

# Introduction to Probability

Lecture 7: Independence, Covariance and Correlation

Mateja Jamnik, Thomas Sauerwald

University of Cambridge, Department of Computer Science and Technology

email: {mateja.jamnik,thomas.sauerwald}@cl.cam.ac.uk



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## Independence of Random Variables

This definition covers the **discrete** and **continuous** case!

Definition of Independence

Two random variables  $X$  and  $Y$  are **independent** if for all values  $a, b$ :

$$\mathbf{P}[X \leq a, Y \leq b] = \mathbf{P}[X \leq a] \cdot \mathbf{P}[Y \leq b].$$

For two **discrete** random variables, an equivalent definition is:

$$\mathbf{P}[X = a, Y = b] = \mathbf{P}[X = a] \cdot \mathbf{P}[Y = b].$$

This is useless for continuous random variables.

Remark

Using the **joint probability distribution**, the above is equivalent to for all  $a, b$ ,

$$F(a, b) = F_X(a) \cdot F_Y(b).$$

All these definitions extend in the natural way to **more than two** variables!



## Factorisation

### Factorisation

The definition of independence of  $X$  and  $Y$  implies the following **factorisation** formula: for any “suitable” sets  $A$  and  $B$ ,

$$\mathbf{P}[X \in A, Y \in B] = \mathbf{P}[X \in A] \cdot \mathbf{P}[Y \in B]$$

For **continuous** distributions one obtains by differentiating both sides in the formula for the joint distribution:

$$f_{X,Y}(x, y) = f_X(x) \cdot f_Y(y)$$

### Example

Let  $X$  and  $Y$  be two independent variables. Let  $I = (a, b]$  be any interval and define  $U := \mathbf{1}_{X \in I}$  and  $V := \mathbf{1}_{Y \in I}$ . Prove  $U$  and  $V$  are independent.

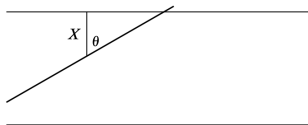
Answer



## Buffon's Needle Problem (1/2)



Georges-Louis Leclerc de Buffon 1707–1788 (Source Wikipedia)



Source: Ross, Probability 8th ed.

- A table is ruled with equidistant, parallel lines a distance  $D$  apart.
- A needle of length  $L$  is thrown randomly on the table.
- **What is the probability that the needle will intersect one of the two lines?**

Let  $X$  be the distance of the **middle point** of the needle to the closest parallel line. Needle intersects a line if hypotenuse of the triangle is less than  $L/2$ , i.e.,

$$\frac{X}{\cos(\theta)} < \frac{L}{2} \quad \Leftrightarrow \quad X < \frac{L}{2} \cos(\theta).$$

We assume that  $X \in [0, D/2]$  and  $\theta \in [0, \pi/2]$  are **independent** and **uniform**.

Can be thought of as: 1. Sample the middle point of needle, 2. Sample the angle.



## Buffon's Needle Problem (2/2)

Let us compute the probability that the line intersects:

$$\begin{aligned} \mathbf{P} \left[ X < \frac{L}{2} \cdot \cos(\theta) \right] &= \iint_{x < (L/2) \cos y} f_{X,\theta}(x, y) \, dx \, dy \\ &= \iint_{x < (L/2) \cos y} f_X(x) f_\theta(y) \, dx \, dy \\ &= \frac{4}{\pi D} \int_0^{\pi/2} \int_0^{L/2 \cos(y)} dx dy \\ &= \frac{4}{\pi D} \int_0^{\pi/2} \frac{L}{2} \cos(y) dy \\ &= \frac{2L}{\pi D}. \end{aligned}$$

This gives us a method to estimate  $\pi$ !



## Covariance

### Definition of Covariance

Let  $X$  and  $Y$  be two random variables. The **covariance** is defined as:

$$\mathbf{Cov}[X, Y] = \mathbf{E}[(X - \mathbf{E}[X]) \cdot (Y - \mathbf{E}[Y])].$$

Interpretation:

- If  $\mathbf{Cov}[X, Y] > 0$  and  $X$  has a realisation larger (smaller) than  $\mathbf{E}[X]$ , then  $Y$  will likely have a realisation larger (smaller) than  $\mathbf{E}[Y]$ .
- If  $\mathbf{Cov}[X, Y] < 0$ , then it is the other way around.

### Alternative Formula

Using the linearity of expectation rule, one has the equivalent definition:

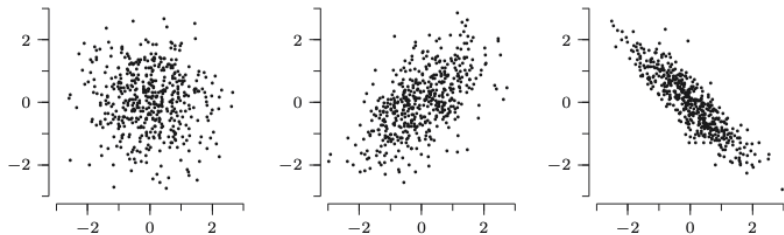
$$\mathbf{Cov}[X, Y] = \mathbf{E}[X \cdot Y] - \mathbf{E}[X] \cdot \mathbf{E}[Y].$$

- Note that  $\mathbf{Cov}[X, X] = \mathbf{V}[X]$ .
- Two variables  $X, Y$  with  $\mathbf{Cov}[X, Y] > 0$  are **positively correlated**.
- Two variables  $X, Y$  with  $\mathbf{Cov}[X, Y] < 0$  are **negatively correlated**.
- Two variables  $X, Y$  with  $\mathbf{Cov}[X, Y] = 0$  are **uncorrelated**.



## Illustration of 3 Cases for Cov [ $X, Y$ ]

500 outcomes of randomly generated pairs of RVs ( $X, Y$ ) with different joint distributions



**Fig. 10.1.** Some scatterplots.  
Source: Textbook by Dekking

1. What is the covariance (positive, negative, neutral)?
2. Where is the covariance the largest (in magnitude)?

## Independence implies Uncorrelated

### Example

Let  $X$  and  $Y$  be two **independent** random variables. Then  $X$  and  $Y$  are **uncorrelated**, i.e.,  $\mathbf{Cov}[X, Y] = 0$ .

Answer

We give a proof for the discrete case:





## Uncorrelated may not imply Independence

### Example

Find a (simple) example of two random variables  $X$  and  $Y$  which are **uncorrelated but dependent**.

Answer

- Let  $X$  be uniformly sampled from  $\{-1, 0, +1\}$  and  $Y := \mathbf{1}_{X=0}$ .  
 $\Rightarrow X \cdot Y = 0$  (for all outcomes), and thus

$$\mathbf{E}[X \cdot Y] = 0.$$

- Further,  $\mathbf{E}[X] = 0$  (and  $\mathbf{E}[Y] = 1/3$ ), and hence:

$$\mathbf{Cov}[X, Y] = \mathbf{E}[X \cdot Y] - \mathbf{E}[X] \cdot \mathbf{E}[Y] = 0.$$

- On the other hand,  $\mathbf{P}[X = 0] = 1/3$  and  $\mathbf{P}[Y = 0] = 2/3$ , and thus

$$1 = \mathbf{P}[X \cdot Y = 0] > \mathbf{P}[X = 0] \cdot \mathbf{P}[Y = 0] = 2/9.$$



### Variance of Sum Formula

- For any two random variables  $X, Y$ ,

$$\mathbf{V}[X + Y] = \mathbf{V}[X] + \mathbf{V}[Y] + 2 \cdot \mathbf{Cov}[X, Y].$$

- Hence if  $X$  and  $Y$  are **uncorrelated** variables,

$$\mathbf{V}[X + Y] = \mathbf{V}[X] + \mathbf{V}[Y].$$

Generalisation of the case where  $X$  and  $Y$  are even **independent**!

- For any random variables  $X_1, X_2, \dots, X_n$ :

$$\mathbf{V}\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n \mathbf{V}[X_i] + 2 \cdot \sum_{i=1}^n \sum_{j=i+1}^n \mathbf{Cov}[X_i, X_j].$$

## Computing Variances of Sums of Uncorrelated Variables

### Example

Recall the example where  $X \in \{-1, 0, +1\}$  uniformly and  $Y := \mathbf{1}_{X=0}$ . Compute  $\mathbf{V}[X + Y]$ .

Answer



## Correlation Coefficient: Normalising the Covariance

The definition of covariance is **not scaling invariant**:

- If  $X$  increases by a factor of  $\alpha$ , then  $\mathbf{Cov}[X, Y]$  increases by a factor of  $\alpha$ .
- ⇒ Even if  $X$  and  $Y$  both increase by  $\alpha$ , then  $\mathbf{Cov}[X, Y]$  will change.  
(Exercise: It changes by?)

### Correlation Coefficient

Let  $X$  and  $Y$  be two random variables. The **correlation coefficient**  $\rho(X, Y)$  is defined as:

$$\rho(X, Y) = \frac{\mathbf{Cov}[X, Y]}{\sqrt{\mathbf{V}[X] \cdot \mathbf{V}[Y]}}$$

If  $\mathbf{V}[X] = 0$  or  $\mathbf{V}[Y] = 0$ , then it is defined as 0.

### Properties:

1. The correlation coefficient is **scaling-invariant**, i.e.,  
 $\rho(X, Y) = \rho(\alpha \cdot X, \beta \cdot Y)$  for any  $\alpha, \beta > 0$ .
2. For any two random variables  $X, Y$ ,  $\rho(X, Y) \in [-1, 1]$ .



## Range of the Correlation Coefficient

### Example

Verify that the correlation coefficients' range satisfies  $\rho(X, Y) \in [-1, 1]$ .

Answer

- We will only prove  $\rho(X, Y) \geq -1$  (the other direction follows in analogous way).
- Let  $\sigma_x^2$  and  $\sigma_y^2$  denote the variances of  $X$  and  $Y$ , and  $\sigma_x$  and  $\sigma_y$  their standard deviations.
- Then:

$$\begin{aligned} 0 &\leq \mathbf{V} \left[ \frac{X}{\sigma_x} + \frac{Y}{\sigma_y} \right] \\ &= \mathbf{V} \left[ \frac{X}{\sigma_x} \right] + \mathbf{V} \left[ \frac{Y}{\sigma_y} \right] + 2 \mathbf{Cov} \left[ \frac{X}{\sigma_x}, \frac{Y}{\sigma_y} \right] \\ &= \frac{\mathbf{V}[X]}{\mathbf{V}[X]} + \frac{\mathbf{V}[Y]}{\mathbf{V}[Y]} + 2 \cdot \frac{\mathbf{Cov}[X, Y]}{\sigma_x \cdot \sigma_y} \\ &= 2 \cdot (1 + \rho(X, Y)). \end{aligned}$$

