## You are encouraged to submit your solutions at student reception or by emailing them to jas289 by 2pm Friday 21th of February

**Question 1.** Let  $X_n$  be the sum of n independent rolls of a fair die. Show that, for any  $k \ge 2$ ,

$$\lim_{n \to \infty} \mathbf{P}[X_n \text{ is divisible by } k] = \frac{1}{k}.$$

**Question 2.** When the U bus arrives outside the Computer Lab, the next bus arrives in 1, 2, ..., 20 minutes with equal probability. You arrive at the bus stop without checking the schedule, at some fixed time n.

- (i) How could you model  $X_n$ , the number of minutes until the next bus when you arrive at time n, as a Markov chain?
- (ii) Buses have been coming and going all day so we can assume the chain has mixed when you arrive. What is the probability of waiting i minutes for a bus in relation to the Chain?
- (iii) How long, on average, do you wait until the next bus arrives?
- (iv) What is the standard deviation of this time?

Solution: Item (i): Let  $X_n$  be the time elapsed from time n till the arrival of the next bus. Then  $X_n$  is a Markov chain on  $\{0, 1, \ldots, 20\}$ , where each state represents how long till the next bus. The transition probabilities are p(0, k) = 1/20 for each  $1 \le k \le 20$  and p(k, k - 1) = 1 for all k > 0.

Item (ii): The idea is that since the chain is stationary when you arrive your just sampling a state according to stationary distribution. Since each state represents the minutes until the next bus, the probability of waiting *i* minutes for a bus is then just the stationary probability  $\pi(i)$  of state *i*.

*Item (iii):* For the stationary distribution observe that

$$\pi(0) = \pi(1)$$
  

$$\pi(1) = \frac{\pi(0)}{20} + \pi(2)$$
  

$$\pi(2) = \frac{\pi(0)}{20} + \pi(3)$$
  

$$\vdots$$
  

$$\pi(19) = \frac{\pi(0)}{20} + \pi(20)$$
  

$$\pi(20) = \frac{\pi(0)}{20}.$$

Thus we have  $\pi(1) = \pi(1)/20 + \pi(2)$  thus since  $\pi(0) = \pi(1)$ ,

$$\pi(2) = \frac{19\pi(0)}{20}.$$

However also we have  $\pi(2) = \pi(0)/20 + \pi(3)$  thus  $\pi(3) = 18\pi(0)/19$  and in general for any  $1 \le k \le 20$ ,

$$\pi(k) = \frac{21 - k}{20}\pi(0).$$

To find the values of  $\pi$  all we must do is find  $\pi(0)$ , we do this by the fact  $\pi$  is a probability distribution,

$$1 = \sum_{k=0}^{20} \pi(k) = \pi(0) \left( 1 + \sum_{k=1}^{20} \frac{21-k}{20} \right),$$

which implies that  $\pi(0) = 2/23$ . Hence we have

$$\mathbb{E}_{\pi}[X_n] = \sum_{k=0}^{20} k\pi(k) = \sum_{k=1}^{20} k \frac{2}{23} \frac{21-k}{20} = \frac{154}{23} \approx 6.7 \text{ mins.}$$

*Item (iv):* The second moment is given by

$$\mathbb{E}_{\pi}[X_n^2] = \sum_{k=0}^{20} k^2 \pi(k) = \sum_{k=1}^{20} k^2 \frac{2}{23} \frac{21-k}{20} = \frac{1617}{23}$$

Thus the s.d. is

$$\sqrt{\frac{1617}{23} - \left(\frac{154}{23}\right)^2} = \frac{35\sqrt{11}}{23} \approx 5 \text{ mins}$$

**Question 3.** Prove the following Lemma from class: For any probability distributions  $\mu$  and  $\eta$  on a countable state space  $\Omega$ 

$$\left\|\mu - \eta\right\|_{tv} = \frac{1}{2} \sum_{\omega \in \Omega} \left|\mu(\omega) - \eta(\omega)\right|$$

 $Solution: \ \ {\rm Let} \ \Omega^+ = \{\omega: \mu(\omega) \geq \eta(\omega)\} \ {\rm and} \ \Omega^- = \{\omega: \mu(\omega) < \eta(\omega)\} \ . \ {\rm Then}$ 

$$\max_{A \subseteq \Omega} \mu(A) - \eta(A) = \mu(\Omega^+) - \eta(\Omega^+)$$

and

$$\max_{A \subseteq \Omega} \eta(A) - \mu(A) = \eta(\Omega^{-}) - \mu(\Omega^{-}).$$

Since  $\Omega = \Omega^+ \cup \Omega^-$  and  $\Omega^+ \cap \Omega^- = \emptyset$  we have

$$\mu(\Omega^+) + \mu(\Omega^-) = 1 \qquad \text{and} \qquad \eta(\Omega^+) + \eta(\Omega^-) = 1,$$

thus

$$\mu(\Omega^+) - \eta(\Omega^+) = \eta(\Omega^-) - \mu(\Omega^-).$$

Hence

$$\sup_{a \in \Omega} |\mu(A) - \eta(A)| = |\mu(\Omega^+) - \eta(\Omega^+)| = |\mu(\Omega^-) - \eta(\Omega^-)|.$$

Combining the above yields

$$2 \|\mu - \eta\|_{tv} = |\mu(\Omega^{+}) - \eta(\Omega^{+})| + |\mu(\Omega^{-}) - \eta(\Omega^{-})| = \sum_{\omega \in \Omega} |\mu(\omega) - \eta(\omega)|.$$

**Question 4.** This question asks you to prove lower bounds on the mixing time of some lazy random walks on graphs.

- 1. Let  $G = (V_1 \cup V_2, E)$  be a graph made of two disjoint complete graphs of n vertices, supported respectively on  $V_1$  and  $V_2$ , connected by a single edge. This is called the Barbell graph. Consider a lazy random walk on G. Prove that  $t_{mix}(G) = \Omega(n^2)$  (recall from Lecture 8 that  $t_{mix} = \tau(1/4)$ ).
- 2. Suppose now we add s < n edges to the Barbell graph, where each edge has one endpoint in  $V_1$  and the other endpoint in  $V_2$ . What happens to  $t_{mix}(G)$ ?

3. Consider now a version of the Barbell graph where  $|V_1| = n, |V_2| = \lfloor \log(n) \rfloor$  and there exists only an edge between  $V_1$  and  $V_2$ . What is the mixing time of this graph?

Solution: For part (i): let  $\pi$  be the stationary distribution of a lazy random walk in G (recall that, for any vertex u,  $\pi(u) = d(u)/2|E|$  where d(u) is the degree of u). Now notice that, by symmetry,  $\sum_{u \in V_1} \pi(u) = \sum_{u \in V_2} \pi(u) = 1/2$ . You can prove this explicitly by using the formula for the stationary distribution mentioned above. Consider a probability distribution p such that  $\sum_{u \in V_2} p(u) \leq \epsilon$  for some small  $\epsilon \geq 0$ . Then,

$$\begin{split} \|p - \pi\|_{TV} &= \frac{1}{2} \sum_{u \in V_1} |p(u) - \pi(u)| + \frac{1}{2} \sum_{u \in V_2} |p(u) - \pi(u)| \\ &\geq \frac{1}{2} \sum_{u \in V_1} (p(u) - \pi(u)) + \frac{1}{2} \sum_{u \in V_2} (\pi(u) - p(u)) \\ &= \frac{1}{2} \left( \sum_{u \in V_1} p(u) - \sum_{u \in V_1} \pi(u) + \sum_{u \in V_2} \pi(u) - \sum_{u \in V_2} p(u) \right) \\ &\geq \frac{1}{2} \left( 1 - \epsilon - \frac{1}{2} \right) + \frac{1}{2} \left( \frac{1}{2} - \epsilon \right) = \frac{1}{2} - \epsilon \end{split}$$

where the last inequality follows from the facts that  $\sum_{u \in V_2} p(u) \leq \epsilon$  and  $\sum_{u \in V_1} \pi(u) = \sum_{u \in V_2} \pi(u) = 1/2$ . Therefore, a walk to be mixed must have at least probability  $\epsilon \geq 1/4$  to be in  $V_2$ .

But now suppose a walk start from a vertex  $u \in V_1$  which is not the only vertex  $v \in V_1$  adjacent to a vertex in  $V_2$ . Then, at each step, if the walk it's still in  $V_1$ , it has probability  $O(1/n^2)$  to move to  $V_2$ (because it must move first to v and then move in  $V_2$ ). Therefore, after t steps,  $\sum_{w \in V_2} P^t(u, w) = O(t/n^2)$ (this follows from a union bounds on the events "at step i the walk moves from  $V_1$  to  $V_2$ " for  $i = 1, \ldots, t$ ). Hence, we need to wait  $\Omega(n^2)$  before the walk is close to stationarity.

For part (*ii*) repeat the same argument as in part (*i*) but now at each step the probability to go from  $V_1$  to  $V_2$  is  $\Omega(s/n^2)$ . Therefore,  $t_{mix} = O(n^2/s)$  (when you reach  $V_2$ , since the subgraph supported on  $V_2$  is complete, after a few steps you are mixed).

For part (*ii*), repeating again the same argument it is clear that to be mixed we just need to move from  $V_2$  to  $V_1$  (it is important here to notice that the worst case is to start in  $V_2$ : since  $V_2$  is very small compared to  $V_1$ , if we start in the latter our argument doesn't work anymore). But this happens with probability  $\Theta(1/(\log n)^2)$ . Therefore mixing happens in  $O(\log n)^2$  steps.

**Question 5.** Let  $\langle \cdot, \cdot \rangle_{\pi}$  be the inner product defined in the lecture. Show that it satisfies the following properties:

Symmetry For any  $f, g \in \ell_2(\pi)$ ,  $\langle f, g \rangle_{\pi} = \langle g, f \rangle_{\pi}$ .

Linearity For any  $f, g, h \in \ell_2(\pi)$  and  $\alpha, \beta \in \mathbb{R}$ ,  $\langle \alpha f + \beta g, h \rangle_{\pi} = \alpha \langle f, h \rangle_{\pi} + \beta \langle g, h \rangle_{\pi}$ .

Positive definiteness For any  $\underline{0} \neq f \in \ell_2(\pi), \langle f, f \rangle_{\pi} > 0.$ 

Do all these properties hold if  $\pi$  is not always positive?

**Question 6.** Given a matrix M such that  $Mf = \lambda f$  (i.e., f is an eigenvector with eigenvalue  $\lambda$  of M), prove that  $M^k f = \lambda^k f$ 

## Hints

**Hint** (Question 1). At face value  $X_n$  is an (infinite) Markov chain on  $\mathbb{N}$ . We would like to consider it as a finite Markov chain, reduction mod m (for some suitable m) will help us achieve this.

**Hint** (Question 4(i)). First prove that, for a distribution p such that  $\sum_{u \in V_2} p(u) \leq \epsilon$  for some small  $\epsilon \geq 0$ ,  $\|p - \pi\|_{TV} \geq \frac{1}{2} - \epsilon$ , where  $\pi$  is the stationary distribution of a random walk on the Barbell graph. Use this fact to obtain a lower bound on the mixing time (think how many steps we need so that  $\sum_{u \in V_2} p^t(u) > \epsilon$ , where  $p^t$  is the random walk distribution after t steps).