Randomised Algorithms

Lecture 4: Markov Chains and Mixing Times

Thomas Sauerwald (tms41@cam.ac.uk)

Lent 2022



Outline

Recap of Markov Chain Basics

Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

Application 1: Card Shuffling

Application 2: Ehrenfest Chain and Hypercubes

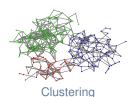
Application 3: Markov Chain Monte Carlo



Broadcasting

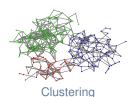


Broadcasting





Broadcasting

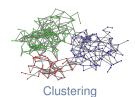




Broadcasting



Ranking Websites

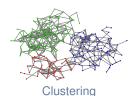




Broadcasting



Ranking Websites



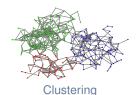
Sampling and Optimisation



Broadcasting



Ranking Websites



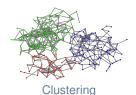
Sampling and Optimisation



Broadcasting



Ranking Websites



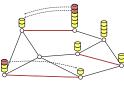
Sampling and Optimisation



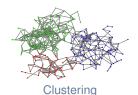
Broadcasting



Ranking Websites



Load Balancing



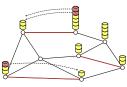
Sampling and Optimisation



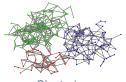
Broadcasting



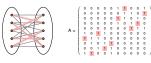
Ranking Websites



Load Balancing



Clustering



Sampling and Optimisation



Particle Processes

- Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

- Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

1. For any $x \in \Omega$, **P** [$X_0 = x$] = $\mu(x)$.

Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

- 1. For any $x \in \Omega$, **P** [$X_0 = x$] = $\mu(x)$.
- 2. The Markov Property holds: for all $t \ge 0$ and any $x_0, \ldots, x_{t+1} \in \Omega$,

$$\mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t, \dots, X_0 = X_0\right] = \mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t\right] \\
:= P(X_t, X_{t+1}).$$

Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

- 1. For any $x \in \Omega$, **P** [$X_0 = x$] = $\mu(x)$.
- 2. The Markov Property holds: for all $t \ge 0$ and any $x_0, \ldots, x_{t+1} \in \Omega$,

$$\mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t, \dots, X_0 = X_0\right] = \mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t\right] \\
:= \mathbf{P}(X_t, X_{t+1}).$$

From the definition one can deduce that (check!)

Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

- 1. For any $x \in \Omega$, **P** [$X_0 = x$] = $\mu(x)$.
- 2. The Markov Property holds: for all $t \ge 0$ and any $x_0, \ldots, x_{t+1} \in \Omega$,

$$\mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t, \dots, X_0 = X_0\right] = \mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t\right] \\
:= \mathbf{P}(X_t, X_{t+1}).$$

From the definition one can deduce that (check!)

• For all $t, x_0, x_1, \ldots, x_t \in \Omega$,

$$P[X_t = x_t, X_{t-1} = x_{t-1}, \dots, X_0 = x_0]$$

= $\mu(x_0) \cdot P(x_0, x_1) \cdot \dots \cdot P(x_{t-2}, x_{t-1}) \cdot P(x_{t-1}, x_t).$

Markov Chain (Discrete Time and State, Time Homogeneous) -

We say that $(X_t)_{t=0}^{\infty}$ is a Markov Chain on State Space Ω with Initial Distribution μ and Transition Matrix P if:

- 1. For any $x \in \Omega$, **P** [$X_0 = x$] = $\mu(x)$.
- 2. The Markov Property holds: for all $t \ge 0$ and any $x_0, \ldots, x_{t+1} \in \Omega$,

$$\mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t, \dots, X_0 = X_0\right] = \mathbf{P}\left[X_{t+1} = X_{t+1} \mid X_t = X_t\right] \\
:= P(X_t, X_{t+1}).$$

From the definition one can deduce that (check!)

• For all $t, x_0, x_1, \ldots, x_t \in \Omega$,

$$P[X_t = x_t, X_{t-1} = x_{t-1}, \dots, X_0 = x_0]$$

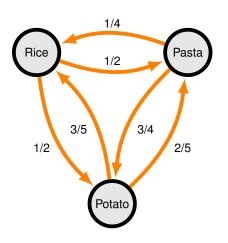
= $\mu(x_0) \cdot P(x_0, x_1) \cdot \dots \cdot P(x_{t-2}, x_{t-1}) \cdot P(x_{t-1}, x_t).$

• For all $0 < t_1 < t_2, x \in \Omega$,

$$\mathbf{P}[X_{t_2} = x] = \sum_{y \in \Omega} \mathbf{P}[X_{t_2} = x \mid X_{t_1} = y] \cdot \mathbf{P}[X_{t_1} = y].$$

What does a Markov Chain Look Like?

Example: the carbohydrate served with lunch in the college cafeteria.



This has transition matrix:

$$P = \begin{bmatrix} \text{Rice} & \text{Pasta} & \text{Potato} \\ 0 & 1/2 & 1/2 \\ 1/4 & 0 & 3/4 \\ 3/5 & 2/5 & 0 \end{bmatrix} \begin{array}{c} \text{Rice} \\ \text{Pasta} \\ \text{Potato} \\ \text{Potato} \\ \end{array}$$



The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

• $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

- $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).
- Multiplying ρ^t by P corresponds to advancing the chain one step:

$$\rho^t(y) = \sum_{i \in \mathcal{O}} \rho^{t-1}(x) \cdot P(x, y)$$
 and thus $\rho^t = \rho^{t-1} \cdot P$.

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

- $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).
- Multiplying ρ^t by P corresponds to advancing the chain one step:

$$ho^t(y) = \sum_{i \in \Omega}
ho^{t-1}(x) \cdot P(x,y)$$
 and thus $ho^t =
ho^{t-1} \cdot P$.

• The Markov Property and line above imply that for any $t \ge 0$

$$\rho^t = \rho \cdot P^{t-1}$$
 and thus $P^t(x, y) = \mathbf{P}[X_t = y \mid X_0 = x].$

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

- $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).
- Multiplying ρ^t by P corresponds to advancing the chain one step:

$$ho^t(y) = \sum_{i \in \Omega}
ho^{t-1}(x) \cdot P(x, y)$$
 and thus $ho^t =
ho^{t-1} \cdot P$.

• The Markov Property and line above imply that for any $t \ge 0$

$$ho^t =
ho \cdot P^{t-1}$$
 and thus $P^t(x,y) = \mathbf{P}[X_t = y \mid X_0 = x].$

Thus
$$\rho^t(x) = (\mu P^t)(x)$$
 and so $\rho^t = \mu P^t = (\mu P^t(1), \mu P^t(2), \dots, \mu P^t(n))$.

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

- $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).
- Multiplying ρ^t by P corresponds to advancing the chain one step:

$$ho^t(y) = \sum_{i \in \mathcal{O}}
ho^{t-1}(x) \cdot P(x, y)$$
 and thus $ho^t =
ho^{t-1} \cdot P$.

■ The Markov Property and line above imply that for any $t \ge 0$

$$\rho^t = \rho \cdot P^{t-1} \quad \text{and thus} \quad P^t(x,y) = \mathbf{P} \left[X_t = y \mid X_0 = x \right].$$
 Thus $\rho^t(x) = (\mu P^t)(x)$ and so $\rho^t = \mu P^t = (\mu P^t(1), \mu P^t(2), \dots, \mu P^t(n)).$

Everything boils down to deterministic vector/matrix computations

The Transition Matrix P of a Markov chain (μ, P) on $\Omega = \{1, \dots n\}$ is given by

$$P = \begin{pmatrix} P(1,1) & \dots & P(1,n) \\ \vdots & \ddots & \vdots \\ P(n,1) & \dots & P(n,n) \end{pmatrix}.$$

- $\rho^t = (\rho^t(1), \rho^t(2), \dots, \rho^t(n))$: state vector at time t (row vector).
- Multiplying ρ^t by P corresponds to advancing the chain one step:

$$\rho^t(y) = \sum_{i \in \mathcal{O}} \rho^{t-1}(x) \cdot P(x, y)$$
 and thus $\rho^t = \rho^{t-1} \cdot P$.

■ The Markov Property and line above imply that for any $t \ge 0$

$$ho^t =
ho \cdot P^{t-1}$$
 and thus $P^t(x,y) = \mathbf{P}[X_t = y \mid X_0 = x]$.
Thus $ho^t(x) = (\mu P^t)(x)$ and so $ho^t = \mu P^t = (\mu P^t(1), \mu P^t(2), \dots, \mu P^t(n))$.

- Everything boils down to deterministic vector/matrix computations
- \Rightarrow can replace ρ by any (load) vector and view P as a balancing matrix!

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

√ "We had rice yesterday"

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

```
✓ "We had rice yesterday" \rightarrow \tau := \min\{t \ge 1 : X_{t-1} = \text{"rice"}\}
```

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

- ✓ "We had rice yesterday" $\rightarrow \tau := \min\{t \ge 1 : X_{t-1} = \text{"rice"}\}$
- "We are having pasta next Thursday"

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

- ✓ "We had rice yesterday" \sim $\tau := \min\{t \ge 1 : X_{t-1} = \text{"rice"}\}$
- × "We are having pasta next Thursday"

For two states $x, y \in \Omega$ we call h(x, y) the hitting time of y from x:

$$h(x, y) := \mathbf{E}_x[\tau_y] = \mathbf{E}[\tau_y \mid X_0 = x]$$
 where $\tau_y = \min\{t \ge 1 : X_t = y\}$.

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

- ✓ "We had rice yesterday" $\rightarrow \tau := \min\{t \ge 1 : X_{t-1} = \text{"rice"}\}$
- × "We are having pasta next Thursday"

For two states $x, y \in \Omega$ we call h(x, y) the hitting time of y from x:

$$\textit{h}(\textit{x},\textit{y}) := \textbf{E}_{\textit{x}}[\tau_{\textit{y}}] = \textbf{E}\left[\tau_{\textit{y}} \mid \textit{X}_{0} = \textit{x}\right] \quad \text{where } \tau_{\textit{y}} = \min\{t \geq 1 : \textit{X}_{t} = \textit{y}\}.$$

Some distinguish between
$$\tau_y^+ = \min\{t \ge 1 \colon X_t = y\}$$
 and $\tau_y = \min\{t \ge 0 \colon X_t = y\}$

A non-negative integer random variable τ is a stopping time for $(X_t)_{t\geq 0}$ if for every $s\geq 0$ the event $\{\tau=s\}$ depends only on X_0,\ldots,X_s .

Example - College Carbs Stopping times:

- ✓ "We had rice yesterday" $\rightarrow \tau := \min\{t \ge 1 : X_{t-1} = \text{"rice"}\}$
- × "We are having pasta next Thursday"

For two states $x, y \in \Omega$ we call h(x, y) the hitting time of y from x:

$$h(x, y) := \mathbf{E}_x[\tau_y] = \mathbf{E}[\tau_y \mid X_0 = x]$$
 where $\tau_y = \min\{t \ge 1 : X_t = y\}$.

A Useful Identity ———

Hitting times are the solution to a set of linear equations:

$$h(x,y) \stackrel{\mathsf{Markov} \ \mathsf{Prop.}}{=} 1 + \sum_{z \in \Omega \setminus \{y\}} P(x,z) \cdot h(z,y) \qquad \forall x \neq y \in \Omega.$$

Outline

Recap of Markov Chain Basics

Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

Application 1: Card Shuffling

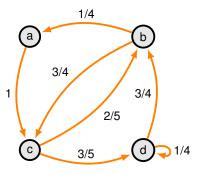
Application 2: Ehrenfest Chain and Hypercubes

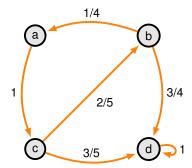
Application 3: Markov Chain Monte Carlo

Irreducible Markov Chains

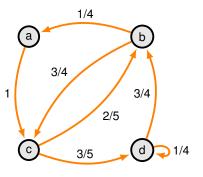
A Markov Chain is irreducible if for every state $x \in \Omega$ there is an integer $k \ge 0$ such that $P^k(x,x) > 0$.

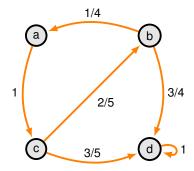
A Markov Chain is irreducible if for every state $x \in \Omega$ there is an integer $k \ge 0$ such that $P^k(x,x) > 0$.





A Markov Chain is irreducible if for every state $x \in \Omega$ there is an integer $k \ge 0$ such that $P^k(x,x) > 0$.

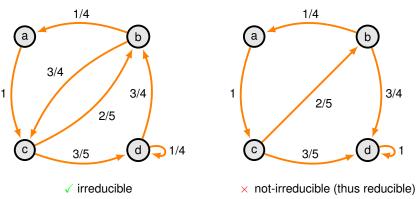






Exercise: Which of the two chains (if any) are irreducible?

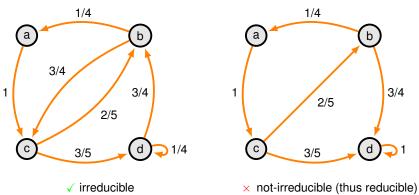
A Markov Chain is irreducible if for every state $x \in \Omega$ there is an integer $k \ge 0$ such that $P^k(x,x) > 0$.





Exercise: Which of the two chains (if any) are irreducible?

A Markov Chain is irreducible if for every state $x \in \Omega$ there is an integer $k \ge 0$ such that $P^k(x,x) > 0$.



Finite Hitting Time Theorem -

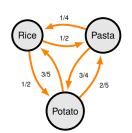
For any states x and y of a finite irreducible Markov Chain $h(x, y) < \infty$.

A probability distribution $\pi = (\pi(1), \dots, \pi(n))$ is the stationary distribution of a Markov Chain if $\pi P = \pi$ (π is a left eigenvector with eigenvalue 1)

A probability distribution $\pi = (\pi(1), \dots, \pi(n))$ is the stationary distribution of a Markov Chain if $\pi P = \pi$ (π is a left eigenvector with eigenvalue 1)

College carbs example:

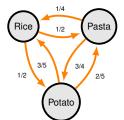
$$\left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right) \cdot \begin{pmatrix} 0 & 1/2 & 1/2 \\ 1/4 & 0 & 3/4 \\ 3/5 & 2/5 & 0 \end{pmatrix} = \left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right)$$



A probability distribution $\pi = (\pi(1), \dots, \pi(n))$ is the stationary distribution of a Markov Chain if $\pi P = \pi$ (π is a left eigenvector with eigenvalue 1)

College carbs example:

$$\begin{pmatrix} \frac{4}{13}, \frac{4}{13}, \frac{5}{13} \end{pmatrix} \cdot \begin{pmatrix} 0 & 1/2 & 1/2 \\ 1/4 & 0 & 3/4 \\ 3/5 & 2/5 & 0 \end{pmatrix} = \begin{pmatrix} \frac{4}{13}, \frac{4}{13}, \frac{5}{13} \end{pmatrix}$$

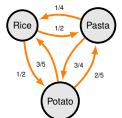


• A Markov Chain reaches stationary distribution if $\rho^t = \pi$ for some t.

A probability distribution $\pi = (\pi(1), \dots, \pi(n))$ is the stationary distribution of a Markov Chain if $\pi P = \pi$ (π is a left eigenvector with eigenvalue 1)

College carbs example:

$$\left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right) \cdot \begin{pmatrix} 0 & 1/2 & 1/2 \\ 1/4 & 0 & 3/4 \\ 3/5 & 2/5 & 0 \end{pmatrix} = \left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right)$$

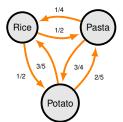


- A Markov Chain reaches stationary distribution if $\rho^t = \pi$ for some t.
- If reached, then it persists: If $\rho^t = \pi$ then $\rho^{t+k} = \pi$ for all $k \ge 0$.

A probability distribution $\pi = (\pi(1), \dots, \pi(n))$ is the stationary distribution of a Markov Chain if $\pi P = \pi$ (π is a left eigenvector with eigenvalue 1)

College carbs example:

$$\left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right) \cdot \begin{pmatrix} 0 & 1/2 & 1/2 \\ 1/4 & 0 & 3/4 \\ 3/5 & 2/5 & 0 \end{pmatrix} = \left(\frac{4}{13}, \frac{4}{13}, \frac{5}{13}\right)$$
Hice



- A Markov Chain reaches stationary distribution if $\rho^t = \pi$ for some t.
- If reached, then it persists: If $\rho^t = \pi$ then $\rho^{t+k} = \pi$ for all $k \ge 0$.

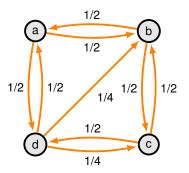
Existence and Uniqueness of a Positive Stationary Distribution -

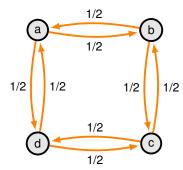
Let *P* be finite, irreducible M.C., then there exists a unique probability distribution π on Ω such that $\pi = \pi P$ and $\pi(x) = 1/h(x, x) > 0$, $\forall x \in \Omega$.

• A Markov Chain is aperiodic if for all $x \in \Omega$, $gcd\{t \ge 1 : P_{x,x}^t > 0\} = 1$.

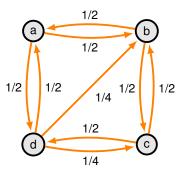
- A Markov Chain is aperiodic if for all $x \in \Omega$, $gcd\{t \ge 1 : P_{x,x}^t > 0\} = 1$.
- Otherwise we say it is periodic.

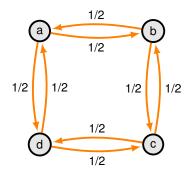
- A Markov Chain is aperiodic if for all $x \in \Omega$, $gcd\{t \ge 1 : P_{x,x}^t > 0\} = 1$.
- Otherwise we say it is periodic.





- A Markov Chain is aperiodic if for all $x \in \Omega$, $gcd\{t \ge 1 : P_{x,x}^t > 0\} = 1$.
- Otherwise we say it is periodic.

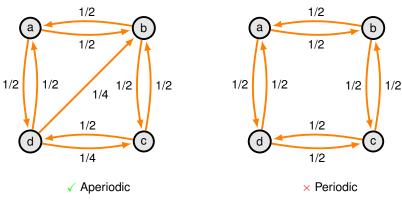






Exercise: Which of the two chains (if any) are aperiodic?

- A Markov Chain is aperiodic if for all $x \in \Omega$, $gcd\{t \ge 1 : P_{x,x}^t > 0\} = 1$.
- Otherwise we say it is periodic.





Exercise: Which of the two chains (if any) are aperiodic?

Convergence Theorem -

Let P be any finite, irreducible, aperiodic Markov Chain with stationary distribution π . Then for any $x, y \in \Omega$,

$$\lim_{t\to\infty} P^t_{x,y} = \pi_y.$$

 ${\sf Ergodic} = {\sf Irreducible} + {\sf Aperiodic}$

Convergence Theorem

Let P be any finite, irreducible, aperiodic Markov Chain with stationary distribution π . Then for any $x,y\in\Omega$,

$$\lim_{t\to\infty} P_{x,y}^t = \pi_y.$$

 ${\sf Ergodic} = {\sf Irreducible} + {\sf Aperiodic}$

Convergence Theorem

Let P be any finite, irreducible, aperiodic Markov Chain with stationary distribution π . Then for any $x,y\in\Omega$,

$$\lim_{t\to\infty} P_{x,y}^t = \pi_y.$$

• mentioned before: For finite irreducible M.C.'s π exists, is unique and

$$\pi_y=\frac{1}{h(y,y)}>0.$$

Ergodic = Irreducible + Aperiodic

Convergence Theorem

Let P be any finite, irreducible, aperiodic Markov Chain with stationary distribution π . Then for any $x,y\in\Omega$,

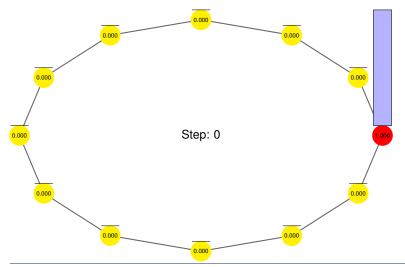
$$\lim_{t\to\infty} P^t_{x,y} = \pi_y.$$

• mentioned before: For finite irreducible M.C.'s π exists, is unique and

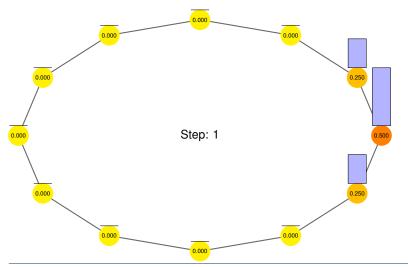
$$\pi_y=\frac{1}{h(y,y)}>0.$$

 We will prove a simpler version of the Convergence Theorem after introducing Spectral Graph Theory.

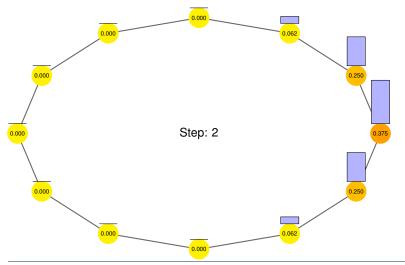
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



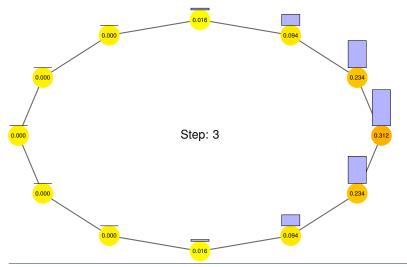
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, \dots, 12\}$ is $P^t(1, x)$.



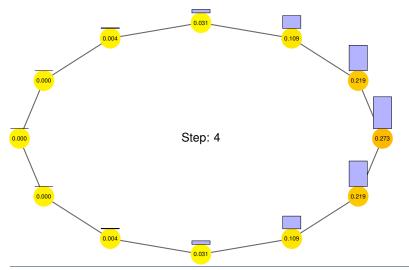
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, \dots, 12\}$ is $P^t(1, x)$.



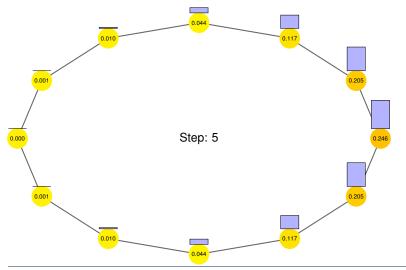
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



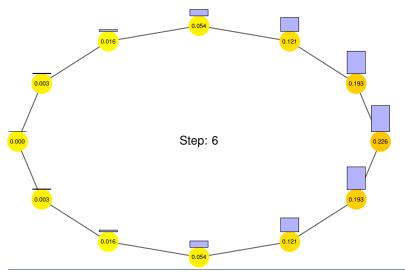
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



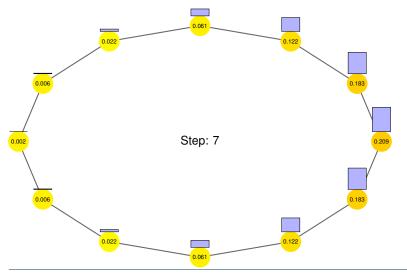
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



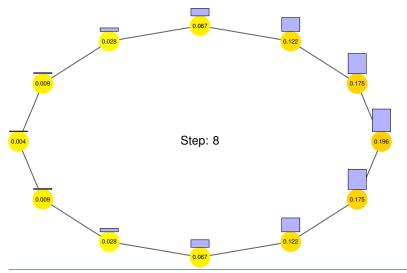
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



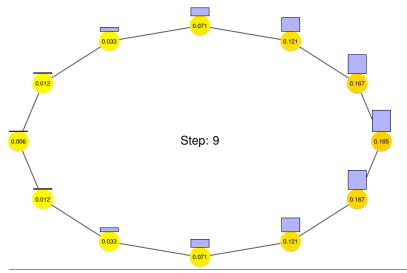
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



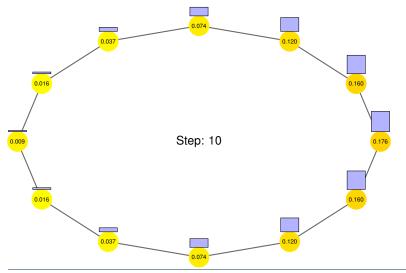
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



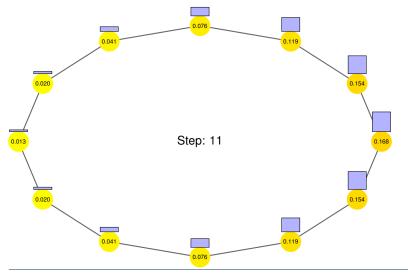
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



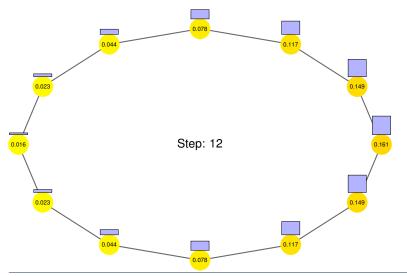
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



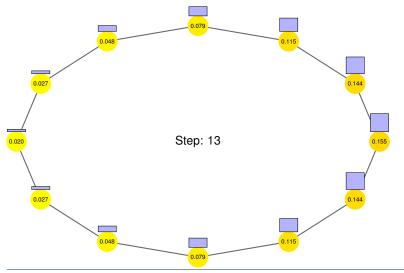
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



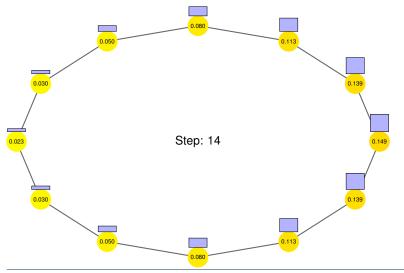
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



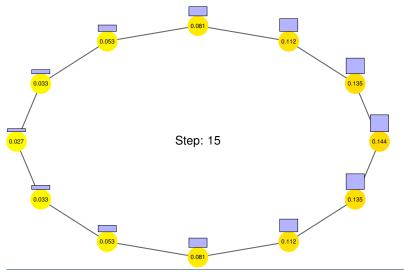
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



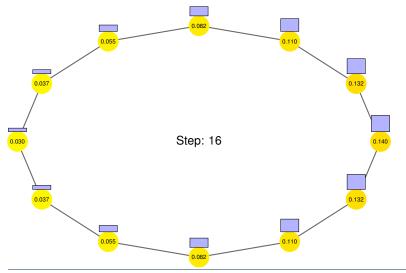
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



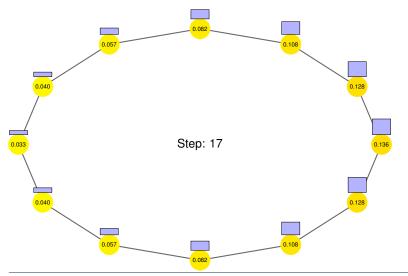
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



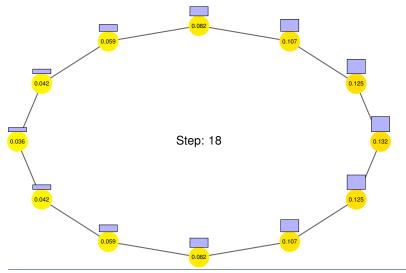
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



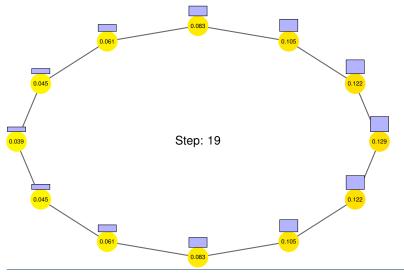
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



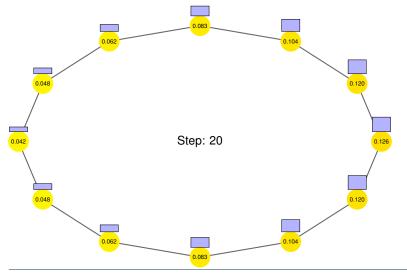
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, \dots, 12\}$ is $P^t(1, x)$.



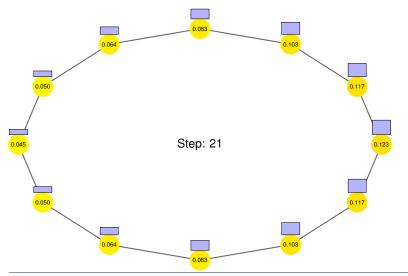
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



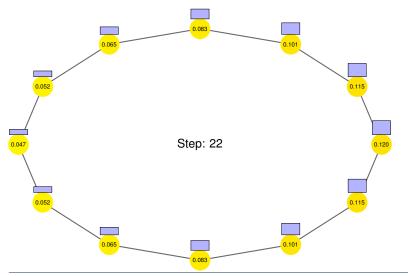
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



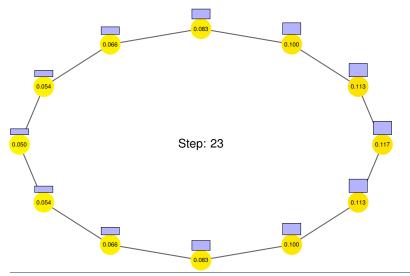
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



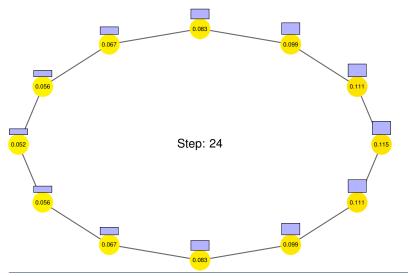
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



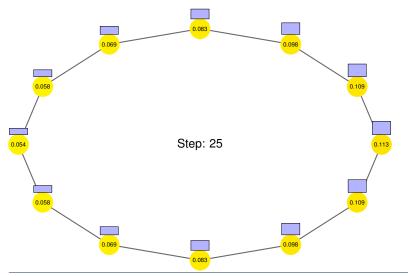
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



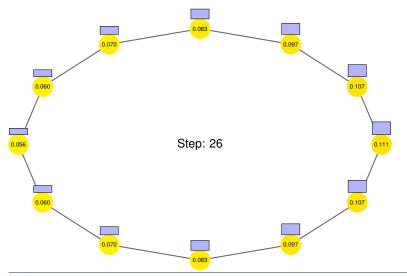
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



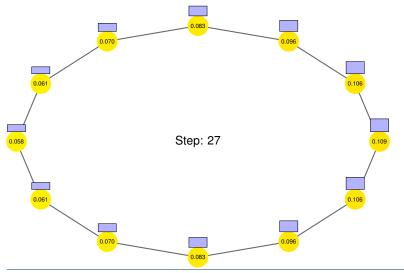
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



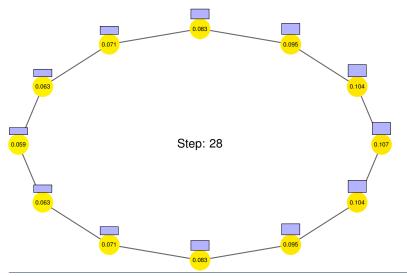
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



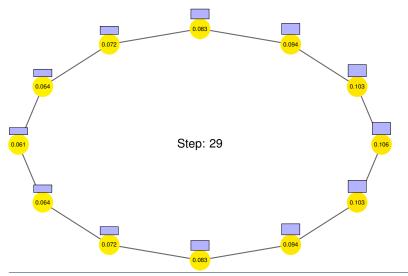
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



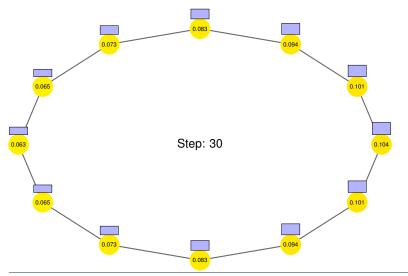
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



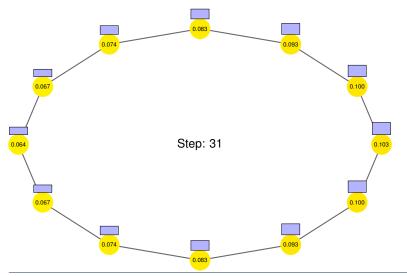
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



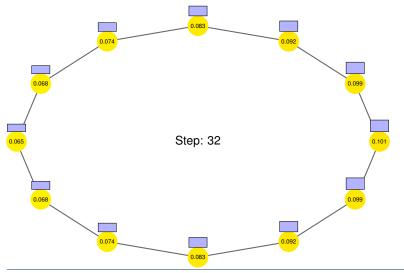
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



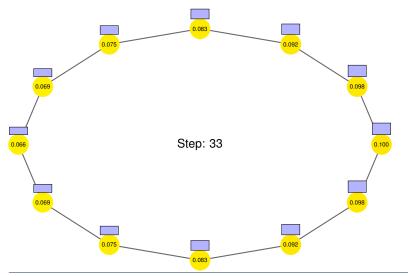
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



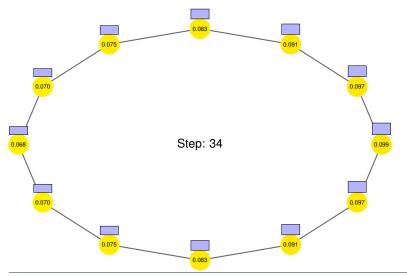
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



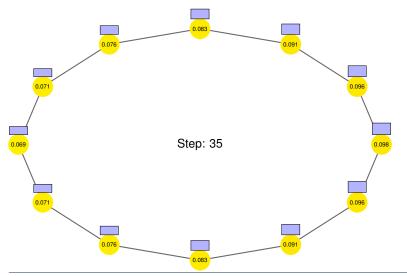
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



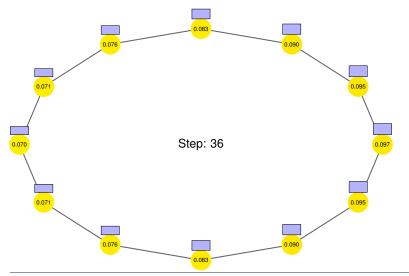
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



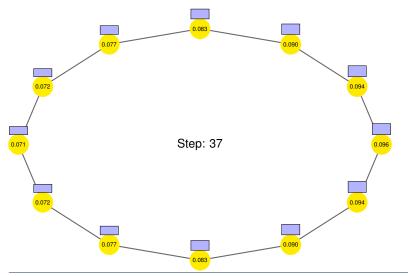
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



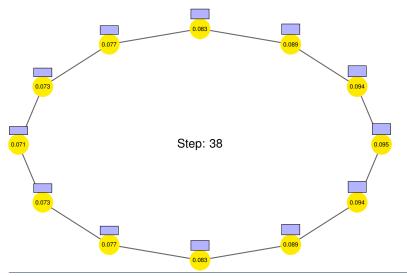
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



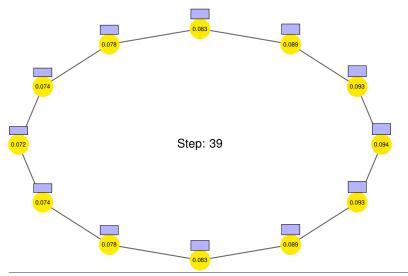
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



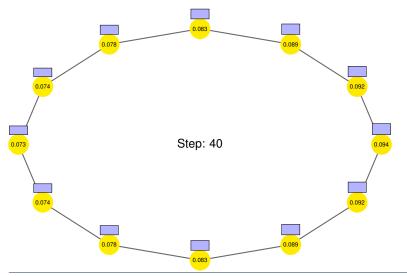
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



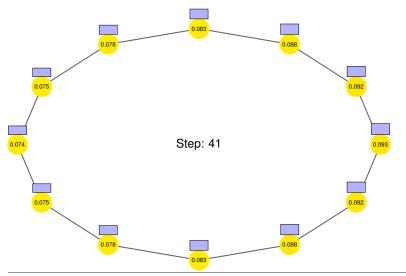
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



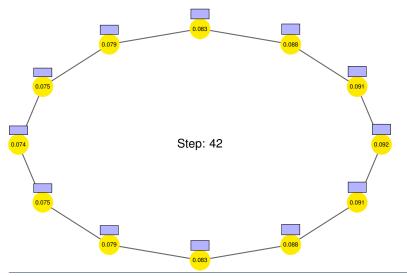
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



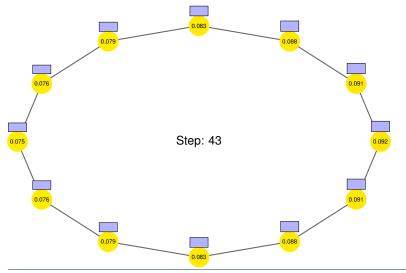
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



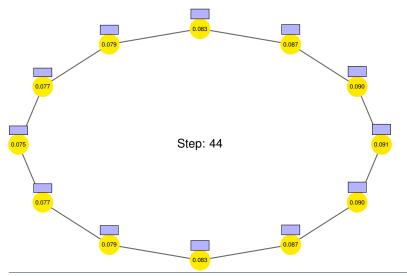
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



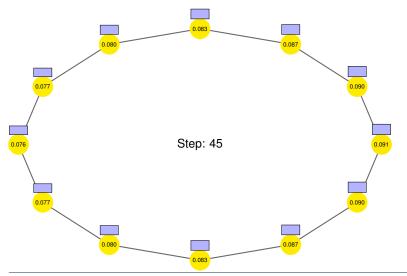
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



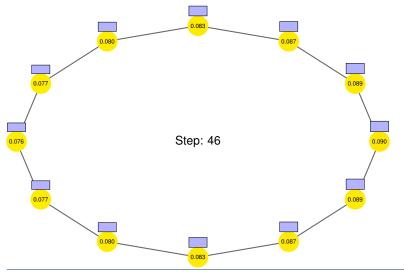
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



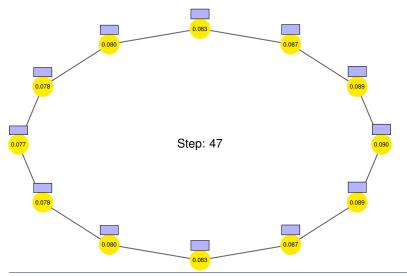
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



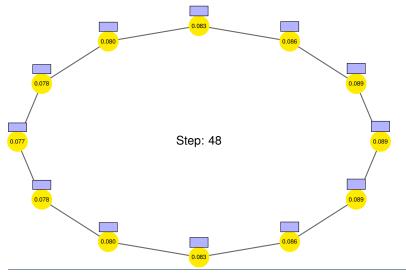
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



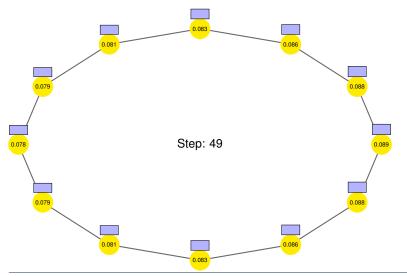
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



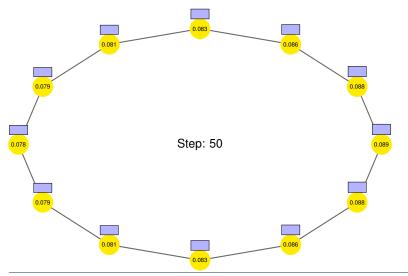
- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



- Markov Chain: stays put with 1/2 and moves left (or right) w.p. 1/4
- At step t the value at vertex $x \in \{1, 2, ..., 12\}$ is $P^t(1, x)$.



Outline

Recap of Markov Chain Basics

Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

Application 1: Card Shuffling

Application 2: Ehrenfest Chain and Hypercubes

Application 3: Markov Chain Monte Carlo

How Similar are Two Probability Measures?

Loaded Dice

• You are presented three loaded (unfair) dice A, B, C:

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24







How Similar are Two Probability Measures?

Loaded Dice -

You are presented three loaded (unfair) dice A, B, C:

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

• Question 1: Which dice is the least fair?







Loaded Dice -

X	1	2	3	4	5	6
$\mathbf{P}[A=x]$	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

- Question 1: Which dice is the least fair?
- Question 2: Which dice is the most fair?



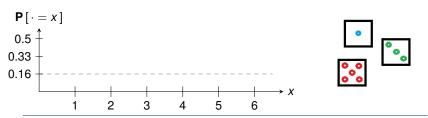




Loaded Dice -

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

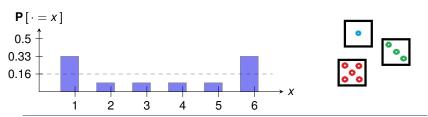
- Question 1: Which dice is the least fair?
- Question 2: Which dice is the most fair?



Loaded Dice -

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

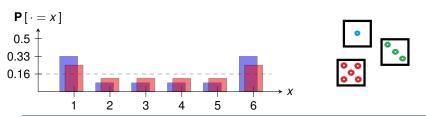
- Question 1: Which dice is the least fair?
- Question 2: Which dice is the most fair?



Loaded Dice -

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

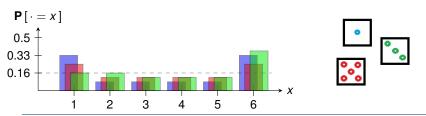
- Question 1: Which dice is the least fair?
- Question 2: Which dice is the most fair?



Loaded Dice -

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

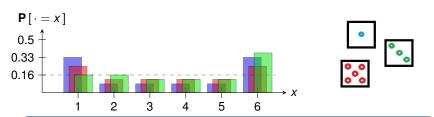
- Question 1: Which dice is the least fair?
- Question 2: Which dice is the most fair?



Loaded Dice -

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

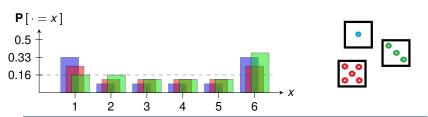
- Question 1: Which dice is the least fair? Most of you choose A.
- Question 2: Which dice is the most fair?



Loaded Dice

X	1	2	3	4	5	6
P[A=x]	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

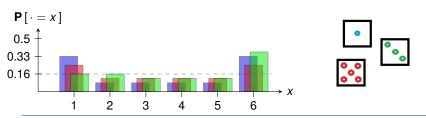
- Question 1: Which dice is the least fair? Most of you choose A. Why?
- Question 2: Which dice is the most fair?



Loaded Dice

X	1	2	3	4	5	6
$\mathbf{P}[A=x]$	1/3	1/12	1/12	1/12	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

- Question 1: Which dice is the least fair? Most of you choose A. Why?
- Question 2: Which dice is the most fair? Dice B and C seem "fairer" than A but which is fairest?



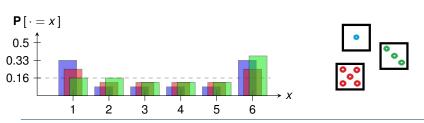
Loaded Dice

You are presented three loaded (unfair) dice A, B, C:

X	1	2	3	4	5	6
P[A=x]	1/3			l	1/12	1/3
P[B=x]	1/4	1/8	1/8	1/8	1/8	1/4
P[C=x]	1/6	1/6	1/8	1/8	1/8	9/24

- Question 1: Which dice is the least fair? Most of you choose A. Why?
- Question 2: Which dice is the most fair? Dice B and C seem "fairer" than A but which is fairest?

We need a formal "fairness measure" to compare probability distributions!



The Total Variation Distance between two probability distributions μ and η on a countable state space Ω is given by

$$\left\|\mu-\eta
ight\|_{tv}=rac{1}{2}\sum_{\omega\in\Omega}|\mu(\omega)-\eta(\omega)|.$$

The Total Variation Distance between two probability distributions μ and η on a countable state space Ω is given by

$$\|\mu - \eta\|_{tv} = \frac{1}{2} \sum_{\omega \in \Omega} |\mu(\omega) - \eta(\omega)|.$$

Loaded Dice: let $D = Unif\{1, 2, 3, 4, 5, 6\}$ be the law of a fair dice:

$$\|D - A\|_{tv} = \frac{1}{2} \left(2 \left| \frac{1}{6} - \frac{1}{3} \right| + 4 \left| \frac{1}{6} - \frac{1}{12} \right| \right) = \frac{1}{3}$$

The Total Variation Distance between two probability distributions μ and η on a countable state space Ω is given by

$$\|\mu - \eta\|_{tv} = \frac{1}{2} \sum_{\omega \in \Omega} |\mu(\omega) - \eta(\omega)|.$$

Loaded Dice: let $D = Unif\{1, 2, 3, 4, 5, 6\}$ be the law of a fair dice:

$$\begin{split} \|D - A\|_{tv} &= \frac{1}{2} \left(2 \left| \frac{1}{6} - \frac{1}{3} \right| + 4 \left| \frac{1}{6} - \frac{1}{12} \right| \right) = \frac{1}{3} \\ \|D - B\|_{tv} &= \frac{1}{2} \left(2 \left| \frac{1}{6} - \frac{1}{4} \right| + 4 \left| \frac{1}{6} - \frac{1}{8} \right| \right) = \frac{1}{6} \\ \|D - C\|_{tv} &= \frac{1}{2} \left(3 \left| \frac{1}{6} - \frac{1}{8} \right| + \left| \frac{1}{6} - \frac{9}{24} \right| \right) = \frac{1}{6}. \end{split}$$

The Total Variation Distance between two probability distributions μ and η on a countable state space Ω is given by

$$\|\mu - \eta\|_{tv} = \frac{1}{2} \sum_{\omega \in \Omega} |\mu(\omega) - \eta(\omega)|.$$

Loaded Dice: let $D = Unif\{1, 2, 3, 4, 5, 6\}$ be the law of a fair dice:

$$\begin{split} \|D - A\|_{tv} &= \frac{1}{2} \left(2 \left| \frac{1}{6} - \frac{1}{3} \right| + 4 \left| \frac{1}{6} - \frac{1}{12} \right| \right) = \frac{1}{3} \\ \|D - B\|_{tv} &= \frac{1}{2} \left(2 \left| \frac{1}{6} - \frac{1}{4} \right| + 4 \left| \frac{1}{6} - \frac{1}{8} \right| \right) = \frac{1}{6} \\ \|D - C\|_{tv} &= \frac{1}{2} \left(3 \left| \frac{1}{6} - \frac{1}{8} \right| + \left| \frac{1}{6} - \frac{9}{24} \right| \right) = \frac{1}{6}. \end{split}$$

Thus

$$\|D - B\|_{tv} = \|D - C\|_{tv}$$
 and $\|D - B\|_{tv}, \|D - C\|_{tv} < \|D - A\|_{tv}.$

So A is the least "fair" however B and C are equally "fair" (in TV distance).

Let P be a finite Markov Chain with stationary distribution π .

Let *P* be a finite Markov Chain with stationary distribution π .

• Let μ be a prob. vector on Ω (might be just one vertex) and $t \geq 0$. Then

$$P_{\mu}^{t} := \mathbf{P} \left[X_{t} = \cdot \mid X_{0} \sim \mu \right],$$

is a probability measure on Ω .

Let *P* be a finite Markov Chain with stationary distribution π .

• Let μ be a prob. vector on Ω (might be just one vertex) and $t \geq 0$. Then

$$P_{\mu}^{t} := \mathbf{P} \left[X_{t} = \cdot \mid X_{0} \sim \mu \right],$$

is a probability measure on Ω .

• For any μ ,

$$\left\| P_{\mu}^{t} - \pi \right\|_{tv} \leq \max_{x \in \Omega} \left\| P_{x}^{t} - \pi \right\|_{tv}.$$

Let *P* be a finite Markov Chain with stationary distribution π .

• Let μ be a prob. vector on Ω (might be just one vertex) and $t \geq 0$. Then

$$P_{\mu}^{t} := \mathbf{P} \left[X_{t} = \cdot \mid X_{0} \sim \mu \right],$$

is a probability measure on Ω .

• For any μ ,

$$\left\| P_{\mu}^{t} - \pi \right\|_{tv} \leq \max_{x \in \Omega} \left\| P_{x}^{t} - \pi \right\|_{tv}.$$

Convergence Theorem (Implication for TV Distance)

For any finite, irreducible, aperiodic Markov Chain

$$\lim_{t\to\infty}\max_{x\in\Omega}\left\|P_x^t-\pi\right\|_{t\nu}=0.$$

Let *P* be a finite Markov Chain with stationary distribution π .

• Let μ be a prob. vector on Ω (might be just one vertex) and $t \geq 0$. Then

$$P_{\mu}^{t} := \mathbf{P} \left[X_{t} = \cdot \mid X_{0} \sim \mu \right],$$

is a probability measure on Ω .

• For any μ ,

$$\left\| P_{\mu}^{t} - \pi \right\|_{tv} \leq \max_{x \in \Omega} \left\| P_{x}^{t} - \pi \right\|_{tv}.$$

Convergence Theorem (Implication for TV Distance)

For any finite, irreducible, aperiodic Markov Chain

$$\lim_{t\to\infty} \max_{x\in\Omega} \left\| P_x^t - \pi \right\|_{tv} = 0.$$

We will prove a similar result later after introducing spectral techniques!

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Question: How fast do they converge?

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Question: How fast do they converge?

Mixing Time -

The Mixing time $\tau_{\rm X}(\epsilon)$ of a finite Markov Chain P with stationary distribution π is defined as

$$\tau_{\mathsf{X}}(\epsilon) = \min \left\{ t \colon \left\| P_{\mathsf{X}}^t - \pi \right\|_{t_{\mathsf{Y}}} \le \epsilon \right\},$$

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Question: How fast do they converge?

- Mixing Time -

The Mixing time $\tau_{\rm x}(\epsilon)$ of a finite Markov Chain P with stationary distribution π is defined as

$$\tau_{\mathsf{X}}(\epsilon) = \min \left\{ t \colon \left\| P_{\mathsf{X}}^t - \pi \right\|_{t\mathsf{Y}} \le \epsilon \right\},$$

and,

$$\tau(\epsilon) = \max_{\mathsf{x}} \tau_{\mathsf{x}}(\epsilon).$$

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Question: How fast do they converge?

- Mixing Time -

The Mixing time $\tau_{\rm X}(\epsilon)$ of a finite Markov Chain P with stationary distribution π is defined as

$$\tau_{\mathsf{X}}(\epsilon) = \min \left\{ t \colon \left\| P_{\mathsf{X}}^t - \pi \right\|_{t\mathsf{Y}} \le \epsilon \right\},$$

and,

$$\tau(\epsilon) = \max_{\mathsf{x}} \tau_{\mathsf{x}}(\epsilon).$$

• This is how long we need to wait until we are " ε -close" to stationarity

Convergence Theorem: "Nice" Markov Chains converge to stationarity.

Question: How fast do they converge?

Mixing Time -

The Mixing time $\tau_{\rm x}(\epsilon)$ of a finite Markov Chain P with stationary distribution π is defined as

$$\tau_{\mathsf{X}}(\epsilon) = \min \left\{ t \colon \left\| P_{\mathsf{X}}^t - \pi \right\|_{t_{\mathsf{Y}}} \le \epsilon \right\},$$

and.

$$\tau(\epsilon) = \max_{\mathbf{x}} \tau_{\mathbf{X}}(\epsilon).$$

- This is how long we need to wait until we are "ε-close" to stationarity
- We often take $\varepsilon = 1/4$, indeed let $t_{mix} := \tau(1/4)$

Outline

Recap of Markov Chain Basics

Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

Application 1: Card Shuffling

Application 2: Ehrenfest Chain and Hypercubes

Application 3: Markov Chain Monte Carlo



Source: wikipedia

How long does it take to shuffle a deck of 52 cards?



Source: wikipedia

How long does it take to shuffle a deck of 52 cards?



Persi Diaconis (Professor of Statistics and former Magician)



Source: wikipedia

How long does it take to shuffle a deck of 52 cards?



His research revealed a lot of beautiful connections between Markov Chains and Algebra.

Persi Diaconis (Professor of Statistics and former Magician)



Source: wikipedia

Here we will focus on one shuffling scheme which is easy to analyse.

How long does it take to shuffle a deck of 52 cards?



His research revealed a lot of beautiful connections between Markov Chains and Algebra.

Persi Diaconis (Professor of Statistics and former Magician)



Source: wikipedia

Here we will focus on one shuffling scheme which is easy to analyse.

How long does it take to shuffle a deck of 52 cards?

How quickly do we converge to the uniform distribution over all n! permutations?



His research revealed a lot of beautiful connections between Markov Chains and Algebra.

Persi Diaconis (Professor of Statistics and former Magician)

The Card Shuffling Markov Chain

TOPTORANDOMSHUFFLE (Input: A pile of *n* cards)

- 1: **For** t = 1, 2, ...
- 2: Pick $i \in \{1, 2, ..., n\}$ uniformly at random
- 3: Take the top card and insert it behind the *i*-th card

The Card Shuffling Markov Chain

TOPTORANDOMSHUFFLE (Input: A pile of *n* cards)

- 1: **For** t = 1, 2, ...
- 2: Pick $i \in \{1, 2, ..., n\}$ uniformly at random
- 3: Take the top card and insert it behind the *i*-th card

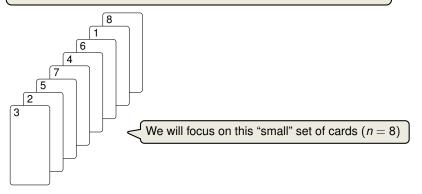
This is a slightly informal definition, so let us look at a small example...

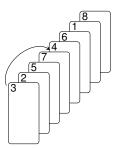
The Card Shuffling Markov Chain

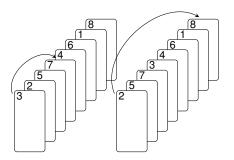
TOPTORANDOMSHUFFLE (Input: A pile of *n* cards)

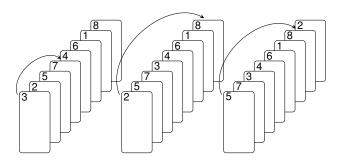
- 1: **For** t = 1, 2, ...
- 2: Pick $i \in \{1, 2, ..., n\}$ uniformly at random
- 3: Take the top card and insert it behind the *i*-th card

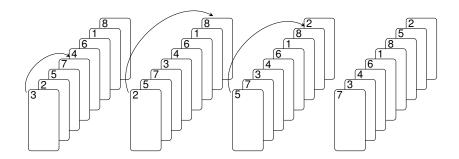
This is a slightly informal definition, so let us look at a small example...

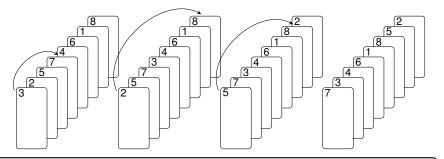


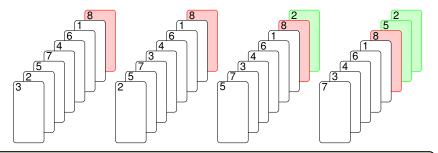


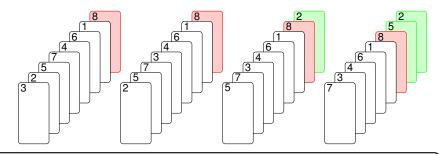


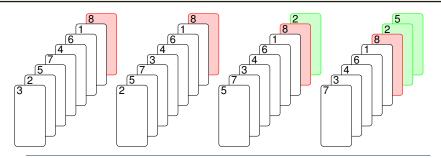


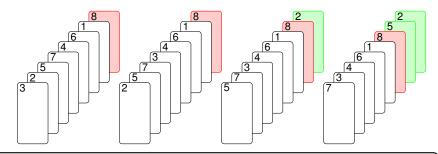


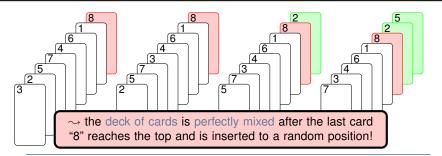


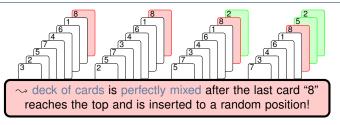


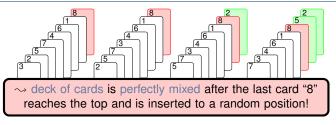




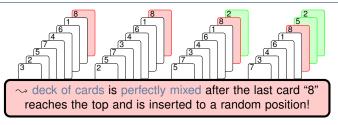




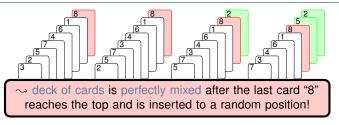




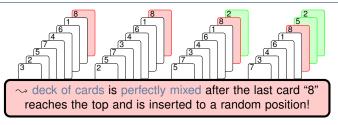
How long does it take for the last card "n" to become top card?



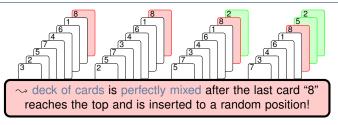
- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step



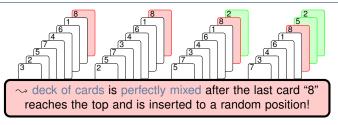
- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$



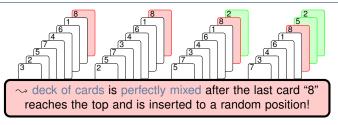
- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$.



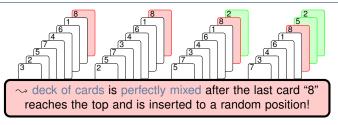
- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$:
- At the second position, card "n" moves up with probability $\frac{n-1}{n}$



- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$
- At the second position, card "n" moves up with probability $\frac{n-1}{n}$
- One final step to randomise card "n"

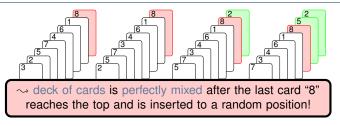


- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$:
- At the second position, card "n" moves up with probability $\frac{n-1}{n}$
- One final step to randomise card "n" (with probability 1)



- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$:
- At the second position, card "n" moves up with probability $\frac{n-1}{n}$
- One final step to randomise card "n" (with probability 1)

This is a "reversed" coupon collector process with n cards, which takes $n \log n$ in expectation.



- How long does it take for the last card "n" to become top card?
- At the last position, card "n" moves up with probability $\frac{1}{n}$ at each step
- At the second last position, card "n" moves up with probability $\frac{2}{n}$:
- At the second position, card "n" moves up with probability $\frac{n-1}{n}$
- One final step to randomise card "n" (with probability 1)

This is a "reversed" coupon collector process with n cards, which takes $n \log n$ in expectation.

Using the so-called coupling method, one could prove $t_{mix} \leq n \log n$.

Riffle Shuffle -

1. Split a deck of *n* cards into two piles (thus the size of each portion will be Binomial)

Riffle Shuffle

- Split a deck of n cards into two piles (thus the size of each portion will be Binomial)
- 2. Riffle the cards together so that the card drops from the left (or right) pile with probability proportional to the number of remaining cards

Riffle Shuffle

- Split a deck of n cards into two piles (thus the size of each portion will be Binomial)
- 2. Riffle the cards together so that the card drops from the left (or right) pile with probability proportional to the number of remaining cards

Riffle Shuffle

- Split a deck of n cards into two piles (thus the size of each portion will be Binomial)
- 2. Riffle the cards together so that the card drops from the left (or right) pile with probability proportional to the number of remaining cards

t | 1 2 3 4 5 6 7 8 9 10
$$||P^t - \pi||_{tv}$$
 | 1.000 1.000 1.000 1.000 0.924 0.614 0.334 0.167 0.085 0.043

Figure: Total Variation Distance for *t* riffle shuffles of 52 cards.

Riffle Shuffle

- Split a deck of n cards into two piles (thus the size of each portion will be Binomial)
- Riffle the cards together so that the card drops from the left (or right) pile with probability proportional to the number of remaining cards





The Annals of Applied Probability 1992, Vol. 2, No. 2, 294-313

TRAILING THE DOVETAIL SHUFFLE TO ITS LAIR

By Dave Bayer1 and Persi Diaconis2

Columbia University and Harvard University

We analyze the most commonly used method for shuffling cards. The main result is a simple expression for the chance of any arrangement after any number of shuffles. This is used to give sharp bounds on the approach to randomness: $\frac{2}{3} \log_2 n + \theta$ shuffles are necessary and sufficient to mix up n cards.

Key ingredients are the analysis of a card trick and the determination of the idempotents of a natural commutative subalgebra in the symmetric group algebra.

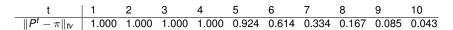


Figure: Total Variation Distance for *t* riffle shuffles of 52 cards.

Outline

Recap of Markov Chain Basics

Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

Application 1: Card Shuffling

Application 2: Ehrenfest Chain and Hypercubes

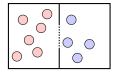
Application 3: Markov Chain Monte Carlo

Ehrenfest Model ——

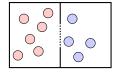
 A simple model for the exchange of molecules between two boxes

Ehrenfest Model ---

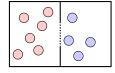
 A simple model for the exchange of molecules between two boxes



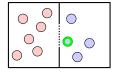
- A simple model for the exchange of molecules between two boxes
- We have d particles



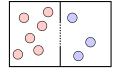
- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box



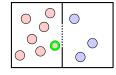
- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box



- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box

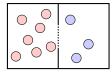


- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box



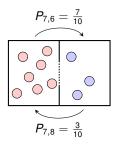
- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.



- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

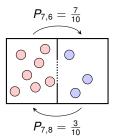
$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.



Ehrenfest Model -

- A simple model for the exchange of molecules between two boxes
- We have d particles
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.

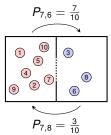


Let us now enlarge the state space by looking at each particle individually!

Ehrenfest Model -

- A simple model for the exchange of molecules between two boxes
- We have d particles labelled 1, 2, . . . , d
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.

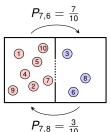


Let us now enlarge the state space by looking at each particle individually!

Ehrenfest Model -

- A simple model for the exchange of molecules between two boxes
- We have d particles labelled 1, 2, ..., d
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.



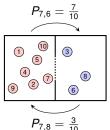
Let us now enlarge the state space by looking at each particle individually!

Random Walk on the Hypercube

Ehrenfest Model -

- A simple model for the exchange of molecules between two boxes
- We have d particles labelled 1, 2, ..., d
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.



Let us now enlarge the state space by looking at each particle individually!

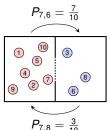
- Random Walk on the Hypercube ————
- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$



Ehrenfest Model -

- A simple model for the exchange of molecules between two boxes
- We have d particles labelled 1, 2, . . . , d
- At each step a particle is selected uniformly at random and switches to the other box
- If $\Omega = \{0, 1, ..., d\}$ denotes the number of particles in the red box, then:

$$P_{x,x-1} = \frac{x}{d}$$
 and $P_{x,x+1} = \frac{d-x}{d}$.



Let us now enlarge the state space by looking at each particle individually!

Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it

Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version) -

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version) -

• At each step t = 0, 1, 2...



(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version) -

- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version) -

- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]
 - With prob. 1/2 flip coordinate.

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version)

- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]
 - With prob. 1/2 flip coordinate.

Lazy Random Walk (2nd Version)

- At each step t = 0, 1, 2...
 - Pick a random coordinate in [d]

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version)

- At each step t = 0, 1, 2...
 - Pick a random coordinate in [d]
 - With prob. 1/2 flip coordinate.

Lazy Random Walk (2nd Version)

- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]
 - Set coordinate to {0,1} uniformly.

(Non-Lazy) Random Walk on the Hypercube

- For each particle an indicator variable $\Rightarrow \Omega = \{0, 1\}^d$
- At each step: pick a random coordinate in [d] and flip it



Problem: This Markov Chain is periodic, as the number of ones always switches from odd to even!

Solution: Add self-loops to break periodic behaviour!

Lazy Random Walk (1st Version) -

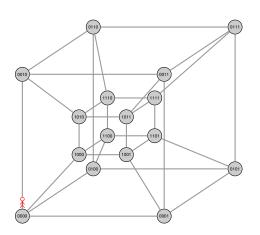
- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]
 - With prob. 1/2 flip coordinate.

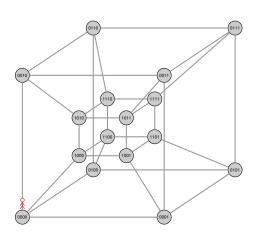
Lazy Random Walk (2nd Version)

- At each step t = 0, 1, 2 . . .
 - Pick a random coordinate in [d]
 - Set coordinate to {0,1} uniformly.

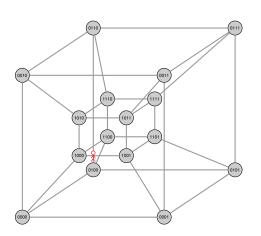


These two chains are equivalent!

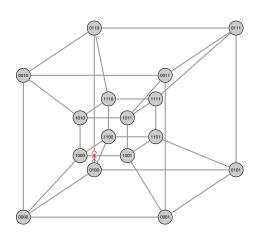




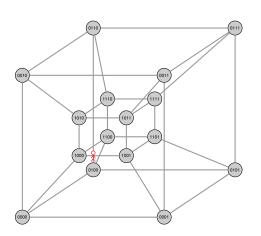
t	Coord.	X_t			
0	2	0	0	0	
1		0	?	0	



t	Coord.	X_t			
0	2	0	0	0	
1		0	1	0	

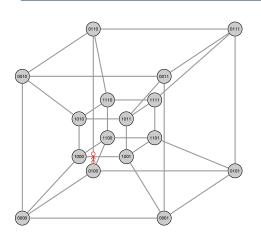


t	Coord.	X_t			
0	2	0	0	0	
1	3	0	1	0	
2		0	1	?	

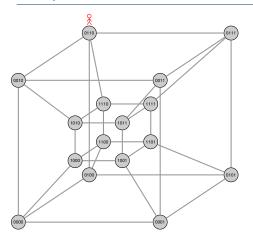


t	Coord.	
)	2	0
I	3	0
2		0

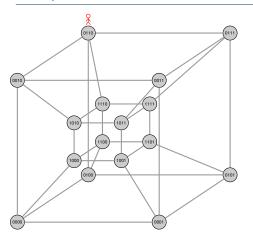
 X_t



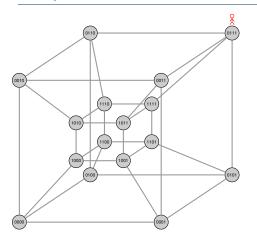
	Coord.	X_t				
)	2	0	0	0	0	
	3	0	1	0	0	
2	3	0	1	0	0	
3		0	1	?	0	



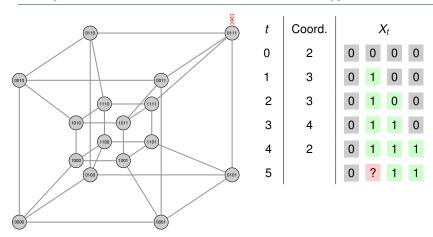
t	Coord.	X_t			
0	2	0	0	0	C
1	3	0	1	0	C
2	3	0	1	0	C
3		0	1	1	C

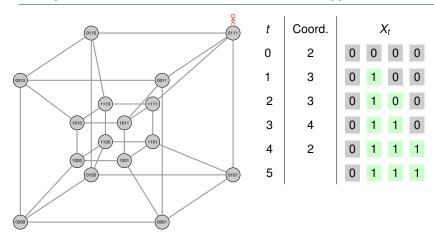


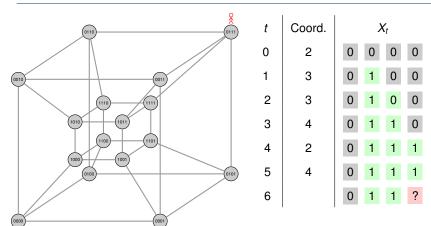
t	Coord.	X_t			
0	2	0	0	0	(
1	3	0	1	0	(
2	3	0	1	0	(
3	4	0	1	1	(
4		0	1	1	•

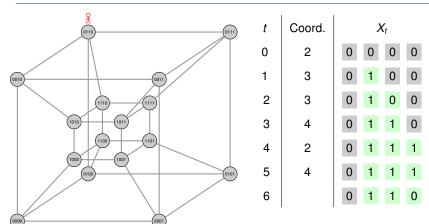


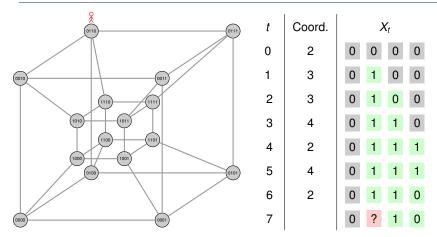
t	Coord.	X_t			
0	2	0	0	0	C
1	3	0	1	0	C
2	3	0	1	0	C
3	4	0	1	1	C
4		0	1	1	1

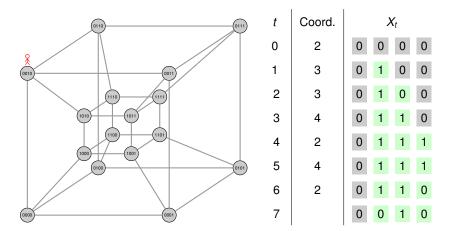


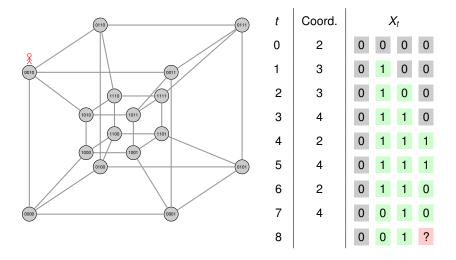


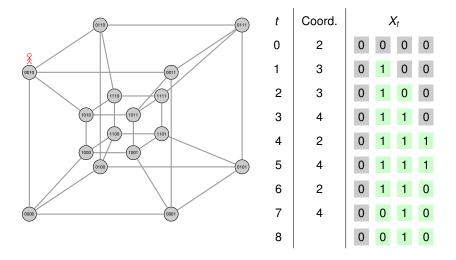


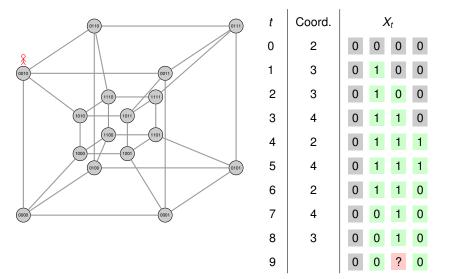


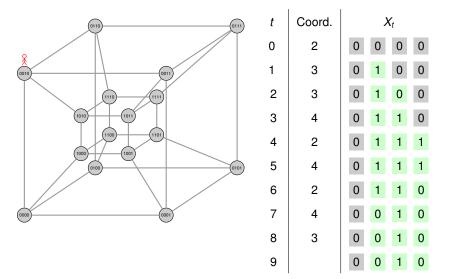


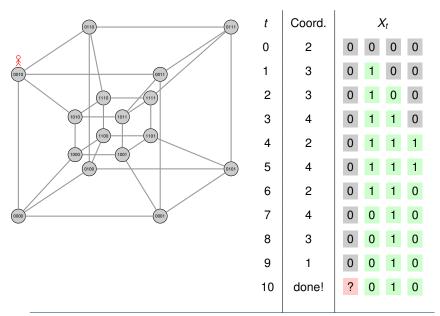


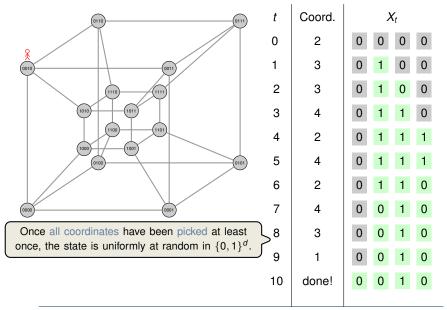


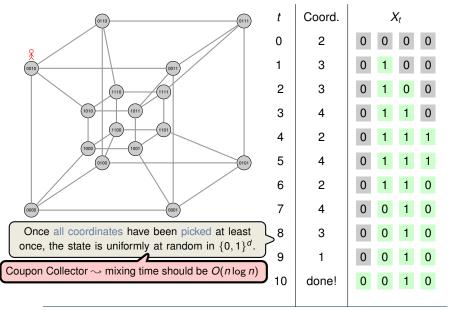


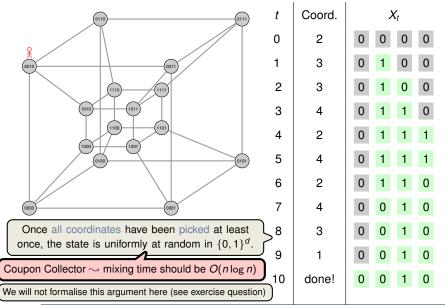




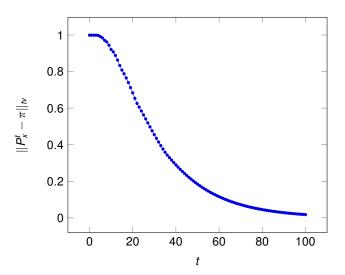




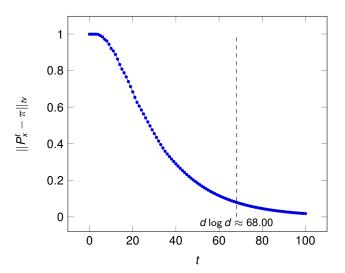




Total Variation Distance of Random Walk on Hypercube (d = 22)



Total Variation Distance of Random Walk on Hypercube (d = 22)



RANDOM WALK ON A HYPERCUBE

53

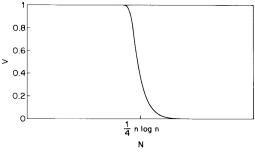


Fig. 1. The variation distance V as a function of N, for $n = 10^{12}$.

Source: "Asymptotic analysis of a random walk on a hypercube with many dimensions", P. Diaconis, R.L. Graham, J.A. Morrison; Random Structures & Algorithms, 1990.





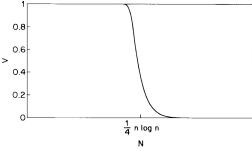


Fig. 1. The variation distance V as a function of N, for $n = 10^{12}$.

Source: "Asymptotic analysis of a random walk on a hypercube with many dimensions", P. Diaconis, R.L. Graham, J.A. Morrison; Random Structures & Algorithms, 1990.

- This is a numerical plot of a theoretical bound, where $d = 10^{12}$ (Minor Remark: This random walk is with a loop probability of 1/(d+1))
- The variation distance exhibits a so-called cut-off phenomena:





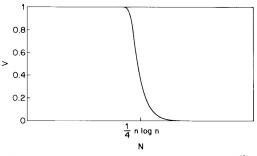


Fig. 1. The variation distance V as a function of N, for $n = 10^{12}$.

Source: "Asymptotic analysis of a random walk on a hypercube with many dimensions", P. Diaconis, R.L. Graham, J.A. Morrison; Random Structures & Algorithms, 1990.

- This is a numerical plot of a theoretical bound, where $d = 10^{12}$ (Minor Remark: This random walk is with a loop probability of 1/(d+1))
- The variation distance exhibits a so-called cut-off phenomena:
 - Distance remains close to its maximum value 1 until step $\frac{1}{4}n \log n \Theta(n)$
 - Then distance moves close to 0 before step $\frac{1}{4}n \log n + \Theta(n)$

Outline

Recap of Markov Chain Basics

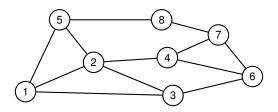
Irreducibility, Periodicity and Convergence

Total Variation Distance and Mixing Times

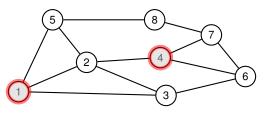
Application 1: Card Shuffling

Application 2: Ehrenfest Chain and Hypercubes

Application 3: Markov Chain Monte Carlo

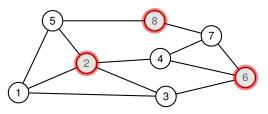


Independent Set



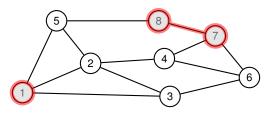
 $S = \{1, 4\}$ is an independent set \checkmark

Independent Set



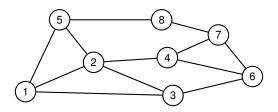
 $S = \{2, 6, 8\}$ is an independent set \checkmark

Independent Set

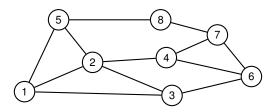


 $S = \{1, 7, 8\}$ is **not** an independent set \times

Independent Set



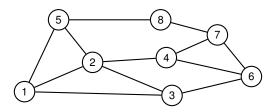
Independent Set



Independent Set

Given an undirected graph G = (V, E), an independent set is a subset $S \subseteq V$ such that there are no two vertices $u, v \in S$ with $\{u, v\} \in E(G)$.

How can we take a sample from the space of all independent sets?

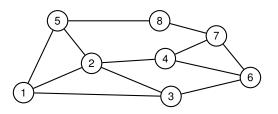


Independent Set

Given an undirected graph G = (V, E), an independent set is a subset $S \subseteq V$ such that there are no two vertices $u, v \in S$ with $\{u, v\} \in E(G)$.

How can we take a sample from the space of all independent sets?

Naive brute-force would take an insane amount of time (and space)!



Independent Set

Given an undirected graph G = (V, E), an independent set is a subset $S \subseteq V$ such that there are no two vertices $u, v \in S$ with $\{u, v\} \in E(G)$.

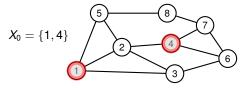
How can we take a sample from the space of all independent sets?

Naive brute-force would take an insane amount of time (and space)!

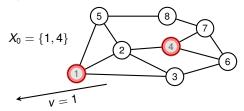
We can use a generic Markov Chain Monte Carlo approach to tackle this problem!

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: **else** $X_{t+1} \leftarrow X_t$

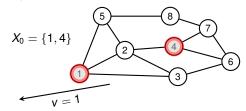
- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

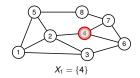


- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

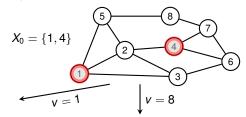


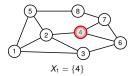
- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$



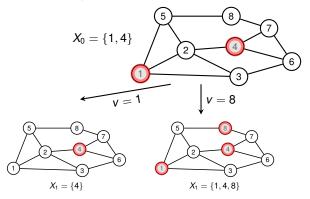


- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

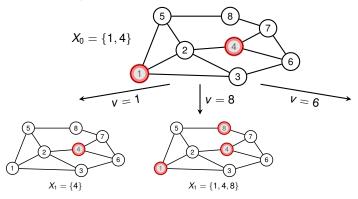




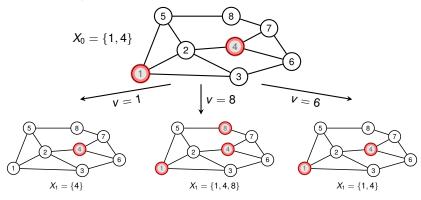
- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$



- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$



- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$



INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: **else** $X_{t+1} \leftarrow X_t$

Remark —

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

Remark -

This is a local definition (no explicit definition of P!)

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

Remark -

- This is a local definition (no explicit definition of P!)
- This chain is irreducible (every independent set is reachable)

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

Remark -

- This is a local definition (no explicit definition of P!)
- This chain is irreducible (every independent set is reachable)
- This chain is aperiodic (Check!)

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: **else** $X_{t+1} \leftarrow X_t$

Remark

- This is a local definition (no explicit definition of P!)
- This chain is irreducible (every independent set is reachable)
- This chain is aperiodic (Check!)
- The stationary distribution is uniform, since $P_{u,v} = P_{v,u}$ (Check!)

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...:
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: else $X_{t+1} \leftarrow X_t$

Remark

- This is a local definition (no explicit definition of P!)
- This chain is irreducible (every independent set is reachable)
- This chain is aperiodic (Check!)
- The stationary distribution is uniform, since $P_{u,v} = P_{v,u}$ (Check!)

Key Question: What is the mixing time of this Markov Chain?

INDEPENDENTSETSAMPLER

- 1: Let X_0 be an arbitrary independent set in G
- 2: **For** t = 1, 2, ...
- 3: Pick a vertex $v \in V(G)$ uniformly at random
- 4: If $v \in X_t$ then $X_{t+1} \leftarrow X_t \setminus \{v\}$
- 5: **elif** $v \notin X_t$ **and** $X_t \cup \{v\}$ is an independent set **then** $X_{t+1} \leftarrow X_t \cup \{v\}$
- 6: **else** $X_{t+1} \leftarrow X_t$

Remark

- This is a local definition (no explicit definition of P!)
- This chain is irreducible (every independent set is reachable)
- This chain is aperiodic (Check!)
- The stationary distribution is uniform, since $P_{u,v} = P_{v,u}$ (Check!)

Key Question: What is the mixing time of this Markov Chain?

not covered here, see the textbook of Mitzenmacher & Upfal