Randomised Algorithms

Lecture 13: Streaming Algorithms

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Introduction

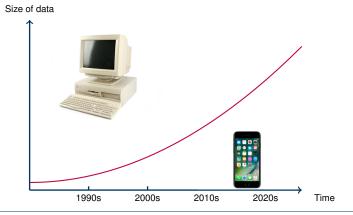
Approximate Counting

Distinct Elements and Frequency Moments

Extra Material (non-examinable): An Algorithm for F₀ in the Turnstile Model

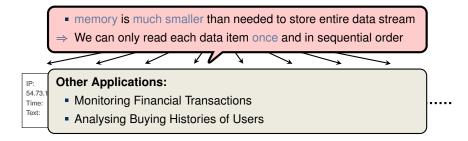
Background of Streaming Algorithms

- The amount of data has been increased exponentially over the last years
- For many applications computational devices' memories are limited
- We need to find good (approximate) solutions without storing the entire input!



Motivation: Analysing Search Engine Queries

- What is the total number queries?
- What is the total number of different IP addresses?
- Extension 1: only consider queries within a certain interval (sliding window)
- Extension 2: also allow the cancellation/removal of a query (turnstile model)
- Extension 3: What if we have different data centers? (distributed streaming)



Streaming algorithms

 The input of a streaming algorithm is given as a data stream, which is a sequence of data

$$S = S_1, S_2, \ldots, S_i, \ldots$$

and every s_i belongs to the universe U.

- Constraints for streaming algorithms: the space complexity should be sublinear in |U| and |S|.
- Quality of the output: The algorithm needs to give a good approximate value with high probability.

– (ε, δ) -approximation —

For confidence parameter δ and approximation parameter ϵ , the algorithm's output Output and the exact answer Exact satisfies

 $\mathbf{P}[\operatorname{Output} \in (1 - \varepsilon, 1 + \varepsilon) \cdot \operatorname{Exact}] \geq 1 - \delta.$

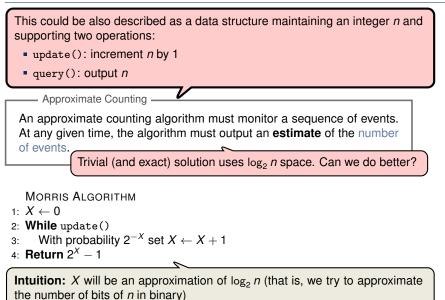
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Approximate Counting and Morris Algorithm



Analysis (1/3)

Lemma (Expectation Analysis)

Let X_n denote the value of X after n updates. For every $n \ge 0$,

$$\mathbf{E}\left[2^{X_n}\right] = n+1.$$
Hence $\Theta_n := 2^{X_n} - 1$ is
an unbiased estimator of n .

Proof:

By Ind

- Base case: For n = 0, we have $X_n = X_0 = 0$
- Induction step: $n \rightarrow n + 1$: By conditioning on X_n ,

$$\mathbf{E}\left[2^{X_{n+1}}\right] = \sum_{j=0}^{\infty} \mathbf{P}\left[X_n = j\right] \cdot \mathbf{E}\left[2^{X_{n+1}} \mid X_n = j\right]$$
$$= \sum_{j=0}^{\infty} \mathbf{P}\left[X_n = j\right] \cdot \left(2^j \cdot \left(1 - \frac{1}{2^j}\right) + 2^{j+1} \cdot \frac{1}{2^j}\right)$$
$$= \sum_{j=0}^{\infty} \mathbf{P}\left[X_n = j\right] \cdot 2^j + \sum_{j=0}^{\infty} \mathbf{P}\left[X_n = j\right]$$
$$= \mathbf{E}\left[2^{X_n}\right] + 1$$
uction Hypothesis
$$\mathbf{Y} = (n+1) + 1.$$

Hence $\Theta_n := 2^{X_n} - 1$ is

Analysis (2/3)

Lemma (Second Moment Analysis) Let X_n denote the value of X after n updates. For every $n \ge 0$, $\mathbf{E}\left[\left(2^{X_n}\right)^2\right] = \mathbf{E}\left[2^{2\cdot X_n}\right] = \frac{3}{2}n^2 + \frac{3}{2}n + 1.$ This is shown similarly to that of the previous Lemma (see supervision sheet)

• Recall
$$\Theta_n = 2^{X_n} - 1$$
.

• Since
$$V[Z] = E[Z^2] - E[Z]^2$$

v

$$\begin{bmatrix} \Theta_n \end{bmatrix} = \mathbf{V} \begin{bmatrix} 2^{X_n} \end{bmatrix} = \mathbf{E} \begin{bmatrix} 2^{2 \cdot X_n} \end{bmatrix} - \left(\mathbf{E} \begin{bmatrix} 2^{X_n} \end{bmatrix} \right)^2$$
$$= \frac{3}{2}n^2 + \frac{3}{2}n + 1 - (n+1)^2 = \frac{n^2 - n}{2}$$

Using Chebysheff's inequality, T

his failure probability (estimate) is at least
$$\frac{1}{2}$$
 ©

$$\mathbf{P}[|\Theta_n - n| \ge \epsilon \cdot n] \le \frac{\mathbf{V}[\Theta_n]}{\epsilon^2 \cdot n^2} \le \frac{\frac{n^2}{2}}{\epsilon^2 \cdot n^2} = \frac{1}{2\epsilon^2}$$

Analysis (3/3)

Idea: Reduce Variance by Running Independent Instances and Taking Average.

IMPROVED MORRIS ALGORITHM(G)

- 1: Let $\Theta^1, \Theta^2, \ldots, \Theta^k$ be k independent instances of MORRIS
- 2: **Return** $\overline{\Theta} := \frac{1}{k} \sum_{i=1}^{k} \Theta^{i}$
- Clearly, $\mathbf{E} \left[\overline{\Theta} \right] = n$. For the variance,

$$\mathbf{V}\left[\overline{\Theta}\right] = \frac{1}{k^2} \cdot \mathbf{V}\left[\sum_{i=1}^k \Theta^i\right] = \frac{1}{k} \cdot \mathbf{V}\left[\Theta^1\right] \le \frac{1}{k} \cdot \frac{n^2}{2}$$

Hence using Chebyshev,

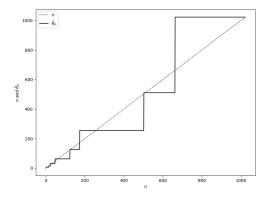
$$\mathbf{P}\left[\left|\overline{\Theta}-n\right|\geq\epsilon\cdot n\right]\leq\frac{1}{2k\epsilon^{2}}.$$

Conclusion

For any $\varepsilon, \delta < 1$, the IMPROVED MORRIS ALG. with $k \geq \frac{1}{2\epsilon^2 \delta}$ satisfies:

$$\mathbf{P}\left[\left|\overline{\Theta}-n\right|\leq\epsilon\cdot n\right]\geq 1-\delta.$$

Simulation



A run of Morris's algorithm on n = 1024 data points

(SOUICE: http://gregorygundersen.com/blog/2019/11/11/morris-algorithm/)

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Extra Material (non-examinable): An Algorithm for F₀ in the Turnstile Model

Norm Estimation: the Alon-Matias-Szegedy algorithm

F_p-norm (Frequency Moments) ——

Let *U* with |U| = n. For $i \in U$, let f_i be the number of occurrences of $i \in U$ in the stream S. Then for any p > 0, the F_{p} -norm is defined by

$$F_{p} := \sum_{i \in U} f_{i}^{p}.$$

- F_1 = total number of items in stream S.
- F_0 = total number of distinct items in stream S.

Alon, Matias, and Szegedy (1996) presented a systematical study for approximating frequency moments.

- F_0, F_1, F_2 can be approximated in space logarithmic in *n* and |S|.
- Approximating F_{ρ} for $\rho \ge 6$ requires $n^{\Omega(1)}$ space.
- The paper won 2005 Gödel Award for "their foundational contribution to streaming algorithms".

We will focus on the simpler case of F_0 , the number of distinct elements.

Pairwise Independence -

A family of functions $H = \{h \mid h : U \mapsto [n]\}$ is pairwise independent if, for any *h* chosen uniformly at random from *H*, the following holds:

1. h(x) is uniformly distributed in $[n] = \{1, 2, ..., n\}$ for any $x \in U$; 2. For any $x_1 \neq x_2 \in U$, $h(x_1)$ and $h(x_2)$ are independent.

Theorem (Fact) -----

Let *n* be a prime number, and let $h_{a,b}(x) = (ax + b) \mod n$. Define

$$H = \{h_{a,b} \mid 0 \le a, b \le n - 1\}.$$

Then *H* is a family of pairwise independent hash functions.

Assume that we have a random hash function h. Define

$$\rho(\mathbf{x}) := \max_{i \ge 0} \left\{ i : \mathbf{x} \mod 2^i = \mathbf{0} \right\},\,$$

which is the number of consecutive 0's among the lowest bits of x.

Example: $\rho(2) = 1$, $\rho(3) = 0$, $\rho(4) = 2$, $\rho(8) = 3$, $\rho(16) = 4$, $\rho(17) = 0$.

Observation. Since h(x) is uniformly distributed over [n], the following holds:

- with probability 1/2, we have $\rho(h(x)) \ge 1$
- with probability 1/4, we have $\rho(h(x)) \ge 2$
- with probability 1/8, we have $\rho(h(x)) \ge 3$
- with probability $1/2^r$, we have $\rho(h(x)) \ge r$

Since *n* is not a power of 2, this probability is in fact equal to $\frac{\lfloor n/2^r \rfloor}{n} \approx 1/2^r - o(1)$.

AMS ALGORITHM

- 1: Choose a random hash function $h : [n] \rightarrow [n]$
- 2: $Z \leftarrow 0$
- 3: while item x from stream S arrives

4: if
$$\rho(h(x)) > Z$$
 then $Z \leftarrow \rho(h(x)) \leq Z \leftarrow \max\{Z, \rho(h(x))\}$
5: return $2^{Z+1/2}$

Analysis of AMS Algorithm

With constant probability > 0, the algorithm's output satisfies

$$2^{Z+1/2} \in [F_0/3, 3 \cdot F_0].$$

We get an $(O(1), \delta)$ -approximation of F_0 by running $\Theta(\log(1/\delta))$ independent copies of the algorithm and returning the median.

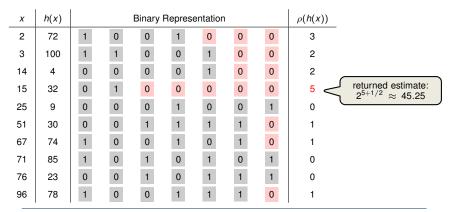
$$\begin{array}{c} & \\ \hline & \\ \hline & \\ \textbf{Recall } (\varepsilon, \delta) \text{-approximation:} \\ & \textbf{P} [\text{Output} \in (1 - \varepsilon, 1 + \varepsilon) \cdot \text{Exact}] \geq 1 - \delta \end{array}$$

Example of the AMS Algorithm

- Assume n = 101 (which is prime)
- The hash function is $h(x) = (ax + b) \mod n$ with a = 28, b = 16
- The data stream is:

S = (25, 76, 14, 51, 25, 14, 76, 76, 3, 51, 96, 14, 67, 3, 15, 25, 2, 76, 14, 71)

• $F_0 = 10$, as the following numbers appeared: {2,3,14,15,25,51,67,71,76,96}



Streaming © Thomas Sauerwald

Analysis (1/2)

Let $X_{r,j}$ be a 0/1 indicator random variable such that

$$X_{r,j} = 1 \Leftrightarrow \rho(h(j)) \geq r.$$

We say item *j* reaches level *r* if $X_{r,j} = 1$.

Let $Y_r = \sum_{j \in S} X_{r,j}$ be the number of items *j* reaching level *r*.

Using that h(j) is uniformly distributed, we conclude

$$\mathbf{E}[X_{r,j}] = \mathbf{P}[\rho(h(j)) \ge r] = \mathbf{P}[h(j) \mod 2^r = 0] = 2^{-r}.$$
definition of function ρ

By linearity of expectation, we have

$$\mathbf{E}[Y_r] = \sum_{j \in \mathcal{S}} \mathbf{E}[X_{r,j}] = \frac{F_0}{2^r}$$

$$\mathbf{V}[Y_r] = \sum_{j \in S} \mathbf{V}[X_{r,j}] \le \sum_{j \in S} \mathbf{E}\left[X_{r,j}^2\right] = \sum_{j \in S} \mathbf{E}[X_{r,j}] = \frac{F_0}{2^r}$$

airwise independence of *h*!

Streaming © Thomas Sauerwald

using pa

Analysis (2/2)

We have proved **E** $[Y_r] = \frac{F_0}{2^r}$ and **V** $[Y_r] \le \frac{F_0}{2^r}$. By Markov's inequality, we have

$$\mathbf{P}[Y_r > 0] = \mathbf{P}[Y_r \ge 1] \le \frac{\mathbf{E}[Y_r]}{1} = \frac{F_0}{2^r}.$$

By Chebyshev's inequality, we have

$$\mathbf{P}[Y_r=0] \leq \mathbf{P}[|Y_r-\mathbf{E}[Y_r]| \geq F_0/2^r] \leq \frac{\mathbf{V}[Y_r]}{(F_0/2^r)^2} \leq \frac{2^r}{F_0}.$$

Let *Z* be the final integer the algo. keeps. So the algo. returns $2^{Z+1/2}$. Let *p* be the smallest integer such that $2^{p+1/2} \ge 3F_0$:

$$\mathbf{P}\left[2^{Z+1/2} \ge 3F_0\right] = \mathbf{P}\left[Z \ge p\right] = \mathbf{P}\left[Y_p > 0\right] \le \frac{F_0}{2^p} \le \frac{\sqrt{2}}{3}$$

Let *q* be the largest integer such that $2^{q+1/2} \le F_0/3$: Union Bound: Error $\le 2 \cdot \frac{\sqrt{2}}{3} < 1$

$$\mathbf{P}\left[2^{Z+1/2} \le F_0/3\right] = \mathbf{P}\left[Z \le q\right] \le \mathbf{P}\left[Y_{q+1} = 0\right] \le \frac{2^{q+1}}{F_0} \le \frac{\sqrt{2}}{3}. \quad \Box$$

- Durand and Flajolet (2003) proposed the LOGLOG algorithm for estimating F_0
- Their algorithm condenses the whole of Shakespeare's works to a table of 256 "small bytes" of 4 bits each
- The estimate of the number of distinct words is $\widetilde{F_0} = 30897$, while the true answer is $F_0 = 28239$, which represents a relative error +9.4%.

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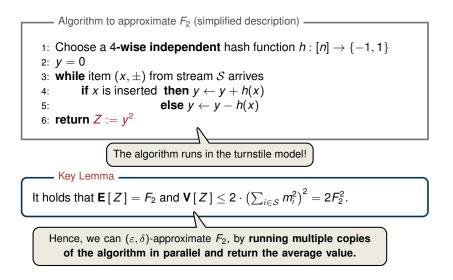
Common approach for designing algorithms in the cash register model:

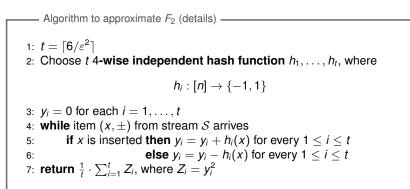
- 1. Sample the data items based on hashed values;
- 2. Store the statistical information of the sampled items, or store the sampled items directly.

Downside of this framework:

- Sampling probability for the current item usually depends on the whole data stream that algorithm has seen so far.
- Deleting an item appeared before could potentially make the current statistical information useless! :(

Sampling techniques are usually non-applicable in the turnstile model.





Analysis

With constant probability, the returned value of the algorithm lies in $(1 - \varepsilon, 1 + \varepsilon) \cdot F_2$. Moreover, the space complexity is $O((1/\varepsilon^2) \log n)$ bits.