

# Probability and Computation: Problem sheet 4 Solutions

You are encouraged to submit your solutions by emailing them to [luca.zanetti@cl.cam.ac.uk](mailto:luca.zanetti@cl.cam.ac.uk) by Wednesday 20th of February.

**Question 1.** Let  $(X_i)_{i \geq 1}$  be independent random variables with  $\mathbf{P}[X_i = 1] = \mathbf{P}[X_i = -1] = 1/2$ . Let  $S_n = \sum_{i=0}^n X_i$ , with  $X_0 = K > 0$ . For  $N > K$  define

$$T = T_{0,N} = \min\{n \geq 0 : S_n = 0 \text{ or } S_n = N\}.$$

1. Prove that  $\mathbf{E}[T] < \infty$  (you cannot use the OST).
2. Find a deterministic sequence of values  $a_n \in \mathbb{R}$  such that  $Z_n = S_n^3 + a_n S_n$  is a martingale w.r.t.  $X_0, X_1, \dots$
3. Find deterministic sequences  $b_n, c_n \in \mathbb{R}$  such that  $W_n = S_n^4 + b_n S_n^2 + c_n$  is a martingale w.r.t.  $X_0, X_1, \dots$

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*Solution:* 1. Let  $T'$  be the first time we win  $N$  times in a row. When  $T'$  happens we either have that  $S_{T'} \geq N$  or not. If  $S_{T'} \geq N$  it means that at some time  $t < T'$  we had  $S_t = N$ . If  $S_{T'} < N$  then  $S_{T'-N} < 0$  then at some time  $t < T' - N$  we had  $S_t = 0$ . In both cases it exist some  $t < T'$  such that  $S_t = 0$  or  $S_t = N$ , therefore  $T < T'$ . The expectation of  $T'$  was computed in Question 6 of the previous problem sheet.

2. Clearly  $Z_n$  is a function of  $X_0, \dots, X_n$  and  $|Z_n| \leq n^3 + |a_n|n$ , which check the first two parts of the definition of a martingale. We need to find  $a_n$  such that

$$\mathbf{E}[Z_{n+1}|X_0, \dots, X_n] = Z_n.$$

note that

$$S_{n+1}^3 = (S_n + X_{n+1})^3 + a_n(S_n + X_{n+1}) = S_n^3 + 3S_n^2 X_{n+1} + 3S_n X_{n+1}^2 + X_{n+1}^3$$

Note that  $X_{n+1}^2 = 1$  and  $X_{n+1}^3 = X_{n+1}$ , then

$$\mathbf{E}[Z_{n+1}|X_0, \dots, X_n] = S_n^3 + 3S_n + a_{n+1}S_n \tag{1}$$

we want to force  $Z_n = S_n^3 + 3S_n + a_{n+1}S_n$ , which implies that  $a_n = 3 + a_{n+1}$ . The solutions of the recursion for  $a_n$  is  $a_n = -3n + a_0$ . We just set the initial value  $a_0 = 0$ . Therefore  $Z_n = S_n^3 - 3nS_n$  is a martingale

3.  $W_n$  is a function of  $X_0, \dots, X_n$  and  $|W_n| \leq n^4 + |b_n|n^2 + |c_n|$ , which check the first two parts of the definition of a martingale.

We proceed to check that  $\mathbf{E}[W_{n+1}|X_0, \dots, X_n] = W_n$  for some values of  $b_n$  and  $c_n$ .

Using that  $X_n^{2k} = 1$  and  $X_n^{2k+1} = X_n$  for  $k \geq 0$ , then

$$S_{n+1}^4 = (S_n + X_{n+1})^4 = S_n^4 + 4S_n^3 X_{n+1} + 6S_n^2 + 4S_n X_{n+1} + 1,$$

and

$$S_{n+1}^2 = S_n^2 + 2S_n X_{n+1} + 1.$$

Hence

$$\mathbf{E}[W_{n+1}|X_0, \dots, X_n] = S_n^4 + (6 + b_{n+1})S_n^2 + (1 + b_{n+1} + c_{n+1}) = S_n^4 + b_n S_n^2 + c_n$$

This suggest that  $b_n = 6 + b_{n+1}$  and  $c_n = 1 + c_{n+1} + b_{n+1}$ . Clearly  $b_n = -6n + b_0$ , so we choose  $b_0 = 0$ . Then  $c_n = 1 - 6n + c_{n+1}$ , hence

$$c_n - c_0 = \sum_{i=1}^n c_i - c_{i-1} = \sum_{i=1}^n [6(i-1) - 1] = 3n^2 - 4n$$

Again, we choose  $c_0 = 0$  and we conclude that

$$W_n = S_n^4 - 6nS_n^2 + (3n^2 - 4n)$$

is a martingale

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### Question 2.

1. Consider out path on vertices  $\{0, \dots, N\}$ , and suppose  $X_0 = K$ . Compute  $h_{K,N}$
2. Compute the cover time of a path on  $\{0, \dots, N\}$  when  $N$  is even. What about when  $N$  is odd?
3. Consider a cycle on vertices  $\{0, 1, \dots, N\}$  where vertex  $i$  is adjacent to  $i + 1$ , and  $0$  is adjacent to  $N$ . Compute the cover time.
4. Consider a cycle on vertices  $\{0, 1, \dots, N\}$ . Define  $T$  by

$$T = \min\{m : \cup_{i=0}^m X_i = \{0, \dots, N\}\}$$

$T$  is the first time all vertices has been covered. Compute  $\mathbf{P}[X_T = i | X_0 = 0]$  for  $i \in \{1, \dots, N\}$ .

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*Solution:* 1. A random walk on the path can be seen as a folded version of a random walk on  $\mathbb{Z}$  where we associate point  $-k$  with point  $k$ . Let  $S_n = \sum_{i=1}^n X_i$  be a random walk on  $\mathbb{Z}$  starting from  $K$  (recall that  $X_i$  are independent,  $\mathbf{P}[X_i = 1] = \mathbf{P}[X_i = -1] = 1/2$ , and  $X_0 = 0$ .)

Therefore, the problem of computing  $h_{K,N}$  is equivalent to the problem to finding  $\mathbf{E}[T_{N-,N} | S_0 = K]$ . It was shown in class that this is equivalent to  $\mathbf{E}[T_{0,2N} | S_n = N + K] = (N + K)(2N - (N + K)) = (N + K)(N - K)$

2. Starting from  $K$  we need to hit either  $0$  or  $N$  which takes  $K(N - K)$  times in expectation. Later from one of the extremes we need to hit the other extreme which takes  $N^2$  times. Then, the total time to cover the graph is  $N^2 + K(N - K)$ . Recall the cover time is considered over the worst starting point, so maximizing  $K$  we get that the cover time is  $N^2 + \lfloor N/2 \rfloor \lceil N/2 \rceil$

3. First of all, note that the set of covered vertices is a path in the cycle. Second, note when we just discover a new vertex this is one of the extreme points of the path. Third, we can always relabel the vertices of the cycle, so the set of cover vertices is  $\{0, 1, \dots, K\}$  and the last discovered vertex is  $K$ . The expected time to discover a new vertex is equivalent the expected time to move outside this path. As a random walk problem this is equivalent to  $\mathbf{E}[T_{-1,K+1} | S_0 = K] = K + 1$ . Now we just need to sum from  $K = 0$  to  $n - 1$  which equals  $n(n - 1)/2$ .

4. We will prove that  $\mathbf{P}[X_T = i | X_0 = 0] = 1/N$ . For that, note that before hitting vertex  $i$  we are either in vertex  $i - 1$  or vertex  $i + 1 \pmod{N}$ . Without lost of generality, suppose we hit vertex  $i - 1$  before vertex  $i + 1$  (otherwise just rename all the vertices). Then if  $i$  is the last vertex we hit, we have to hit  $i + 1$  before vertex  $i$  starting from vertex  $i - 1$ . Note this is equivalent to start a path in  $1$ , and we want to hit  $N - 1$  before vertex  $0$ . In lectures we analysed that problem and deduced that such a probability is  $1/(N - 1)$ . Therefore  $\mathbf{P}[X_T = i | X_0 = 0] \geq 1/(N - 1)$ . Since there are  $N - 1$  other vertices besides  $0$  and for all of them  $\mathbf{P}[X_T = i | X_0 = 0] \geq 1/(N - 1)$ , then the only possibility is that  $\mathbf{P}[X_T = i | X_0 = 0] = 1/(N - 1)$

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**Question 3. Wald's Equation:** Let  $X_1, \dots$ , i.i.d. non-negative random variables with finite expectation. Let  $T$  be a stopping time with respect to this sequence and suppose that  $\mathbf{E}[T] < \infty$  and that  $\mathbf{E}[|X_1|] < \infty$ . Prove that

$$\mathbf{E}\left[\sum_{i=1}^T X_i\right] = \mathbf{E}[T] \mathbf{E}[X_1].$$

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*Solution:* Define  $Z_n = \sum_{i=1}^n X_i - n\mathbf{E}[X_1]$  and  $Z_0 = 0$ . Then  $Z_n$  is a martingale w.r.t  $X_1, \dots, X_n$ . Clearly,  $Z_n$  is a function of  $X_1, \dots, X_n$ . Also

$$\mathbf{E}[|Z_n|] \leq 2n\mathbf{E}[|X_1|] < \infty.$$

Finally,

$$\mathbf{E}[Z_{n+1}|X_1, \dots, X_n] = Z_n + \mathbf{E}[X_{n+1} - \mathbf{E}[X_1]|X_1, \dots, X_n] = Z_n,$$

where in the last equality we use that the  $X_i$ 's are independent and they have the same distribution and thus  $\mathbf{E}[X_1] = \mathbf{E}[X_{n+1}]$ .

Finally, as  $T$  is a stopping time, we can use condition iii) of the OST and then

$$\mathbf{E}[Z_T] = \mathbf{E}[Z_0] = 0,$$

but

$$Z_T = \sum_{i=1}^T X_i - T\mathbf{E}[X_1],$$

from which we conclude the result.

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**Question 4.** A weighted undirected graph  $G = (V, E, w)$  is defined by a set vertices  $V$ , a collection of edges  $E \subseteq V \times V$ , and a weight function  $w: V \times V \rightarrow \mathbb{R}_{\geq 0}$  such that, for any  $u, v \in V$ ,  $w(u, v) = w(v, u)$  and  $w(u, v) > 0$  if and only if  $(u, v) \in E$ . Self-loops of the kind  $(u, u)$  are allowed. A random walk on  $G = (V, E, w)$  is a Markov chain with transition matrix  $P$  such that, for any  $u, v \in V$ ,  $P(u, v) = w(u, v)/d(u)$ , where  $d(u) = \sum_{z \in V} w(u, z)$ .

1. What is the stationary distribution of this Markov chain?
  2. What does being aperiodic amounts to?
  3. Prove that a Markov chain is reversible if and only if it can be represented by a random walk on a weighted undirected graph.
  4. Prove that if  $P$  is reversible, then  $P^t$  is also reversible for any  $t \in \mathbb{N}$ .
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*Solution:*

1. Let  $\pi: V \rightarrow \mathbb{R}$  such that  $\pi(u) = \frac{d(u)}{\sum_{z \in V} d(z)}$ . Then,  $\pi$  is stationary for  $P$ . To prove this, we just need to check that  $\pi P = \pi$ . Let  $u$  be an arbitrary vertex. Then,

$$(\pi P)(u) = \sum_{v \in V} \pi(v)P(v, u) = \sum_{v \in V} \frac{d(v)P(v, u)}{\sum_{z \in V} d(z)} = \sum_{v \in V} \frac{w(v, u)}{\sum_{z \in V} d(z)} = \frac{d(u)}{\sum_{z \in V} d(z)} = \pi(u).$$

where the third equality follows from  $P(v, u) = w(v, u)/d(v)$ .

2. Being aperiodic for a random walk on an undirected graphs simply means the graph does not contain bipartite connected components.
3. We start showing that a random walk on an undirected graph is always reversible, i.e., it satisfies the detailed balance condition:  $\pi(u)P(u, v) = \pi(v)P(v, u)$  for any  $u, v \in V$ . Let  $u, v$  be arbitrary vertices. Then,

$$\pi(u)P(u, v) = \frac{d(u)}{\sum_{z \in V} d(z)} \cdot \frac{w(u, v)}{d(u)} = \frac{d(v)}{\sum_{z \in V} d(z)} \cdot \frac{w(v, u)}{d(v)} = \pi(v)P(v, u)$$

where the second equation follows from the fact that, since  $G$  is undirected,  $w(u, v) = w(v, u)$ .

We now show the reverse implication: we are given a transition matrix  $P$  on  $\Omega$  with stationary distribution  $\pi$  such that  $\pi(u)P(u, v) = \pi(v)P(v, u)$  for any  $u, v \in \Omega$ , and we want to show we can construct an undirected weighted graph  $G = (V, E, w)$  such that  $P$  is the transition matrix

of a random walk on  $G$ . First of all, we choose  $\Omega = V$ . Then, we construct the weight function  $w: V \rightarrow \mathbb{R}_{\geq 0}$  as  $w(u, v) = \pi(u)P(u, v)$ , and we set  $E = \{\{u, v\}: w(u, v) > 0\}$ . We need to show that  $w$  is a proper weight function. Clearly,  $w$  is nonnegative and strictly positive exactly on  $E$ . Moreover,  $w(u, v) = \pi(u)P(u, v) = \pi(v)P(v, u) = w(v, u)$  since  $P$  is reversible. Finally,

$$\frac{w(u)}{\sum_{z \in V} w(u, z)} = \frac{\pi(u)P(u, v)}{\sum_{z \in V} \pi(u)P(u, z)} = \frac{P(u, v)}{\sum_{z \in V} P(u, z)} = P(u, v),$$

where the last equality follows from the fact that each row of  $P$  sum up to 1 (by definition of transition matrix). Hence, we have shown that a random walk on  $G$  has transition matrix  $P$ .

4. There are several ways to show this fact. The easiest way is probably to use the fact that  $P$  is reversible if and only if  $\langle Pf, g \rangle_\pi = \langle f, Pg \rangle_\pi$  for any  $f, g \in V \rightarrow \mathbb{R}$ . Hence, we need to show that, for arbitrary  $f, g \in V \rightarrow \mathbb{R}$ ,  $\langle P^t f, g \rangle_\pi = \langle f, P^t g \rangle_\pi$ :

$$\langle P^t f, g \rangle_\pi = \langle P^{t-1} f, Pg \rangle_\pi = \langle P^{t-2} f, P^2 g \rangle_\pi = \dots = \langle f, P^t g \rangle_\pi$$

where at each step we have applied the reversibility of  $P$ .

**Question 5.** Let  $P$  be the transition matrix of a (simple) random walk on an undirected graph  $G = (V, E)$ . Let  $\lambda_1 \geq \dots \geq \lambda_n$ . Prove the following.

1.  $\lambda_1 = 1$ .
2.  $\lambda_2 = 1$  if and only if the graph is disconnected.
3.  $\lambda_n = -1$  if and only if there exists a bipartite connected component.
4. Suppose now that the random walk is lazy (i.e.,  $P(u, u) \geq 1/2$  for any  $u \in V$ ). Prove that all the eigenvalues of  $P$  are non-negative.

*Solution:*

1. Since the all-one vector is an eigenvector of  $P$  with eigenvalue 1,  $\lambda_1 \geq 1$ . We just have to show that we cannot have an eigenvalue strictly greater than 1. Actually, we will show something stronger:  $P$  doesn't have eigenvalues strictly greater than 1 *in absolute value*. Let  $f$  be an eigenvector of eigenvalue  $\lambda$  for  $P$ , i.e.,  $Pf = \lambda f$ . Take  $x$  maximising  $|f(x)|$ . Since  $f$  is an eigenvector of eigenvalue  $\lambda$ ,

$$|\lambda| |f(x)| = |(Pf)(x)| = \left| \sum_y P(x, y) f(y) \right| \leq |f(x)| \left| \sum_y P(x, y) \right| = |f(x)|.$$

Clearly,  $f(x) \neq 0$ . Therefore,  $|\lambda| \leq 1$ .

2. We will first show that if  $G$  is disconnected, then  $\lambda_2 = 1$ . Recall the variational characterisation of  $\lambda_2$ :

$$1 - \lambda_2 = \min_{0 \neq f \perp 1} \frac{\sum_{x, y} (f(x) - f(y))^2 P(x, y) \pi(x)}{\|f\|_{2, \pi}^2} \quad (2)$$

This expression is clearly nonnegative. Hence, we just need to find a nonzero  $f \perp 1$  such that the numerator is 0. Since  $G$  is disconnected, we can partition  $V$  in two sets,  $S$  and  $V \setminus S$ , such that there are no edges between the two. Then, as long as  $f$  is constant on  $S$  and  $V \setminus S$ , the numerator is 0 ( $P(x, y)$  is nonnegative only if there exists an edge between  $x$  and  $y$ ). We just need to be careful to construct  $f$  so that  $\langle f, 1 \rangle_\pi = 0$ . The following is a good choice (check!):

$$f(u) = \begin{cases} 1/\text{vol}(S) & \text{if } u \in S \\ -1/\text{vol}(V \setminus S) & \text{if } u \notin S \end{cases}$$

where  $\text{vol}(S) = \sum_{u \in S} d(u) = 2|E| \sum_{u \in S} \pi(u)$ .

We now show that, if  $G$  is connected,  $\lambda_2 < 1$ . Suppose by contradiction  $0 \neq f \perp 1$  is an eigenvector of eigenvalue  $\lambda_2 = 1$ . Again, by (2),

$$\sum_{x,y} (f(x) - f(y))^2 P(x,y) \pi(x) = \sum_{\{x,y\} \in E} (f(x) - f(y))^2 P(x,y) \pi(x) = 0.$$

Therefore, for any  $\{x,y\} \in E$ ,  $f(x) = f(y)$ . But notice that  $0 = \langle f, 1 \rangle_\pi = \sum_z f(z) \pi(z)$  implies that  $f$  must have strictly positive and strictly negative entries. Take  $u, v \in V$  such that  $f(u) > 0$  and  $f(v) < 0$ . Since  $G$  is connected, there exists a path  $u = x_0, x_1, \dots, x_\ell = v$  from  $u$  to  $v$ . But then, since for any  $\{x,y\} \in E$   $f(x) = f(y)$ , we have that  $0 > f(u) = f(x_0) = f(x_1) = \dots = f(x_\ell) = f(v) < 0$ , reaching a contradiction.

3. Without loss of generality, we assume  $G$  is connected and prove that  $P$  has eigenvalue  $-1$  if and only if  $G$  is bipartite. This is without loss of generality because if  $G$  were disconnected, the eigenvalues of  $P$  would simply be the union of the eigenvalues of the transition matrices of the simple random walks in each one of the connected component of  $G$ . First we assume that  $G$  is bipartite and prove it has eigenvalue  $-1$ . Recall that a graph is bipartite if it can be partitioned in two sets  $S$  and  $V \setminus S$  such that no edge connect two vertices in  $S$  or two vertices in  $V \setminus S$  (i.e., it has a cut containing all the edges in the graph). Assuming such  $S$  and  $V \setminus S$  exist, we construct a function  $f: V \rightarrow \mathbb{R}$  as follows:

$$f(u) = \begin{cases} 1 & \text{if } u \in S \\ -1 & \text{if } u \notin S \end{cases}$$

We claim this function is an eigenvector of eigenvalue  $-1$  for  $P$ . Let  $u \in S$ . then,

$$(Pf)(u) = \sum_{v \in V} P(u,v) f(v) = \sum_{v \in S} P(u,v) - \sum_{v \notin S} P(u,v) = - \sum_{v: \{u,v\} \in E} \frac{1}{d(u)} = -1 = -f(u),$$

where the third equality follows from the fact that all the neighbours of  $u$  are in  $V \setminus S$  and the fourth follows from  $u$  having  $d(u)$  neighbours. Analogously, we can prove that  $(Pf)(u) = -f(u)$  for any  $u \in V \setminus S$ . This proves that  $Pf = -f$ , which completes one direction of the proof.

For the reverse direction, we assume that  $P$  has eigenvalue  $-1$  and prove that  $G$  must be bipartite. Let  $f$  (nonzero) such that  $Pf = -f$ . We define  $S = \{u \in V: f(u) \geq 0\}$  and we claim that  $(S, V \setminus S)$  is a bipartition of the graph such that no edge connects two vertices in  $S$  or in  $V \setminus S$ . To this end, pick any  $u \in V$  such that  $|f(u)| = \max_{v \in V} |f(v)|$ . Then,

$$-f(u) = (Pf)(u) = \sum_{v \in V} P(u,v) f(v) = \sum_{v: \{u,v\} \in E} \frac{f(v)}{d(u)}. \quad (3)$$

But since  $u$  has only  $d(u)$  neighbours and  $|f(u)|$  achieves the maximum of  $f$  in absolute value,  $f(v) = -f(u)$  for any  $v$  such that  $\{u,v\} \in E$ . Therefore, since  $G$  is connected, we have proved that  $|f(u)| = |f(v)| \neq 0$  for any  $v \in V$ . Hence, (3) must hold for any  $u \in V$ . This implies that, for any  $\{u,v\} \in E$ ,  $f(u) = -f(v)$ . Therefore,  $u \in S$  and  $v \notin S$ , or vice versa. In other words, we have proved we can partition  $V$  in two sets such that no edge connects two vertices in the same set, i.e.,  $G$  is bipartite.

4. Since  $P$  is lazy,  $P = \frac{1}{2}(I + P')$  where  $I$  is the identity matrix and  $P'$  is the transition matrix of a lazy random walk on  $G$ . We proved above that the eigenvalues of  $P'$  are between 1 and  $-1$ . Hence,  $\lambda_n(P) = \frac{1}{2} + \frac{1}{2}\lambda_n(P') \geq 0$ .