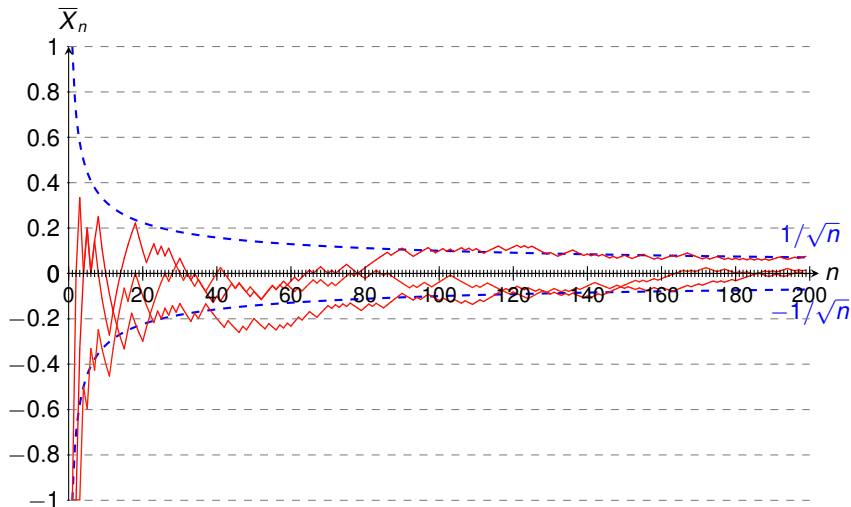


Illustration of Weak Law of Large Numbers (4/4)

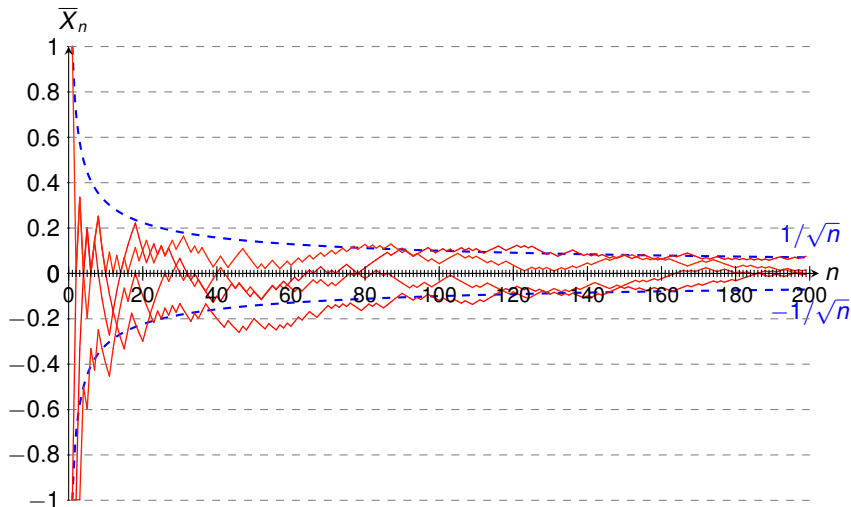
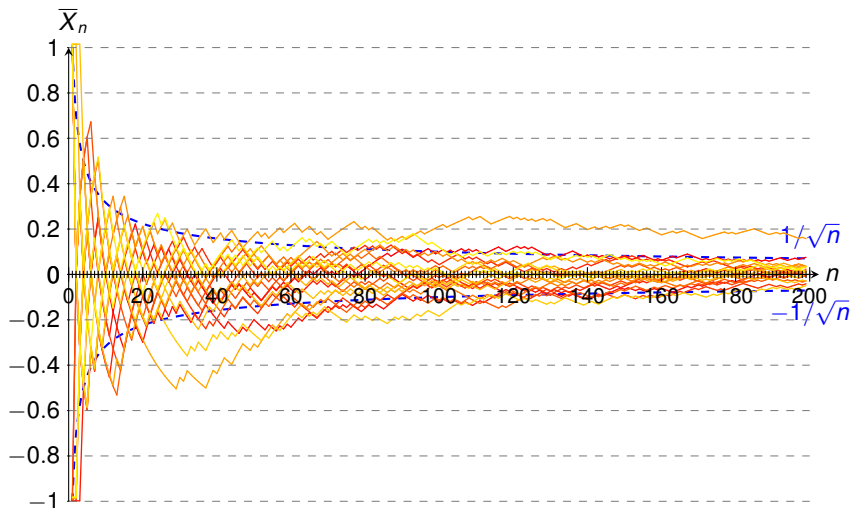


Illustration of Weak Law of Large Numbers (4/4)



Proof of the Weak Law of Large Numbers

The Weak Law of Large Numbers

Let $\bar{X}_n := 1/n \cdot \sum_{i=1}^n X_i$, where the X_i 's are i.i.d. with finite expectation μ and finite variance σ^2 . Then, for any $\epsilon > 0$,

$$\lim_{n \rightarrow \infty} \mathbf{P} \left[|\bar{X}_n - \mu| > \epsilon \right] = 0$$

Proof

Inferring Probabilities of an Event

Example 4

Suppose that, instead of the expectation μ , we want to estimate the probability of an **event**, e.g.,

$$p := \mathbf{P}[X \in (a, b)], \text{ where } a < b.$$

How can we use the **Law of Large Numbers**?

Answer

Appendix: Sum of Two Uniform R.V. (non-examinable)

Example

Let X and Y be two independent random variables, both uniformly distributed on $[0, 1]$. How does the probability density of $X + Y$ look like?

Answer

Appendix: Sum of Two Uniform R.V. (non-examinable)

Example

Let X and Y be two independent random variables, both uniformly distributed on $[0, 1]$. How does the probability density of $X + Y$ look like?

Answer

We have

$$f_{X+Y}(a) \stackrel{(*)}{=} \int_{-\infty}^{+\infty} f_X(a-y)f_Y(y)dy,$$

where for $(*)$, see Chapter 6.3 in Ross (Chapter 11.2 in Dekking et al.). Since $f_Y(y) = 1$ if $0 \leq y \leq 1$ and $f_Y(y) = 0$ otherwise, we have

$$f_{X+Y}(a) = \int_0^1 f_X(a-y)dy.$$

Further, for $0 \leq a \leq 1$ we have $f_X(a-y) = 1$ and $f_X(a-y) = 0$ otherwise, and thus

$$f_{X+Y}(a) = \int_0^a dy = a.$$

Similarly, for $1 < a < 2$, $f_{X+Y}(a) = \int_a^2 dy = 2 - a$. Therefore,

$$f_{X+Y}(a) = \begin{cases} a & \text{if } 0 \leq a \leq 1, \\ 2 - a & \text{if } 1 \leq a \leq 2, \\ 0 & \text{otherwise.} \end{cases}$$