

# Randomised Algorithms

## Lecture 2-3: Concentration Inequalities

Thomas Sauerwald (tms41@cam.ac.uk)

Lent 2022



UNIVERSITY OF  
CAMBRIDGE

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

Application 1: Balls into Bins

Application 2: Randomised QuickSort

Extensions of Chernoff Bounds

Applications of Method of Bounded Differences

Appendix

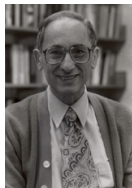
- **Concentration** refers to the phenomena where random variables are very close to their mean

- **Concentration** refers to the phenomena where random variables are very close to their mean
- This is very useful in randomised algorithms as it ensures an **almost** deterministic behaviour

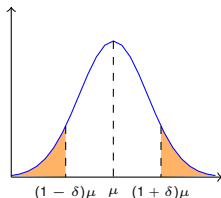
- **Concentration** refers to the phenomena where random variables are very close to their mean
- This is very useful in randomised algorithms as it ensures an **almost** deterministic behaviour
- It gives us the best of two worlds:
  1. **Randomised Algorithms:** Easy to Design and Implement
  2. **Deterministic Algorithms:** They do what they claim

## Chernoff Bounds: A Tool for Concentration

- Chernoff's bounds are “strong” bounds on the tail probabilities of **sums of independent random variables**
- random variables can be **discrete** (or continuous)
- usually these bounds decrease **exponentially** as opposed to a polynomial decrease in Markov's or Chebyshev's inequality (see example)

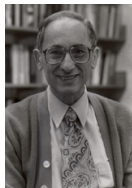


Hermann Chernoff (1923-)

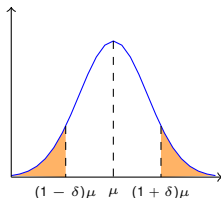


## Chernoff Bounds: A Tool for Concentration

- Chernoff's bounds are “strong” bounds on the tail probabilities of **sums of independent random variables**
- random variables can be **discrete** (or continuous)
- usually these bounds decrease **exponentially** as opposed to a polynomial decrease in Markov's or Chebyshev's inequality (see example)
- easy to apply, but **requires independence**

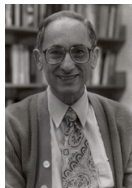


Hermann Chernoff (1923-)

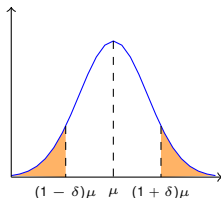


## Chernoff Bounds: A Tool for Concentration

- Chernoff's bounds are “strong” bounds on the tail probabilities of **sums of independent random variables**
- random variables can be **discrete** (or continuous)
- usually these bounds decrease **exponentially** as opposed to a polynomial decrease in Markov's or Chebyshev's inequality (see example)
- easy to apply, but **requires independence**
- have found various applications in:
  - Randomised Algorithms
  - Statistics
  - Random Projections and Dimensionality Reduction
  - Learning Theory (e.g., PAC-learning)
- $\vdots$



Hermann Chernoff (1923-)





## Recap: Markov and Chebyshev

---

Markov's Inequality

If  $X$  is a non-negative random variable, then for any  $a > 0$ ,

$$\mathbf{P}[X \geq a] \leq \mathbf{E}[X]/a.$$

Chebyshev's Inequality

If  $X$  is a random variable, then for any  $a > 0$ ,

$$\mathbf{P}[|X - \mathbf{E}[X]| \geq a] \leq \mathbf{V}[X]/a^2.$$

## Recap: Markov and Chebyshev

---

Markov's Inequality

If  $X$  is a non-negative random variable, then for any  $a > 0$ ,

$$\mathbf{P}[X \geq a] \leq \mathbf{E}[X]/a.$$

Chebyshev's Inequality

If  $X$  is a random variable, then for any  $a > 0$ ,

$$\mathbf{P}[|X - \mathbf{E}[X]| \geq a] \leq \mathbf{V}[X]/a^2.$$

- Let  $f : \mathbb{R} \rightarrow [0, \infty)$  and **increasing**, then  $f(X) \geq 0$ , and thus

$$\mathbf{P}[X \geq a] \leq \mathbf{P}[f(X) \geq f(a)] \leq \mathbf{E}[f(X)]/f(a).$$

## Recap: Markov and Chebyshev

Markov's Inequality

If  $X$  is a non-negative random variable, then for any  $a > 0$ ,

$$\mathbf{P}[X \geq a] \leq \mathbf{E}[X]/a.$$

Chebyshev's Inequality

If  $X$  is a random variable, then for any  $a > 0$ ,

$$\mathbf{P}[|X - \mathbf{E}[X]| \geq a] \leq \mathbf{V}[X]/a^2.$$

- Let  $f : \mathbb{R} \rightarrow [0, \infty)$  and **increasing**, then  $f(X) \geq 0$ , and thus

$$\mathbf{P}[X \geq a] \leq \mathbf{P}[f(X) \geq f(a)] \leq \mathbf{E}[f(X)]/f(a).$$

- Similarly, if  $g : \mathbb{R} \rightarrow [0, \infty)$  and **decreasing**, then  $g(X) \geq 0$ , and thus

$$\mathbf{P}[X \leq a] \leq \mathbf{P}[g(X) \geq g(a)] \leq \mathbf{E}[g(X)]/g(a).$$

## Recap: Markov and Chebyshev

### Markov's Inequality

If  $X$  is a non-negative random variable, then for any  $a > 0$ ,

$$\mathbf{P}[X \geq a] \leq \mathbf{E}[X]/a.$$

### Chebyshev's Inequality

If  $X$  is a random variable, then for any  $a > 0$ ,

$$\mathbf{P}[|X - \mathbf{E}[X]| \geq a] \leq \mathbf{V}[X]/a^2.$$

- Let  $f : \mathbb{R} \rightarrow [0, \infty)$  and **increasing**, then  $f(X) \geq 0$ , and thus

$$\mathbf{P}[X \geq a] \leq \mathbf{P}[f(X) \geq f(a)] \leq \mathbf{E}[f(X)]/f(a).$$

- Similarly, if  $g : \mathbb{R} \rightarrow [0, \infty)$  and **decreasing**, then  $g(X) \geq 0$ , and thus

$$\mathbf{P}[X \leq a] \leq \mathbf{P}[g(X) \geq g(a)] \leq \mathbf{E}[g(X)]/g(a).$$

Chebyshev's inequality (or Markov) can be obtained by choosing  $f(X) := (X - \mu)^2$  (or  $f(X) := X$ , respectively).

## From Markov and Chebyshev to Chernoff

---

Markov and Chebyshev use the **first and second moment** of the random variable. Can we keep going?

Markov and Chebyshev use the **first and second moment** of the random variable. Can we keep going?

- **Yes!**

Markov and Chebyshev use the **first and second moment** of the random variable. Can we keep going?

- **Yes!**

We can consider the first, second, **third and more** moments! That is the basic idea behind the **Chernoff Bounds**

## Our First Chernoff Bound

---

Chernoff Bounds (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1 + \delta)}} \right]^\mu. \quad (\star)$$



## Our First Chernoff Bound

### Chernoff Bounds (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are **independent Bernoulli** random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu. \quad (\star)$$

While  $(\star)$  is one of the easiest (and most generic) Chernoff bounds to derive, the bound is complicated and hard to apply...

## Our First Chernoff Bound

### Chernoff Bounds (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are **independent Bernoulli** random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu. \quad (\star)$$

This implies that for any  $t > \mu$ ,

$$\mathbf{P}[X \geq t] \leq e^{-\mu} \left( \frac{e\mu}{t} \right)^t.$$

While  $(\star)$  is one of the easiest (and most generic) Chernoff bounds to derive, the bound is complicated and hard to apply...

## Example: Coin Flips (1/3)

---

- Consider throwing a fair coin  $n$  times and count the total number of heads

## Example: Coin Flips (1/3)

---

- Consider throwing a fair coin  $n$  times and count the total number of heads
- $X_i \in \{0, 1\}$ ,  $X = \sum_{i=1}^n X_i$  and  $\mathbf{E}[X] = n \cdot 1/2 = n/2$

## Example: Coin Flips (1/3)

---

- Consider throwing a **fair coin**  $n$  times and count the **total number of heads**
- $X_i \in \{0, 1\}$ ,  $X = \sum_{i=1}^n X_i$  and  $\mathbf{E}[X] = n \cdot 1/2 = n/2$
- The **Chernoff Bound** gives for any  $\delta > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)(n/2)] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^{n/2}.$$

## Example: Coin Flips (1/3)

---

- Consider throwing a **fair coin**  $n$  times and count the **total number of heads**
- $X_i \in \{0, 1\}$ ,  $X = \sum_{i=1}^n X_i$  and  $\mathbf{E}[X] = n \cdot 1/2 = n/2$
- The **Chernoff Bound** gives for any  $\delta > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)(n/2)] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^{n/2}.$$

- The above expression equals 1 only for  $\delta = 0$ , and then it gives a value strictly less than 1 (**check this!**)

## Example: Coin Flips (1/3)

---

- Consider throwing a **fair coin**  $n$  times and count the **total number of heads**
- $X_i \in \{0, 1\}$ ,  $X = \sum_{i=1}^n X_i$  and  $\mathbf{E}[X] = n \cdot 1/2 = n/2$
- The **Chernoff Bound** gives for any  $\delta > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)(n/2)] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^{n/2}.$$

- The above expression equals 1 only for  $\delta = 0$ , and then it gives a value strictly less than 1 (**check this!**)
- The inequality is **exponential in  $n$** , (for fixed  $\delta$ ) which is much better than Chebyshev's inequality.

## Example: Coin Flips (1/3)

---

- Consider throwing a **fair coin**  $n$  times and count the **total number of heads**
- $X_i \in \{0, 1\}$ ,  $X = \sum_{i=1}^n X_i$  and  $\mathbf{E}[X] = n \cdot 1/2 = n/2$
- The **Chernoff Bound** gives for any  $\delta > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)(n/2)] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^{n/2}.$$

- The above expression equals 1 only for  $\delta = 0$ , and then it gives a value strictly less than 1 (**check this!**)
- The inequality is **exponential in  $n$** , (for fixed  $\delta$ ) which is much better than Chebyshev's inequality.

What about a **concrete value** of  $n$ , say  $n = 100$ ?



## Example: Coin Flips (2/3)

---

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

## Example: Coin Flips (2/3)

---

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

- Markov's inequality:  $\mathbf{E}[X] = 100/2 = 50$ .

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq 2/3 = \mathbf{0.666}.$$

## Example: Coin Flips (2/3)

---

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

- Markov's inequality:  $\mathbf{E}[X] = 100/2 = 50$ .

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq 2/3 = \mathbf{0.666}.$$

- Chebyshev's inequality:  $\mathbf{V}[X] = \sum_{i=1}^{100} \mathbf{V}[X_i] = 100 \cdot (1/2)^2 = 25$ .

$$\mathbf{P}[|X - \mu| \geq t] \leq \frac{\mathbf{V}[X]}{t^2},$$

and plugging in  $t = 25$  gives an upper bound of  $25/25^2 = 1/25 = \mathbf{0.04}$ , much better than what we obtained by Markov's inequality.

## Example: Coin Flips (2/3)

---

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

- Markov's inequality:  $\mathbf{E}[X] = 100/2 = 50$ .

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq 2/3 = \mathbf{0.666}.$$

- Chebyshev's inequality:  $\mathbf{V}[X] = \sum_{i=1}^{100} \mathbf{V}[X_i] = 100 \cdot (1/2)^2 = 25$ .

$$\mathbf{P}[|X - \mu| \geq t] \leq \frac{\mathbf{V}[X]}{t^2},$$

and plugging in  $t = 25$  gives an upper bound of  $25/25^2 = 1/25 = \mathbf{0.04}$ , much better than what we obtained by Markov's inequality.

- The Chernoff bound: with  $\delta = 1/2$  gives:

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq \left( \frac{e^{1/2}}{(3/2)^{3/2}} \right)^{50} = \mathbf{0.004472}.$$

## Example: Coin Flips (2/3)

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

- Markov's inequality:  $\mathbf{E}[X] = 100/2 = 50$ .

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq 2/3 = \mathbf{0.666}.$$

- Chebyshev's inequality:  $\mathbf{V}[X] = \sum_{i=1}^{100} \mathbf{V}[X_i] = 100 \cdot (1/2)^2 = 25$ .

$$\mathbf{P}[|X - \mu| \geq t] \leq \frac{\mathbf{V}[X]}{t^2},$$

and plugging in  $t = 25$  gives an upper bound of  $25/25^2 = 1/25 = \mathbf{0.04}$ , much better than what we obtained by Markov's inequality.

- The Chernoff bound: with  $\delta = 1/2$  gives:

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq \left( \frac{e^{1/2}}{(3/2)^{3/2}} \right)^{50} = \mathbf{0.004472}.$$

- Remark: The exact probability is  $\mathbf{0.0000028 \dots}$

## Example: Coin Flips (2/3)

Consider 100 independent coin flips. We wish to find an upper bound on the probability that the number of heads is greater or equal than 75.

- Markov's inequality:  $\mathbf{E}[X] = 100/2 = 50$ .

$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq 2/3 = \mathbf{0.666}.$$

- Chebyshev's inequality:  $\mathbf{V}[X] = \sum_{i=1}^{100} \mathbf{V}[X_i] = 100 \cdot (1/2)^2 = 25$ .

$$\mathbf{P}[|X - \mu| \geq t] \leq \frac{\mathbf{V}[X]}{t^2},$$

and plugging in  $t = 25$  gives an upper bound of  $25/25^2 = 1/25 = \mathbf{0.04}$ , much better than what we obtained by Markov's inequality.

- The Chernoff bound: with  $\delta = 1/2$  gives:

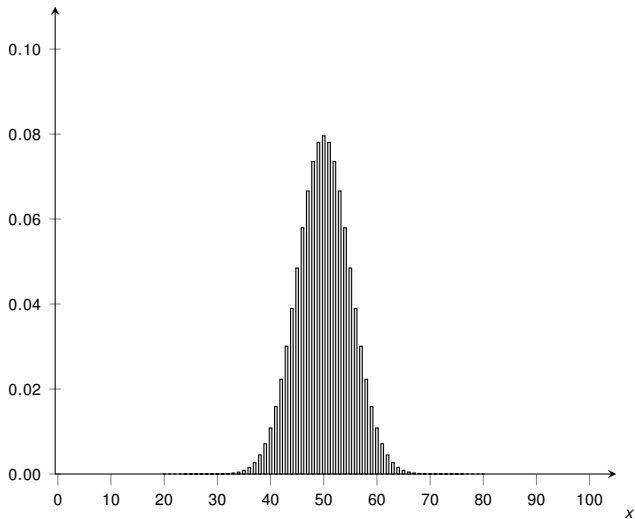
$$\mathbf{P}[X \geq 3/2 \cdot \mathbf{E}[X]] \leq \left( \frac{e^{1/2}}{(3/2)^{3/2}} \right)^{50} = \mathbf{0.004472}.$$

- Remark: The exact probability is  $\mathbf{0.00000028 \dots}$

Chernoff bound yields a much better result (but needs independence!)

## Example: Coin Flips (3/3)

$$P[\text{Bin}(100, 1/2) = x]$$



# Outline

---

Introduction to Chernoff Bounds

**How to Derive Chernoff Bounds**

Application 1: Balls into Bins

Application 2: Randomised QuickSort

Extensions of Chernoff Bounds

Applications of Method of Bounded Differences

Appendix



## General Recipe for Deriving Chernoff Bounds

---

Recipe

The **three main steps** in deriving Chernoff bounds for sums of **independent** random variables  $X = X_1 + \dots + X_n$  are:

## General Recipe for Deriving Chernoff Bounds

---

### Recipe

The **three main steps** in deriving Chernoff bounds for sums of **independent** random variables  $X = X_1 + \dots + X_n$  are:

1. Instead of working with  $X$ , we switch to the **moment generating function**  $e^{\lambda X}$ ,  $\lambda > 0$  and apply Markov's inequality  $\leadsto \mathbf{E} [ e^{\lambda X} ]$

## General Recipe for Deriving Chernoff Bounds

---

### Recipe

The **three main steps** in deriving Chernoff bounds for sums of **independent** random variables  $X = X_1 + \dots + X_n$  are:

1. Instead of working with  $X$ , we switch to the **moment generating function**  $e^{\lambda X}$ ,  $\lambda > 0$  and apply Markov's inequality  $\leadsto \mathbf{E} [ e^{\lambda X} ]$
2. Compute an upper bound for  $\mathbf{E} [ e^{\lambda X} ]$  (using independence)

## General Recipe for Deriving Chernoff Bounds

---

### Recipe

The **three main steps** in deriving Chernoff bounds for sums of **independent** random variables  $X = X_1 + \dots + X_n$  are:

1. Instead of working with  $X$ , we switch to the **moment generating function**  $e^{\lambda X}$ ,  $\lambda > 0$  and apply Markov's inequality  $\leadsto \mathbf{E} [ e^{\lambda X} ]$
2. Compute an upper bound for  $\mathbf{E} [ e^{\lambda X} ]$  (using independence)
3. Optimise value of  $\lambda$  to obtain best tail bound

## Chernoff Bound: Proof

Chernoff Bound (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu.$$

Proof:

## Chernoff Bound: Proof

Chernoff Bound (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu.$$

Proof:

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda X} \text{ is incr}}{\leq} \mathbf{P}\left[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}\right] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}\left[e^{\lambda X}\right]$$

## Chernoff Bound: Proof

Chernoff Bound (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu.$$

Proof:

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda X} \text{ is incr}}{\leq} \mathbf{P}\left[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}\right] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}\left[e^{\lambda X}\right]$$

$$2. \mathbf{E}\left[e^{\lambda X}\right] = \mathbf{E}\left[e^{\lambda \sum_{i=1}^n X_i}\right] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X_i}\right]$$

## Chernoff Bound: Proof

Chernoff Bound (General Form, Upper Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq \left[ \frac{e^\delta}{(1 + \delta)^{(1+\delta)}} \right]^\mu.$$

Proof:

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda x} \text{ is incr}}{\leq} \mathbf{P}\left[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}\right] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}\left[e^{\lambda X}\right]$$

$$2. \mathbf{E}\left[e^{\lambda X}\right] = \mathbf{E}\left[e^{\lambda \sum_{i=1}^n X_i}\right] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X_i}\right]$$

3.

$$\mathbf{E}\left[e^{\lambda X_i}\right] = e^\lambda p_i + (1 - p_i) = 1 + p_i(e^\lambda - 1) \stackrel{1+x \leq e^x}{\leq} e^{p_i(e^\lambda - 1)}$$



# Chernoff Bound: Proof

---

## Chernoff Bound: Proof

---

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda x} \text{ is incr}}{=} \mathbf{P}\left[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}\right] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}\left[e^{\lambda X}\right]$$

$$2. \mathbf{E}\left[e^{\lambda X}\right] = \mathbf{E}\left[e^{\lambda \sum_{i=1}^n X_i}\right] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X_i}\right]$$

3.

$$\mathbf{E}\left[e^{\lambda X_i}\right] = e^{\lambda} p_i + (1 - p_i) = 1 + p_i(e^{\lambda} - 1) \stackrel{1+x \leq e^x}{\leq} e^{p_i(e^{\lambda} - 1)}$$

## Chernoff Bound: Proof

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda x} \text{ is incr}}{=} \mathbf{P}\left[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}\right] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}\left[e^{\lambda X}\right]$$

$$2. \mathbf{E}\left[e^{\lambda X}\right] = \mathbf{E}\left[e^{\lambda \sum_{i=1}^n X_i}\right] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X_i}\right]$$

3.

$$\mathbf{E}\left[e^{\lambda X_i}\right] = e^{\lambda} p_i + (1 - p_i) = 1 + p_i(e^{\lambda} - 1) \stackrel{1+x \leq e^x}{\leq} e^{p_i(e^{\lambda} - 1)}$$

4. Putting all together

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq e^{-\lambda(1+\delta)\mu} \prod_{i=1}^n e^{p_i(e^{\lambda} - 1)} = e^{-\lambda(1+\delta)\mu} e^{\mu(e^{\lambda} - 1)}$$

## Chernoff Bound: Proof

1. For  $\lambda > 0$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mu] \stackrel{e^{\lambda x} \text{ is incr}}{=} \mathbf{P}[e^{\lambda X} \geq e^{\lambda(1+\delta)\mu}] \stackrel{\text{Markov}}{\leq} e^{-\lambda(1+\delta)\mu} \mathbf{E}[e^{\lambda X}]$$

$$2. \mathbf{E}[e^{\lambda X}] = \mathbf{E}\left[e^{\lambda \sum_{i=1}^n X_i}\right] \stackrel{\text{indep}}{=} \prod_{i=1}^n \mathbf{E}[e^{\lambda X_i}]$$

3.

$$\mathbf{E}[e^{\lambda X_i}] = e^{\lambda} p_i + (1 - p_i) = 1 + p_i(e^{\lambda} - 1) \stackrel{1+x \leq e^x}{\leq} e^{p_i(e^{\lambda} - 1)}$$

4. Putting all together

$$\mathbf{P}[X \geq (1 + \delta)\mu] \leq e^{-\lambda(1+\delta)\mu} \prod_{i=1}^n e^{p_i(e^{\lambda} - 1)} = e^{-\lambda(1+\delta)\mu} e^{\mu(e^{\lambda} - 1)}$$

5. Choose  $\lambda = \log(1 + \delta) > 0$  to get the result.

## Chernoff Bounds: Lower Tails

We can also use Chernoff Bounds to show a random variable is **not too small** compared to its mean:

Chernoff Bounds (General Form, Lower Tail)

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum p_i$ . Then, for any  $\delta > 0$  it holds that

$$\mathbf{P}[X \leq (1 - \delta)\mu] \leq \left[ \frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right]^\mu,$$

and thus, by substitution, for any  $t < \mu$ ,

$$\mathbf{P}[X \leq t] \leq e^{-\mu} \left( \frac{e\mu}{t} \right)^t.$$

### Exercise on Supervision Sheet

**Hint:** multiply both sides by  $-1$  and repeat the proof of the Chernoff Bound

# Nicer Chernoff Bounds

---

## Nicer Chernoff Bounds

---

— “Nicer” Chernoff Bounds —

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then,

## Nicer Chernoff Bounds

— “Nicer” Chernoff Bounds —

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then,

- For all  $t > 0$ ,

$$\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$$

$$\mathbf{P}[X \leq \mathbf{E}[X] - t] \leq e^{-2t^2/n}$$



## Nicer Chernoff Bounds

### “Nicer” Chernoff Bounds

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then,

- For all  $t > 0$ ,

$$\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$$

$$\mathbf{P}[X \leq \mathbf{E}[X] - t] \leq e^{-2t^2/n}$$

- For  $0 < \delta < 1$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{3}\right)$$

$$\mathbf{P}[X \leq (1 - \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{2}\right)$$

## Nicer Chernoff Bounds

“Nicer” Chernoff Bounds

Suppose  $X_1, \dots, X_n$  are independent Bernoulli random variables with parameter  $p_i$ . Let  $X = X_1 + \dots + X_n$  and  $\mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$ . Then,

- For all  $t > 0$ ,

$$\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$$

$$\mathbf{P}[X \leq \mathbf{E}[X] - t] \leq e^{-2t^2/n}$$

- For  $0 < \delta < 1$ ,

$$\mathbf{P}[X \geq (1 + \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{3}\right)$$

$$\mathbf{P}[X \leq (1 - \delta)\mathbf{E}[X]] \leq \exp\left(-\frac{\delta^2 \mathbf{E}[X]}{2}\right)$$

All upper tail bounds hold even under a relaxed independence assumption:  
For all  $1 \leq i \leq n$  and  $x_1, x_2, \dots, x_{i-1} \in \{0, 1\}$ ,

$$\mathbf{P}[X_i = 1 \mid X_1 = x_1, \dots, X_{i-1} = x_{i-1}] \leq p_i.$$

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

**Application 1: Balls into Bins**

Application 2: Randomised QuickSort

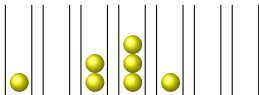
Extensions of Chernoff Bounds

Applications of Method of Bounded Differences

Appendix

## Balls into Bins

---

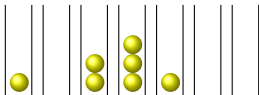


Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

## Balls into Bins

---



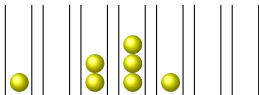
### Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

- A very natural but also rich mathematical model

## Balls into Bins

---



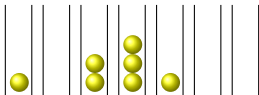
### Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

- A very natural but also rich mathematical model
- In computer science, there are several interpretations:

## Balls into Bins

---

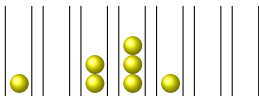


### Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked **independently and uniformly at random**.

- A very natural but also rich **mathematical** model
- In **computer science**, there are several interpretations:
  1. Bins are a hash table, balls are items
  2. Bins are processors and balls are jobs
  3. Bins are data servers and balls are queries

## Balls into Bins



### Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

- A very natural but also rich mathematical model
- In computer science, there are several interpretations:
  1. Bins are a hash table, balls are items
  2. Bins are processors and balls are jobs
  3. Bins are data servers and balls are queries

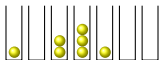


**Exercise:** Think about the relation between the Balls into Bins Model and the Coupon Collector Problem.



## Balls into Bins: Bounding the Maximum Load (1/4)

---

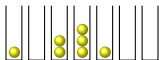


Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

## Balls into Bins: Bounding the Maximum Load (1/4)

---

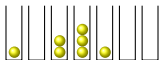


Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

## Balls into Bins: Bounding the Maximum Load (1/4)



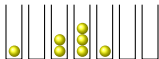
Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

- Focus on an arbitrary single bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .

## Balls into Bins: Bounding the Maximum Load (1/4)



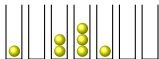
Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

- Focus on an arbitrary single bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .
- The total balls in the bin is given by  $X := \sum_{i=1}^n X_i$ .

## Balls into Bins: Bounding the Maximum Load (1/4)



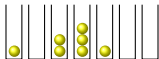
Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

- Focus on an arbitrary single bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .
- The total balls in the bin is given by  $X := \sum_{i=1}^m X_i$ .
- Since  $m = 2n \log n$ , then  $\mu = \mathbf{E}[X] = 2 \log n$

## Balls into Bins: Bounding the Maximum Load (1/4)



Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

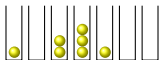
- Focus on an arbitrary single bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .
- The total balls in the bin is given by  $X := \sum_{i=1}^n X_i$ .
- Since  $m = 2n \log n$ , then  $\mu = \mathbf{E}[X] = 2 \log n$

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

- By the Chernoff Bound,

$$\mathbf{P}[X \geq 6 \log n] \leq e^{-2 \log n} \left( \frac{2e \log n}{6 \log n} \right)^{6 \log n} \leq e^{-2 \log n} = n^{-2}$$

## Balls into Bins: Bounding the Maximum Load (1/4)



Balls into Bins Model

You have  $m$  balls and  $n$  bins. Each ball is allocated in a bin picked independently and uniformly at random.

**Question 1:** How large is the maximum load if  $m = 2n \log n$ ?

- Focus on an arbitrary single bin. Let  $X_i$  the indicator variable which is 1 iff ball  $i$  is assigned to this bin. Note that  $p_i = \mathbf{P}[X_i = 1] = 1/n$ .
- The total balls in the bin is given by  $X := \sum_{i=1}^n X_i$ .
- Since  $m = 2n \log n$ , then  $\mu = \mathbf{E}[X] = 2 \log n$

here we could have used the “nicer” bounds as well!

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

- By the Chernoff Bound,

$$\mathbf{P}[X \geq 6 \log n] \leq e^{-2 \log n} \left( \frac{2e \log n}{6 \log n} \right)^{6 \log n} \leq e^{-2 \log n} = n^{-2}$$

## Balls into Bins: Bounding the Maximum Load (2/4)

---

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.



## Balls into Bins: Bounding the Maximum Load (2/4)

---

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$

## Balls into Bins: Bounding the Maximum Load (2/4)

---

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$
- By the **Union Bound**,

$$\mathbf{P} \left[ \bigcup_{j=1}^n \mathcal{E}_j \right] \leq \sum_{j=1}^n \mathbf{P}[\mathcal{E}_j] \leq n \cdot n^{-2} = n^{-1}.$$

## Balls into Bins: Bounding the Maximum Load (2/4)

---

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$
- By the **Union Bound**,

$$\mathbf{P} \left[ \bigcup_{j=1}^n \mathcal{E}_j \right] \leq \sum_{j=1}^n \mathbf{P}[\mathcal{E}_j] \leq n \cdot n^{-2} = n^{-1}.$$

- Therefore **whp**, no bin receives at least  $6 \log n$  balls

## Balls into Bins: Bounding the Maximum Load (2/4)

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$
- By the **Union Bound**,

$$\mathbf{P} \left[ \bigcup_{j=1}^n \mathcal{E}_j \right] \leq \sum_{j=1}^n \mathbf{P}[\mathcal{E}_j] \leq n \cdot n^{-2} = n^{-1}.$$

- Therefore **whp**, no bin receives at least  $6 \log n$  balls

*whp* stands for *with high probability*:

An event  $\mathcal{E}$  (that implicitly depends on an input parameter  $n$ ) occurs **whp** if

$$\mathbf{P}[\mathcal{E}] \rightarrow 1 \text{ as } n \rightarrow \infty.$$

This is a very standard notation in randomised algorithms  
but it may vary from author to author. **Be careful!**

## Balls into Bins: Bounding the Maximum Load (2/4)

- Let  $\mathcal{E}_j := \{X(j) \geq 6 \log n\}$ , that is, bin  $j$  receives at least  $6 \log n$  balls.
- We are interested in the probability that **at least** one bin receives at least  $6 \log n$  balls  $\Rightarrow$  this is the event  $\bigcup_{j=1}^n \mathcal{E}_j$
- By the **Union Bound**,

$$\mathbf{P} \left[ \bigcup_{j=1}^n \mathcal{E}_j \right] \leq \sum_{j=1}^n \mathbf{P}[\mathcal{E}_j] \leq n \cdot n^{-2} = n^{-1}.$$

- Therefore **whp**, no bin receives at least  $6 \log n$  balls
- By **pigeonhole principle**, the max loaded bin receives at least  $2 \log n$  balls. Hence our bound is pretty sharp.

*whp* stands for *with high probability*:

An event  $\mathcal{E}$  (that implicitly depends on an input parameter  $n$ ) occurs **whp** if

$$\mathbf{P}[\mathcal{E}] \rightarrow 1 \text{ as } n \rightarrow \infty.$$

This is a very standard notation in randomised algorithms but it may vary from author to author. **Be careful!**

## Balls into Bins: Bounding the Maximum Load (3/4)

---

**Question 2:** How large is the maximum load if  $m = n$ ?

## Balls into Bins: Bounding the Maximum Load (3/4)

**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

## Balls into Bins: Bounding the Maximum Load (3/4)

**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .



**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:

$$\mathbf{P}[X \geq t] \leq e^{-\mu} (e\mu/t)^t$$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .
- Indeed:

$$\left(\frac{e \log \log n}{4 \log n}\right)^{4 \log n / \log \log n} = \exp\left(\frac{4 \log n}{\log \log n} \cdot \log\left(\frac{e \log \log n}{4 \log n}\right)\right)$$

**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:

$$\mathbf{P}[X \geq t] \leq e^{-\mu}(e\mu/t)^t$$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .
- Indeed:

$$\left(\frac{e \log \log n}{4 \log n}\right)^{4 \log n / \log \log n} = \exp\left(\frac{4 \log n}{\log \log n} \cdot \log\left(\frac{e \log \log n}{4 \log n}\right)\right)$$

- The term inside the exponential is

$$\frac{4 \log n}{\log \log n} \cdot (\log(4/e) + \log \log \log n - \log \log n)$$

## Balls into Bins: Bounding the Maximum Load (3/4)

**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:

$$\mathbf{P}[X \geq t] \leq e^{-\mu}(e\mu/t)^t$$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .
- Indeed:

$$\left(\frac{e \log \log n}{4 \log n}\right)^{4 \log n / \log \log n} = \exp\left(\frac{4 \log n}{\log \log n} \cdot \log\left(\frac{e \log \log n}{4 \log n}\right)\right)$$

- The term inside the exponential is

$$\frac{4 \log n}{\log \log n} \cdot (\log(4/e) + \log \log \log n - \log \log n) \leq \frac{4 \log n}{\log \log n} \left(-\frac{1}{2} \log \log n\right),$$

This inequality only works for large enough  $n$ .

**Question 2:** How large is the maximum load if  $m = n$ ?

- Using the Chernoff Bound:  $\mathbf{P}[X \geq t] \leq e^{-\mu}(e\mu/t)^t$

$$\mathbf{P}[X \geq t] \leq e^{-1} \left(\frac{e}{t}\right)^t \leq \left(\frac{e}{t}\right)^t$$

- By setting  $t = 4 \log n / \log \log n$ , we claim to obtain  $\mathbf{P}[X \geq t] \leq n^{-2}$ .
- Indeed:

$$\left(\frac{e \log \log n}{4 \log n}\right)^{4 \log n / \log \log n} = \exp\left(\frac{4 \log n}{\log \log n} \cdot \log\left(\frac{e \log \log n}{4 \log n}\right)\right)$$

- The term inside the exponential is

$$\frac{4 \log n}{\log \log n} \cdot (\log(4/e) + \log \log \log n - \log \log n) \leq \frac{4 \log n}{\log \log n} \left(-\frac{1}{2} \log \log n\right),$$

obtaining that  $\mathbf{P}[X \geq t] \leq n^{-4/2} = n^{-2}$ .

This inequality only works for large enough  $n$ .

We just proved that

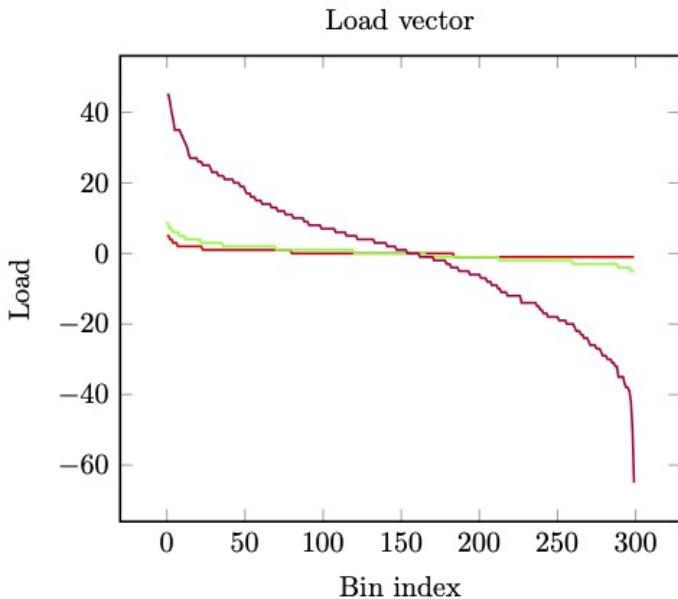
$$\mathbf{P}[X \geq 4 \log n / \log \log n] \leq n^{-2},$$

thus by the **Union Bound**, no bin receives more than  $\Omega(\log n / \log \log n)$  balls with probability at least  $1 - 1/n$ .  $\square$

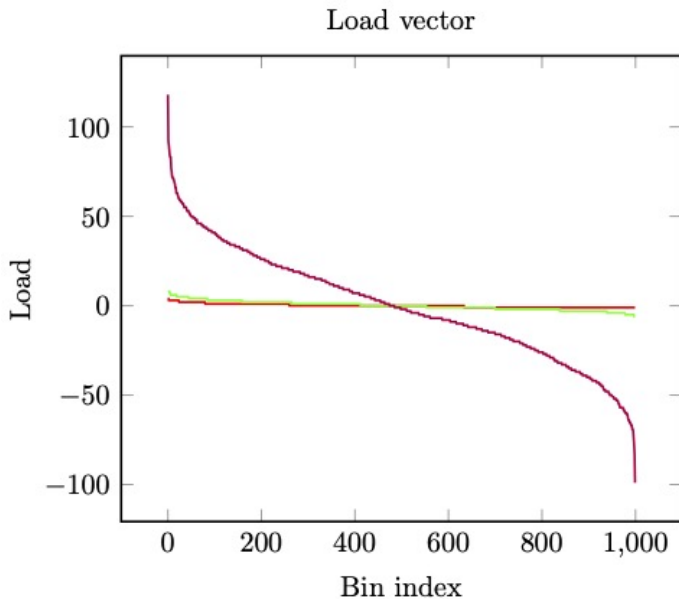
- We plot the load configuration for  $m \in \{n, n \log n, n^2\}$
- We consider  $n \in \{300, 1000, 100000\}$
- In plots, we take the **normalised load**, that is, actual bin load minus average load

Acknowledgements: experiments and plots created by Dimitris Los

## Balls-into-Bins Plot (1/3)

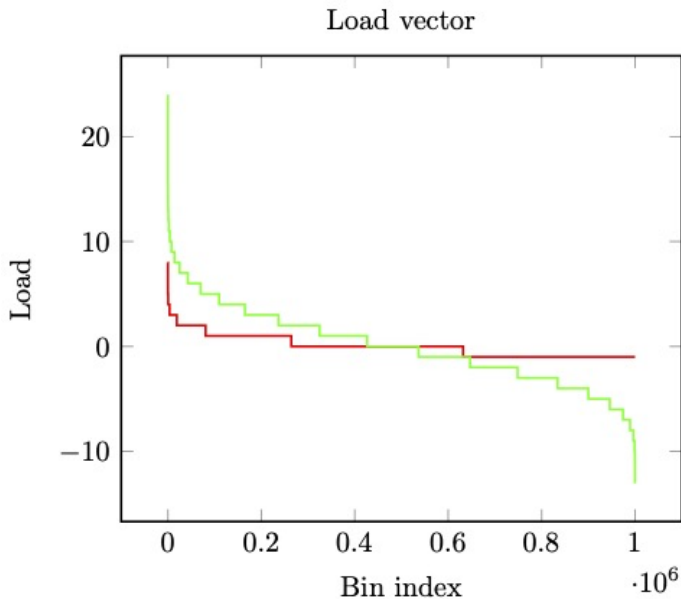


## Balls-into-Bins Plot (2/3)





## Balls-into-Bins Plot (3/3) (only $m \in \{n, n \log n\}$ )



## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**

## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$

## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$
- For the case  $m = n$ , the algorithm is **not good**, since the maximum load is whp.  $\Theta(\log n / \log \log n)$ , while the average load is 1.

## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$
- For the case  $m = n$ , the algorithm is **not good**, since the maximum load is whp.  $\Theta(\log n / \log \log n)$ , while the average load is 1.

### A Better Load Balancing Approach

For any  $m \geq n$ , we can improve the balls into bin process by sampling **two bins** in each step, then assigning the ball into the bin with lesser load.

## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$
- For the case  $m = n$ , the algorithm is **not good**, since the maximum load is whp.  $\Theta(\log n / \log \log n)$ , while the average load is 1.

### A Better Load Balancing Approach

For any  $m \geq n$ , we can improve the balls into bin process by sampling **two bins** in each step, then assigning the ball into the bin with lesser load.  
 $\Rightarrow$  gives a (normalised) maximum load  $\Theta(\log \log n)$  w.p.  $1 - 1/n$ .

## Conclusions

---

- If the number of balls is  $2 \log n$  times  $n$  (the number of bins), then to distribute balls at random is a **good algorithm**
  - This is because the worst case maximum load is whp.  $6 \log n$ , while the average load is  $2 \log n$
- For the case  $m = n$ , the algorithm is **not good**, since the maximum load is whp.  $\Theta(\log n / \log \log n)$ , while the average load is 1.

### A Better Load Balancing Approach

For any  $m \geq n$ , we can improve the balls into bin process by sampling **two bins** in each step, then assigning the ball into the bin with lesser load.  
 $\Rightarrow$  gives a (normalised) maximum load  $\Theta(\log \log n)$  w.p.  $1 - 1/n$ .

This is called the **power of two choices**: It is a common technique to improve the performance of randomised algorithms.

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

Application 1: Balls into Bins

**Application 2: Randomised QuickSort**

Extensions of Chernoff Bounds

Applications of Method of Bounded Differences

Appendix



QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10:    **return**  $A$

QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10:    **return**  $A$

- **Example:** Let  $A = (2, 8, 9, 1, 7, 5, 6, 3, 4)$  with  $A[7] = 6$  as pivot.

QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10:    **return**  $A$

- **Example:** Let  $A = (2, 8, 9, 1, 7, 5, 6, 3, 4)$  with  $A[7] = 6$  as pivot.  
⇒  $A_1 = (2, 1, 5, 3, 4)$  and  $A_2 = (8, 9, 7)$

QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10:    **return**  $A$

- **Example:** Let  $A = (2, 8, 9, 1, 7, 5, 6, 3, 4)$  with  $A[7] = 6$  as pivot.  
⇒  $A_1 = (2, 1, 5, 3, 4)$  and  $A_2 = (8, 9, 7)$
- **Worst-Case Complexity** (number of comparisons) is  $\Theta(n^2)$ ,

QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10:  **return**  $A$

- **Example:** Let  $A = (2, 8, 9, 1, 7, 5, 6, 3, 4)$  with  $A[7] = 6$  as pivot.  
⇒  $A_1 = (2, 1, 5, 3, 4)$  and  $A_2 = (8, 9, 7)$
- **Worst-Case Complexity** (number of comparisons) is  $\Theta(n^2)$ ,  
while **Average-Case Complexity** is  $O(n \log n)$ .

QUICKSORT (Input  $A[1], A[2], \dots, A[n]$ )

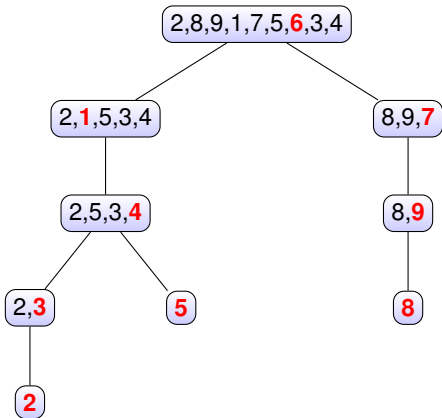
- 1: Pick an element from the array, the so-called **pivot**
- 2: **If**  $|A| = 0$  or  $|A| = 1$  **then**
- 3:     **return**  $A$
- 4: **else**
- 5:     Create two subarrays  $A_1$  and  $A_2$  (without the pivot) such that:
- 6:          $A_1$  contains the elements that are **smaller than the pivot**
- 7:          $A_2$  contains the elements that are **greater (or equal) than the pivot**
- 8:     QUICKSORT( $A_1$ )
- 9:     QUICKSORT( $A_2$ )
- 10: **return**  $A$

- **Example:** Let  $A = (2, 8, 9, 1, 7, 5, 6, 3, 4)$  with  $A[7] = 6$  as pivot.  
⇒  $A_1 = (2, 1, 5, 3, 4)$  and  $A_2 = (8, 9, 7)$
- **Worst-Case Complexity** (number of comparisons) is  $\Theta(n^2)$ ,  
while **Average-Case Complexity** is  $O(n \log n)$ .

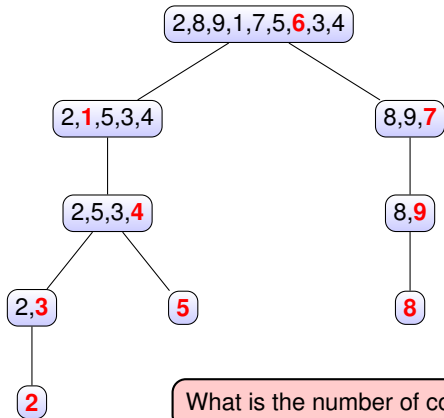
We will now give a proof of this “well-known” result!

## QuickSort: How to Count Comparisons

---

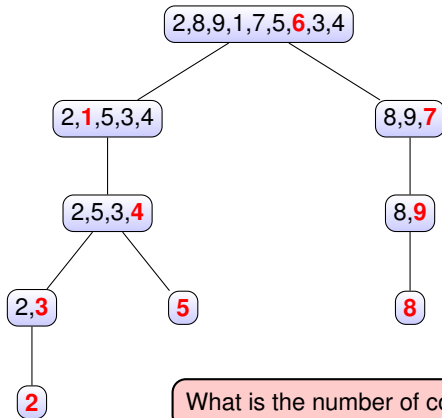


## QuickSort: How to Count Comparisons



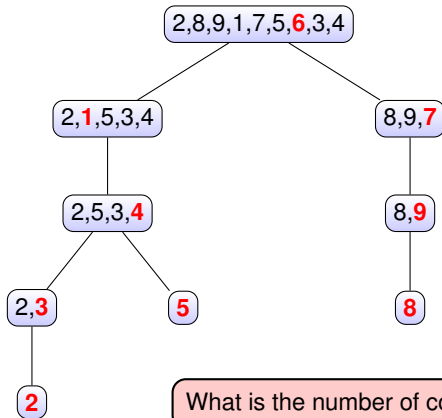


## QuickSort: How to Count Comparisons



Note that the **number of comparison** by QUICKSORT is equivalent to the **sum of the height** of all nodes in the tree (why?).

## QuickSort: How to Count Comparisons



What is the number of comparisons?

Note that the number of comparison by QUICKSORT is equivalent to the sum of the height of all nodes in the tree (why?). In this case:

$$0 + 1 + 1 + 2 + 2 + 3 + 3 + 3 + 4 = 19.$$

## Randomised QuickSort: Analysis (1/4)

---

How to pick a good pivot? We don't, **just pick one at random.**

## Randomised QuickSort: Analysis (1/4)

---

How to pick a good pivot? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

## Randomised QuickSort: Analysis (1/4)

---

How to pick a good pivot? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

Let us analyse QUICKSORT with random pivots.

## Randomised QuickSort: Analysis (1/4)

---

How to pick a good pivot? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

Let us analyse QUICKSORT with random pivots.

1. Assume  $A$  consists of  $n$  different numbers, w.l.o.g.,  $\{1, 2, \dots, n\}$

## Randomised QuickSort: Analysis (1/4)

---

How to pick a **good pivot**? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

Let us analyse QUICKSORT with **random** pivots.

1. Assume  $A$  consists of  $n$  different numbers, w.l.o.g.,  $\{1, 2, \dots, n\}$
2. Let  $H_i$  be the **deepest level** where element  $i$  appears in the tree.  
Then the number of comparison is  $H = \sum_{i=1}^n H_i$

## Randomised QuickSort: Analysis (1/4)

---

How to pick a **good pivot**? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

Let us analyse QUICKSORT with **random** pivots.

1. Assume  $A$  consists of  $n$  different numbers, w.l.o.g.,  $\{1, 2, \dots, n\}$
2. Let  $H_i$  be the **deepest level** where element  $i$  appears in the tree.  
Then the number of comparison is  $H = \sum_{i=1}^n H_i$
3. We will prove that exists  $C > 0$  such that

$$\mathbf{P}[H \leq Cn \log n] \geq 1 - n^{-1}.$$



## Randomised QuickSort: Analysis (1/4)

How to pick a good pivot? We don't, **just pick one at random.**

This should be your standard answer in this course 😊

Let us analyse QUICKSORT with random pivots.

1. Assume  $A$  consists of  $n$  different numbers, w.l.o.g.,  $\{1, 2, \dots, n\}$
2. Let  $H_i$  be the deepest level where element  $i$  appears in the tree.  
Then the number of comparison is  $H = \sum_{i=1}^n H_i$
3. We will prove that exists  $C > 0$  such that

$$\mathbf{P}[H \leq Cn \log n] \geq 1 - n^{-1}.$$

4. Actually, we will prove sth slightly stronger:

$$\mathbf{P}\left[\bigcap_{i=1}^n \{H_i \leq C \log n\}\right] \geq 1 - n^{-1}.$$

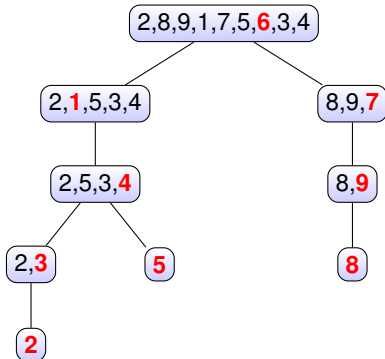
## Randomised QuickSort: Analysis (2/4)

---

- Let  $P$  be a path from the root to the deepest level of some element

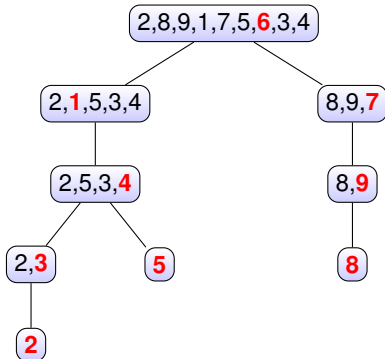
## Randomised QuickSort: Analysis (2/4)

- Let  $P$  be a path from the root to the deepest level of some element



## Randomised QuickSort: Analysis (2/4)

- Let  $P$  be a path from the root to the deepest level of some element

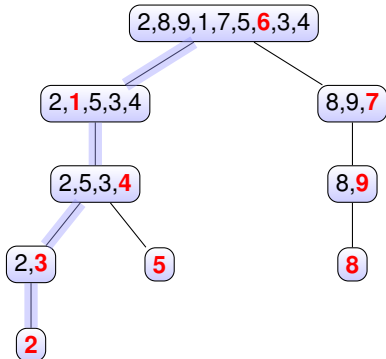


- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$



## Randomised QuickSort: Analysis (2/4)

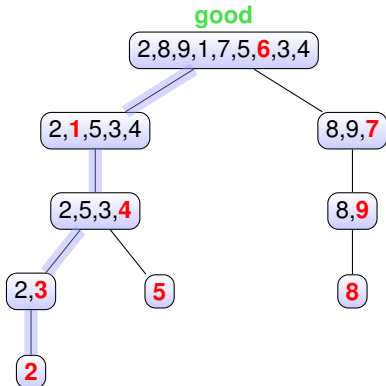
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

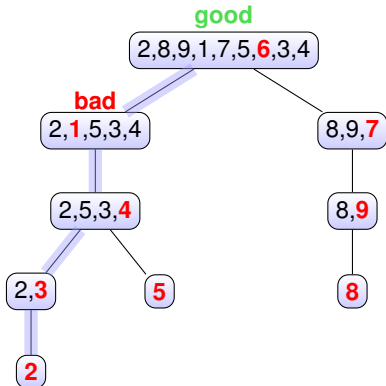
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**

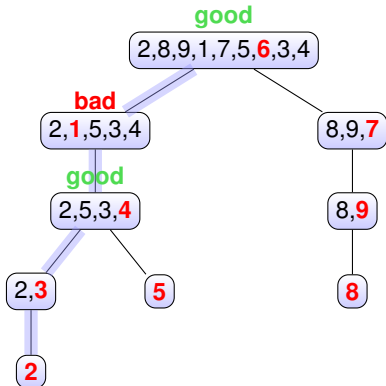


- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$



## Randomised QuickSort: Analysis (2/4)

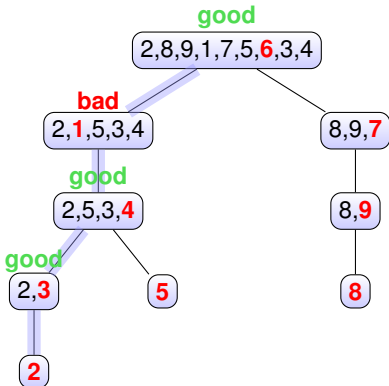
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

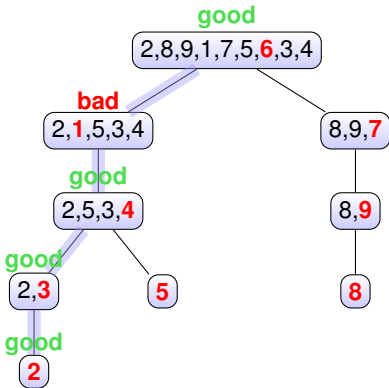
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

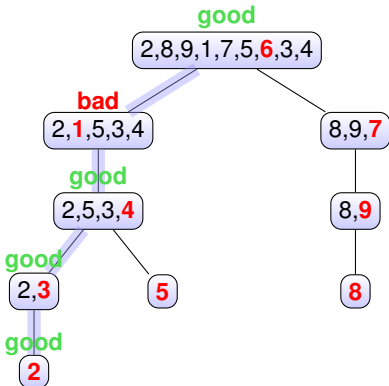
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

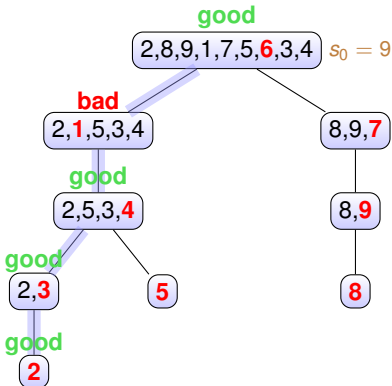
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

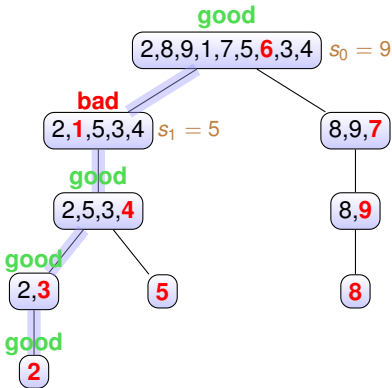
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

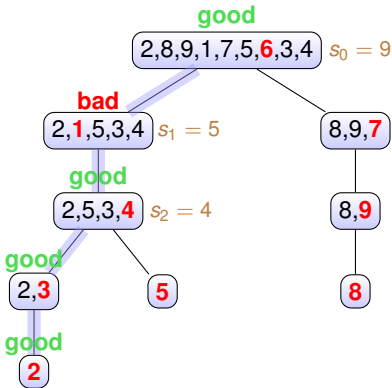
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

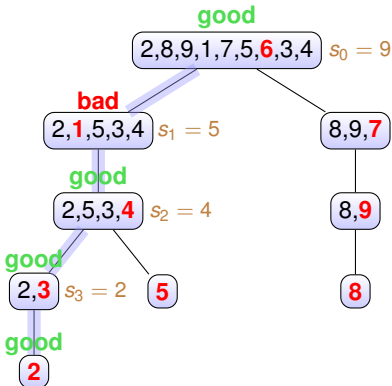
- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (2/4)

- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .

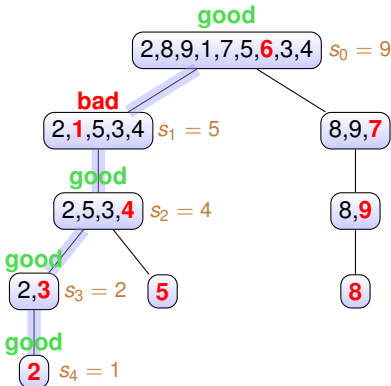


- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$



## Randomised QuickSort: Analysis (2/4)

- Let  $P$  be a path from the root to the deepest level of some element
  - A node in  $P$  is called **good** if the corresponding pivot partitions the array into two subarrays each of size at most  $2/3$  of the previous one
  - otherwise, the node is **bad**
- Further let  $s_t$  be the **size** of the array at level  $t$  in  $P$ .



- Element 2:  $(2, 8, 9, 1, 7, 5, 6, 3, 4) \rightarrow (2, 1, 5, 3, 4) \rightarrow (2, 5, 3, 4) \rightarrow (2, 3) \rightarrow (2)$

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .

## Randomised QuickSort: Analysis (3/4)

---

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
- Second Case, **bad** node:  $s_{k+1} \leq s_k$ .

## Randomised QuickSort: Analysis (3/4)

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
  - First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
  - Second Case, **bad** node:  $s_{k+1} \leq s_k$ .
- $\Rightarrow$  There are at most  $T = \frac{\log n}{\log(3/2)} < 3 \log n$  many **good** nodes on any path  $P$ .



## Randomised QuickSort: Analysis (3/4)

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
- Second Case, **bad** node:  $s_{k+1} \leq s_k$ .

This even holds always,  
i.e., deterministically!

$\Rightarrow$  There are at most  $T = \frac{\log n}{\log(3/2)} < 3 \log n$  many **good** nodes on any path  $P$ .

## Randomised QuickSort: Analysis (3/4)

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
- Second Case, **bad** node:  $s_{k+1} \leq s_k$ .

This even holds always,  
i.e., deterministically!

- ⇒ There are at most  $T = \frac{\log n}{\log(3/2)} < 3 \log n$  many **good** nodes on any path  $P$ .
- Assume  $|P| \geq C \log n$  for  $C := 24$

## Randomised QuickSort: Analysis (3/4)

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
- Second Case, **bad** node:  $s_{k+1} \leq s_k$ .

This even holds always,  
i.e., deterministically!

- ⇒ There are at most  $T = \frac{\log n}{\log(3/2)} < 3 \log n$  many **good** nodes on any path  $P$ .
- Assume  $|P| \geq C \log n$  for  $C := 24$   
⇒ number of **bad** vertices in the first  $24 \log n$  levels is more than  $21 \log n$ .

## Randomised QuickSort: Analysis (3/4)

- Consider now any element  $i \in \{1, 2, \dots, n\}$  and construct the path  $P = P(i)$  one level by one
- For  $P$  to proceed from level  $k$  to  $k + 1$ , the condition  $s_k > 1$  is necessary

How far could such a path  $P$  possibly run until we have  $s_k = 1$ ?

- We start with  $s_0 = n$
- First Case, **good** node:  $s_{k+1} \leq \frac{2}{3} \cdot s_k$ .
- Second Case, **bad** node:  $s_{k+1} \leq s_k$ .

This even holds always,  
i.e., deterministically!

⇒ There are at most  $T = \frac{\log n}{\log(3/2)} < 3 \log n$  many **good** nodes on any path  $P$ .

- Assume  $|P| \geq C \log n$  for  $C := 24$

⇒ number of **bad** vertices in the first  $24 \log n$  levels is more than  $21 \log n$ .

Let us now upper bound the probability that this “bad event” happens!

## Randomised QuickSort: Analysis (4/4)

---

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .

## Randomised QuickSort: Analysis (4/4)

---

- Consider the first  $24 \log n$  vertices of  $P$  to the **deepest level** of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :

## Randomised QuickSort: Analysis (4/4)

---

- Consider the first  $24 \log n$  vertices of  $P$  to the **deepest level** of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.

## Randomised QuickSort: Analysis (4/4)

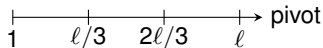
---

- Consider the first  $24 \log n$  vertices of  $P$  to the **deepest level** of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$



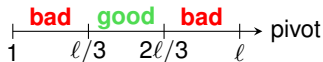
## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$



## Randomised QuickSort: Analysis (4/4)

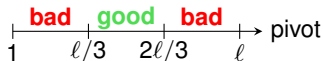
- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$



## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :

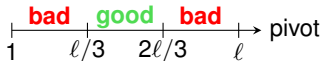
- $X_j = 1$  if the node at level  $j$  is **bad**
- $X_j = 0$  if the node at level  $j$  is **good**.



- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)

## Randomised QuickSort: Analysis (4/4)

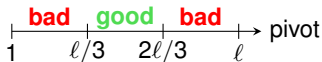
- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



**Question:** But what if the path  $P$  does not reach level  $j$ ?

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



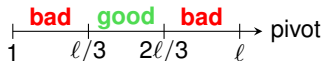
**Question:** But what if the path  $P$  does not reach level  $j$ ?

**Answer:** We can then simply define  $X_j$  as the result of an independent coin flip with probability  $2/3$ .

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :

- $X_j = 1$  if the node at level  $j$  is **bad**
- $X_j = 0$  if the node at level  $j$  is **good**.

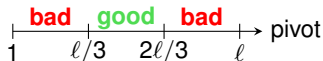


- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :

- $X_j = 1$  if the node at level  $j$  is **bad**
- $X_j = 0$  if the node at level  $j$  is **good**.

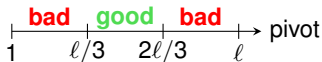


- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)

We can now apply the “nicer” **Chernoff Bound!**

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



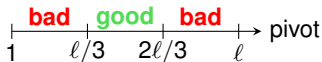
We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$



## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)

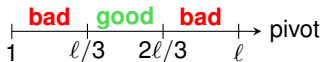


We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)

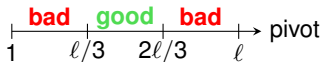


We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



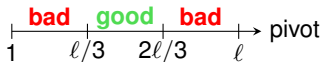
We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\mathbf{P}[X > 21 \log n] \leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n]$$

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



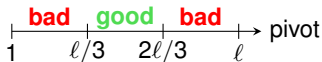
We can now apply the “nicer” **Chernoff Bound**!

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\mathbf{P}[X > 21 \log n] \leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)}$$

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



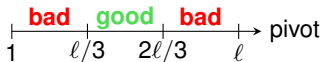
We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\begin{aligned} \mathbf{P}[X > 21 \log n] &\leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)} \\ &= e^{-(50/24) \log n} \end{aligned}$$

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



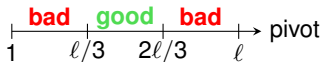
We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\begin{aligned} \mathbf{P}[X > 21 \log n] &\leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)} \\ &= e^{-(50/24) \log n} \leq n^{-2}. \end{aligned}$$

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



We can now apply the “nicer” **Chernoff Bound!**

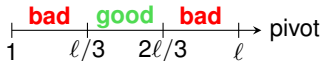
- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\begin{aligned} \mathbf{P}[X > 21 \log n] &\leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)} \\ &= e^{-(50/24) \log n} \leq n^{-2}. \end{aligned}$$

- Hence  $P$  has more than  $24 \log n$  nodes with probability at most  $n^{-2}$ .

## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

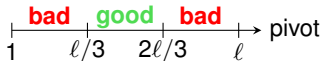
$$\begin{aligned}\mathbf{P}[X > 21 \log n] &\leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)} \\ &= e^{-(50/24) \log n} \leq n^{-2}.\end{aligned}$$

- Hence  $P$  has more than  $24 \log n$  nodes with probability at most  $n^{-2}$ .
- As there are in total  $n$  paths, by the union bound, the probability that at least one of them has more than  $24 \log n$  nodes is at most  $n^{-1}$ .



## Randomised QuickSort: Analysis (4/4)

- Consider the first  $24 \log n$  vertices of  $P$  to the deepest level of element  $i$ .
- For any level  $j \in \{0, 1, \dots, 24 \log n - 1\}$ , define an indicator variable  $X_j$ :
  - $X_j = 1$  if the node at level  $j$  is **bad**
  - $X_j = 0$  if the node at level  $j$  is **good**.
- $\mathbf{P}[X_j = 1 \mid X_0 = x_0, \dots, X_{j-1} = x_{j-1}] \leq \frac{2}{3}$
- $X := \sum_{j=0}^{24 \log n - 1} X_j$  satisfies relaxed independence assumption (slide 16)



We can now apply the “nicer” **Chernoff Bound!**

- We have  $\mathbf{E}[X] \leq (2/3) \cdot 24 \log n = 16 \log n$
- Then, by the “nicer” Chernoff Bounds  $\mathbf{P}[X \geq \mathbf{E}[X] + t] \leq e^{-2t^2/n}$

$$\begin{aligned}\mathbf{P}[X > 21 \log n] &\leq \mathbf{P}[X > \mathbf{E}[X] + 5 \log n] \leq e^{-2(5 \log n)^2 / (24 \log n)} \\ &= e^{-(50/24) \log n} \leq n^{-2}.\end{aligned}$$

- Hence  $P$  has more than  $24 \log n$  nodes with probability at most  $n^{-2}$ .
- As there are in total  $n$  paths, by the union bound, the probability that at least one of them has more than  $24 \log n$  nodes is at most  $n^{-1}$ .  $\square$

## Randomised QuickSort: Final Remarks

---

- Well-known: any comparison-based sorting algorithm needs  $\Omega(n \log n)$

## Randomised QuickSort: Final Remarks

---

- Well-known: any comparison-based sorting algorithm needs  $\Omega(n \log n)$
- A classical result: **expected number** of comparison of **randomised QUICKSORT** is  $2n \log n + O(n)$  (see, e.g., book by Mitzenmacher & Upfal)

## Randomised QuickSort: Final Remarks

---

- Well-known: any comparison-based sorting algorithm needs  $\Omega(n \log n)$
- A classical result: **expected number** of comparison of **randomised QUICKSORT** is  $2n \log n + O(n)$  (see, e.g., book by Mitzenmacher & Upfal)

**Supervision Exercise:** Our upper bound of  $O(n \log n)$  **whp** also immediately implies a  $O(n \log n)$  bound on the expected number of comparisons!

## Randomised QuickSort: Final Remarks

---

- Well-known: any comparison-based sorting algorithm needs  $\Omega(n \log n)$
- A classical result: **expected number** of comparison of **randomised QUICKSORT** is  $2n \log n + O(n)$  (see, e.g., book by Mitzenmacher & Upfal)

**Supervision Exercise:** Our upper bound of  $O(n \log n)$  **whp** also immediately implies a  $O(n \log n)$  bound on the expected number of comparisons!

- It is possible to **deterministically** find the best pivot element that divides the array into two subarrays of the same size.
- The latter requires to compute the **median** of the array in linear time, which is not easy...
- The presented **randomised** algorithm for QUICKSORT is much **easier to implement!**

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

Application 1: Balls into Bins

Application 2: Randomised QuickSort

**Extensions of Chernoff Bounds**

Applications of Method of Bounded Differences

Appendix

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.
- Unfortunately the distribution of the  $X_i$  may be unknown or hard to compute, thus it will be hard to compute the moment-generating function.



## Hoeffding's Extension

---

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.
- Unfortunately the distribution of the  $X_i$  may be unknown or hard to compute, thus it will be hard to compute the moment-generating function.
- Hoeffding's Lemma helps us here:

## Hoeffding's Extension

---

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.
- Unfortunately the distribution of the  $X_i$  may be unknown or hard to compute, thus it will be hard to compute the moment-generating function.
- Hoeffding's Lemma helps us here:

— Hoeffding's Extension Lemma —

Let  $X$  be a random variable with mean 0 such that  $a \leq X \leq b$ . Then for all  $\lambda \in \mathbb{R}$ ,

$$\mathbf{E} \left[ e^{\lambda X} \right] \leq \exp \left( \frac{(b-a)^2 \lambda^2}{8} \right)$$

## Hoeffding's Extension

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.
- Unfortunately the distribution of the  $X_i$  may be unknown or hard to compute, thus it will be hard to compute the moment-generating function.
- Hoeffding's Lemma helps us here:

You can always consider  
 $X' = X - \mathbf{E}[X]$

Hoeffding's Extension Lemma

Let  $X$  be a random variable with mean 0 such that  $a \leq X \leq b$ . Then for all  $\lambda \in \mathbb{R}$ ,

$$\mathbf{E} \left[ e^{\lambda X} \right] \leq \exp \left( \frac{(b-a)^2 \lambda^2}{8} \right)$$

## Hoeffding's Extension

- Besides **sums of independent bernoulli** random variables, **sums of independent and bounded** random variables are very frequent in applications.
- Unfortunately the distribution of the  $X_i$  may be unknown or hard to compute, thus it will be hard to compute the moment-generating function.
- Hoeffding's Lemma helps us here:

You can always consider  
 $X' = X - \mathbf{E}[X]$

Hoeffding's Extension Lemma

Let  $X$  be a random variable with mean 0 such that  $a \leq X \leq b$ . Then for all  $\lambda \in \mathbb{R}$ ,

$$\mathbf{E} \left[ e^{\lambda X} \right] \leq \exp \left( \frac{(b-a)^2 \lambda^2}{8} \right)$$

We omit the proof of this lemma!

## Hoeffding Bounds

### Hoeffding's Inequality

Let  $X_1, \dots, X_n$  be independent random variable with mean  $\mu_i$  such that  $a_i \leq X_i \leq b_i$ . Let  $X = X_1 + \dots + X_n$ , and let  $\mu = \mathbf{E}[X] = \sum_{i=1}^n \mu_i$ . Then for any  $t > 0$

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

## Hoeffding Bounds

### Hoeffding's Inequality

Let  $X_1, \dots, X_n$  be independent random variable with mean  $\mu_i$  such that  $a_i \leq X_i \leq b_i$ . Let  $X = X_1 + \dots + X_n$ , and let  $\mu = \mathbf{E}[X] = \sum_{i=1}^n \mu_i$ . Then for any  $t > 0$

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

Proof Outline (skipped):

- Let  $X'_i = X_i - \mu_i$  and  $X' = X'_1 + \dots + X'_n$ , then  $\mathbf{P}[X \geq \mu + t] = \mathbf{P}[X' \geq t]$

## Hoeffding Bounds

### Hoeffding's Inequality

Let  $X_1, \dots, X_n$  be independent random variable with mean  $\mu_i$  such that  $a_i \leq X_i \leq b_i$ . Let  $X = X_1 + \dots + X_n$ , and let  $\mu = \mathbf{E}[X] = \sum_{i=1}^n \mu_i$ . Then for any  $t > 0$

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

### Proof Outline (skipped):

- Let  $X'_i = X_i - \mu_i$  and  $X' = X'_1 + \dots + X'_n$ , then  $\mathbf{P}[X \geq \mu + t] = \mathbf{P}[X' \geq t]$
- $\mathbf{P}[X' \geq t] \leq e^{-\lambda t} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X'_i}\right] \leq \exp\left[-\lambda t + \frac{\lambda^2}{8} \sum_{i=1}^n (b_i - a_i)^2\right]$

## Hoeffding Bounds

### Hoeffding's Inequality

Let  $X_1, \dots, X_n$  be independent random variable with mean  $\mu_i$  such that  $a_i \leq X_i \leq b_i$ . Let  $X = X_1 + \dots + X_n$ , and let  $\mu = \mathbf{E}[X] = \sum_{i=1}^n \mu_i$ . Then for any  $t > 0$

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

### Proof Outline (skipped):

- Let  $X'_i = X_i - \mu_i$  and  $X' = X'_1 + \dots + X'_n$ , then  $\mathbf{P}[X \geq \mu + t] = \mathbf{P}[X' \geq t]$
- $\mathbf{P}[X' \geq t] \leq e^{-\lambda t} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X'_i}\right] \leq \exp\left[-\lambda t + \frac{\lambda^2}{8} \sum_{i=1}^n (b_i - a_i)^2\right]$
- Choose  $\lambda = \frac{4t}{\sum_{i=1}^n (b_i - a_i)^2}$  to get the result.



## Hoeffding Bounds

### Hoeffding's Inequality

Let  $X_1, \dots, X_n$  be independent random variable with mean  $\mu_i$  such that  $a_i \leq X_i \leq b_i$ . Let  $X = X_1 + \dots + X_n$ , and let  $\mu = \mathbf{E}[X] = \sum_{i=1}^n \mu_i$ . Then for any  $t > 0$

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2}\right).$$

### Proof Outline (skipped):

- Let  $X'_i = X_i - \mu_i$  and  $X' = X'_1 + \dots + X'_n$ , then  $\mathbf{P}[X \geq \mu + t] = \mathbf{P}[X' \geq t]$
- $\mathbf{P}[X' \geq t] \leq e^{-\lambda t} \prod_{i=1}^n \mathbf{E}\left[e^{\lambda X'_i}\right] \leq \exp\left[-\lambda t + \frac{\lambda^2}{8} \sum_{i=1}^n (b_i - a_i)^2\right]$
- Choose  $\lambda = \frac{4t}{\sum_{i=1}^n (b_i - a_i)^2}$  to get the result.

This is not magic! you just need to optimise  $\lambda$ !

## Method of Bounded Differences

---

— Framework —

Suppose, we have **independent** random variables  $X_1, \dots, X_n$ . We want to study the random variable:

$$f(X_1, \dots, X_n)$$

Framework

Suppose, we have **independent** random variables  $X_1, \dots, X_n$ . We want to study the random variable:

$$f(X_1, \dots, X_n)$$

Some examples:

1.  $X = X_1 + \dots + X_n$

## Method of Bounded Differences

---

### Framework

Suppose, we have independent random variables  $X_1, \dots, X_n$ . We want to study the random variable:

$$f(X_1, \dots, X_n)$$

Some examples:

1.  $X = X_1 + \dots + X_n$
2. In balls into bins,  $X_i$  indicates where ball  $i$  is allocated, and  $f(X_1, \dots, X_m)$  is the number of empty bins

## Method of Bounded Differences

---

Framework

Suppose, we have independent random variables  $X_1, \dots, X_n$ . We want to study the random variable:

$$f(X_1, \dots, X_n)$$

Some examples:

1.  $X = X_1 + \dots + X_n$
2. In balls into bins,  $X_i$  indicates where ball  $i$  is allocated, and  $f(X_1, \dots, X_m)$  is the number of empty bins
3.  $X_i$  indicates if the  $i$ -th edge is present in a graph, and  $f(X_1, \dots, X_m)$  represents the number of connected components of  $G$

## Method of Bounded Differences

---

### Framework

Suppose, we have independent random variables  $X_1, \dots, X_n$ . We want to study the random variable:

$$f(X_1, \dots, X_n)$$

Some examples:

1.  $X = X_1 + \dots + X_n$
2. In balls into bins,  $X_i$  indicates where ball  $i$  is allocated, and  $f(X_1, \dots, X_m)$  is the number of empty bins
3.  $X_i$  indicates if the  $i$ -th edge is present in a graph, and  $f(X_1, \dots, X_m)$  represents the number of connected components of  $G$

In all those cases (and more) we can easily prove concentration of  $f(X_1, \dots, X_n)$  around its mean by the so-called **Method of Bounded Differences**.

## Method of Bounded Differences

---

A function  $f$  is called Lipschitz with parameters  $\mathbf{c} = (c_1, \dots, c_n)$  if for all  $i = 1, 2, \dots, n$ ,

$$|f(x_1, x_2, \dots, x_{i-1}, \mathbf{x}_i, x_{i+1}, \dots, x_n) - f(x_1, x_2, \dots, x_{i-1}, \mathbf{y}_i, x_{i+1}, \dots, x_n)| \leq c_i,$$

where  $x_i$  and  $y_i$  are in the domain of the  $i$ -th coordinate.

## Method of Bounded Differences

A function  $f$  is called **Lipschitz with parameters**  $\mathbf{c} = (c_1, \dots, c_n)$  if for all  $i = 1, 2, \dots, n$ ,

$$|f(x_1, x_2, \dots, x_{i-1}, \mathbf{x}_i, x_{i+1}, \dots, x_n) - f(x_1, x_2, \dots, x_{i-1}, \mathbf{y}_i, x_{i+1}, \dots, x_n)| \leq c_i,$$

where  $x_i$  and  $y_i$  are in the domain of the  $i$ -th coordinate.

McDiarmid's inequality

Let  $X_1, \dots, X_n$  be **independent** random variables. Let  $f$  be **Lipschitz** with parameters  $\mathbf{c} = (c_1, \dots, c_n)$ . Let  $X = f(X_1, \dots, X_n)$ . Then for any  $t > 0$ ,

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right).$$



## Method of Bounded Differences

A function  $f$  is called **Lipschitz with parameters**  $\mathbf{c} = (c_1, \dots, c_n)$  if for all  $i = 1, 2, \dots, n$ ,

$$|f(x_1, x_2, \dots, x_{i-1}, \mathbf{x}_i, x_{i+1}, \dots, x_n) - f(x_1, x_2, \dots, x_{i-1}, \mathbf{y}_i, x_{i+1}, \dots, x_n)| \leq c_i,$$

where  $x_i$  and  $y_i$  are in the domain of the  $i$ -th coordinate.

McDiarmid's inequality

Let  $X_1, \dots, X_n$  be **independent** random variables. Let  $f$  be **Lipschitz** with parameters  $\mathbf{c} = (c_1, \dots, c_n)$ . Let  $X = f(X_1, \dots, X_n)$ . Then for any  $t > 0$ ,

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right).$$

- Notice the similarity with Hoeffding's inequality!

## Method of Bounded Differences

A function  $f$  is called **Lipschitz with parameters**  $\mathbf{c} = (c_1, \dots, c_n)$  if for all  $i = 1, 2, \dots, n$ ,

$$|f(x_1, x_2, \dots, x_{i-1}, \mathbf{x}_i, x_{i+1}, \dots, x_n) - f(x_1, x_2, \dots, x_{i-1}, \mathbf{y}_i, x_{i+1}, \dots, x_n)| \leq c_i,$$

where  $x_i$  and  $y_i$  are in the domain of the  $i$ -th coordinate.

McDiarmid's inequality

Let  $X_1, \dots, X_n$  be **independent** random variables. Let  $f$  be **Lipschitz** with parameters  $\mathbf{c} = (c_1, \dots, c_n)$ . Let  $X = f(X_1, \dots, X_n)$ . Then for any  $t > 0$ ,

$$\mathbf{P}[X \geq \mu + t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right),$$

and

$$\mathbf{P}[X \leq \mu - t] \leq \exp\left(-\frac{2t^2}{\sum_{i=1}^n c_i^2}\right).$$

- Notice the similarity with Hoeffding's inequality!
- The proof is omitted here (it requires the concept of **martingales**).

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

Application 1: Balls into Bins

Application 2: Randomised QuickSort

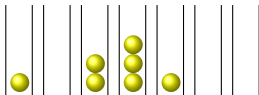
Extensions of Chernoff Bounds

**Applications of Method of Bounded Differences**

Appendix

### Application 3: Balls into Bins (again...)

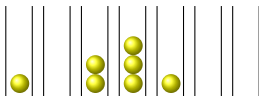
---



- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.

### Application 3: Balls into Bins (again...)

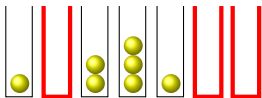
---



- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$

## Application 3: Balls into Bins (again...)

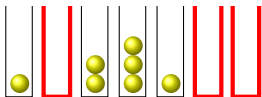
---



- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$
- Let  $Z$  be the number of empty bins (after assigning the  $m$  balls)

## Application 3: Balls into Bins (again...)

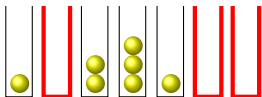
---



- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$
- Let  $Z$  be the number of empty bins (after assigning the  $m$  balls)
- $Z = Z(X_1, \dots, X_m)$  and  $Z$  is Lipschitz with  $\mathbf{c} = (1, \dots, 1)$

## Application 3: Balls into Bins (again...)

---

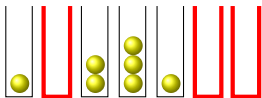


- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$
- Let  $Z$  be the number of empty bins (after assigning the  $m$  balls)
- $Z = Z(X_1, \dots, X_m)$  and  $Z$  is Lipschitz with  $\mathbf{c} = (1, \dots, 1)$   
(If we move one ball to another bin, number of empty bins changes by  $\leq 1$ .)



## Application 3: Balls into Bins (again...)

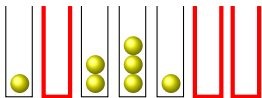
---



- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$
- Let  $Z$  be the number of empty bins (after assigning the  $m$  balls)
- $Z = Z(X_1, \dots, X_m)$  and  $Z$  is Lipschitz with  $\mathbf{c} = (1, \dots, 1)$   
(If we move one ball to another bin, number of empty bins changes by  $\leq 1$ .)
- By McDiarmid's inequality, for any  $t \geq 0$ ,

$$\mathbf{P}[|Z - \mathbf{E}[Z]| > t] \leq 2 \cdot e^{-2t^2/m}.$$

## Application 3: Balls into Bins (again...)

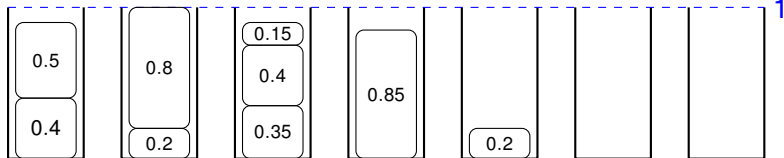


- Consider again  $m$  balls assigned uniformly at random into  $n$  bins.
- Enumerate the balls from 1 to  $m$ . Ball  $i$  is assigned to a random bin  $X_i$
- Let  $Z$  be the number of empty bins (after assigning the  $m$  balls)
- $Z = Z(X_1, \dots, X_m)$  and  $Z$  is Lipschitz with  $\mathbf{c} = (1, \dots, 1)$   
(If we move one ball to another bin, number of empty bins changes by  $\leq 1$ .)
- By McDiarmid's inequality, for any  $t \geq 0$ ,

$$\mathbf{P}[|Z - \mathbf{E}[Z]| > t] \leq 2 \cdot e^{-2t^2/m}.$$

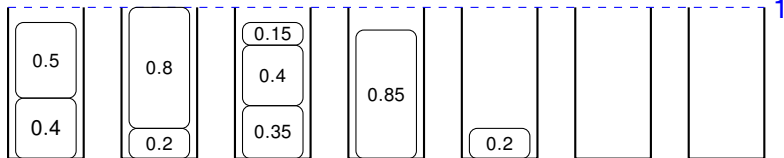
This is a decent bound, but for some values of  $m$  it is far from tight and stronger bounds are possible through a refined analysis.

## Application 4: Bin Packing



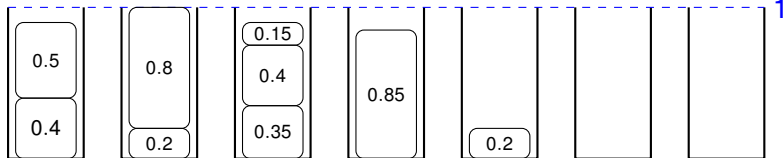
- We are given  $n$  items of sizes in the unit interval  $[0, 1]$

## Application 4: Bin Packing



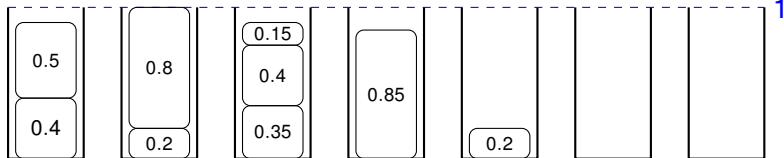
- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**

## Application 4: Bin Packing



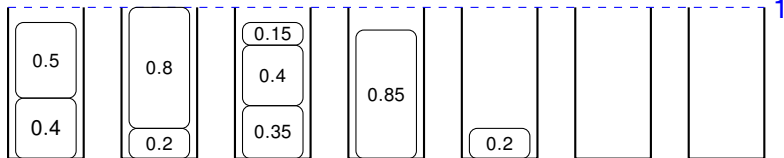
- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**
- Suppose the item sizes  $X_i$  are **independent random variables** in  $[0, 1]$

## Application 4: Bin Packing



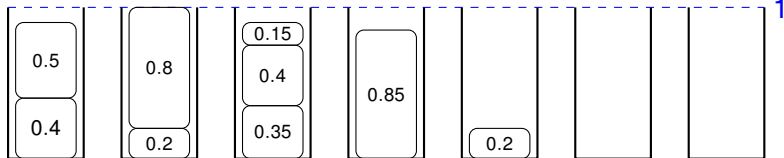
- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**
- Suppose the item sizes  $X_i$  are **independent random variables** in  $[0, 1]$
- Let  $B = B(X_1, \dots, X_n)$  be the **optimal number of bins**

## Application 4: Bin Packing



- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**
- Suppose the item sizes  $X_i$  are **independent random variables** in  $[0, 1]$
- Let  $B = B(X_1, \dots, X_n)$  be the **optimal number of bins**
- The Lipschitz conditions holds with  $\mathbf{c} = (1, \dots, 1)$ . **Why?**

## Application 4: Bin Packing

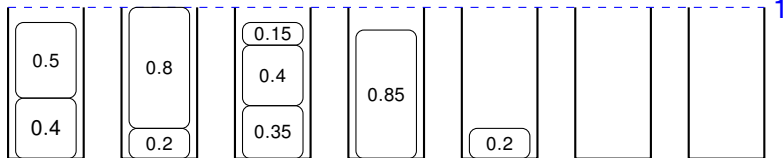


- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**
- Suppose the item sizes  $X_i$  are **independent random variables** in  $[0, 1]$
  
- Let  $B = B(X_1, \dots, X_n)$  be the **optimal number of bins**
- The Lipschitz conditions holds with  $\mathbf{c} = (1, \dots, 1)$ . **Why?**
- Therefore

$$\mathbf{P}[|B - \mathbf{E}[B]| \geq t] \leq 2 \cdot e^{-2t^2/n}.$$



## Application 4: Bin Packing



- We are given  $n$  items of sizes in the unit interval  $[0, 1]$
- We want to pack those items into the **fewest number of unit-capacity bins**
- Suppose the item sizes  $X_i$  are **independent random variables** in  $[0, 1]$
  
- Let  $B = B(X_1, \dots, X_n)$  be the **optimal number of bins**
- The Lipschitz conditions holds with  $\mathbf{c} = (1, \dots, 1)$ . **Why?**
- Therefore

$$\mathbf{P}[|B - \mathbf{E}[B]| \geq t] \leq 2 \cdot e^{-2t^2/n}.$$

This is a typical example where proving concentration is much easier than calculating (or estimating) the expectation!

# Outline

---

Introduction to Chernoff Bounds

How to Derive Chernoff Bounds

Application 1: Balls into Bins

Application 2: Randomised QuickSort

Extensions of Chernoff Bounds

Applications of Method of Bounded Differences

**Appendix**

## Moment Generating Functions

---

— Moment-Generating Function —

The **moment-generating** function of a random variable  $X$  is

$$M_X(t) = \mathbf{E} \left[ e^{tX} \right], \quad \text{where } t \in \mathbb{R}.$$

## Moment Generating Functions

Moment-Generating Function

The **moment-generating** function of a random variable  $X$  is

$$M_X(t) = \mathbf{E} \left[ e^{tX} \right], \quad \text{where } t \in \mathbb{R}.$$

Using power series of  $e$  and differentiating shows that  $M_X(t)$  encapsulates all moments of  $X$ .

## Moment Generating Functions

— Moment-Generating Function —

The **moment-generating** function of a random variable  $X$  is

$$M_X(t) = \mathbf{E} \left[ e^{tX} \right], \quad \text{where } t \in \mathbb{R}.$$

Using power series of  $e$  and differentiating shows that  $M_X(t)$  encapsulates all moments of  $X$ .

— Lemma —

1. If  $X$  and  $Y$  are two r.v.'s with  $M_X(t) = M_Y(t)$  for all  $t \in (-\delta, +\delta)$  for some  $\delta > 0$ , then the distributions  $X$  and  $Y$  are identical.
2. If  $X$  and  $Y$  are **independent** random variables, then

$$M_{X+Y}(t) = M_X(t) \cdot M_Y(t).$$

## Moment Generating Functions

### Moment-Generating Function

The **moment-generating** function of a random variable  $X$  is

$$M_X(t) = \mathbf{E} \left[ e^{tX} \right], \quad \text{where } t \in \mathbb{R}.$$

Using power series of  $e$  and differentiating shows that  $M_X(t)$  encapsulates all moments of  $X$ .

### Lemma

1. If  $X$  and  $Y$  are two r.v.'s with  $M_X(t) = M_Y(t)$  for all  $t \in (-\delta, +\delta)$  for some  $\delta > 0$ , then the distributions  $X$  and  $Y$  are identical.
2. If  $X$  and  $Y$  are **independent** random variables, then

$$M_{X+Y}(t) = M_X(t) \cdot M_Y(t).$$

Proof of 2:

$$M_{X+Y}(t) = \mathbf{E} \left[ e^{t(X+Y)} \right] = \mathbf{E} \left[ e^{tX} \cdot e^{tY} \right] \stackrel{(!)}{=} \mathbf{E} \left[ e^{tX} \right] \cdot \mathbf{E} \left[ e^{tY} \right] = M_X(t)M_Y(t) \quad \square$$